

**September 13, 2007**  
**SC705: Advanced Statistics**  
**Instructor: Natasha Sarkisian**  
**Class notes: Two-Level HLM Models**

We continue working with High School and Beyond data (included with HLM6 – see HLM folder → Examples → Chapter 2, data files: HSB1.sav and HSB2.sav).

After estimating a null model and assuring that we observe a significant amount of group-level variance, we proceed to build a multilevel explanatory model. A typical approach is to build such a model from bottom up.

**Model 1. Conditional model with random intercept (one way ANCOVA with random intercept)**

**LEVEL 1 MODEL**

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij}) + r_{ij}$$

**LEVEL 2 MODEL**

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10}$$

**MIXED MODEL**

$$\text{MATHACH}_{ij} = \gamma_{00} + \gamma_{10} * \text{SES}_{ij} + u_{0j} + r_{ij}$$

Sigma\_squared = 37.03440

Tau  
 INTRCPT1,B0 4.76815

Tau (as correlations)  
 INTRCPT1,B0 1.000

```

-----
Random level-1 coefficient   Reliability estimate
-----
INTRCPT1, B0                0.843
-----

```

The value of the likelihood function at iteration 6 = -2.332167E+004

The outcome variable is MATHACH  
 Final estimation of fixed effects:

```

-----
Fixed Effect                Coefficient   Standard
                             Error          T-ratio
-----
For      INTRCPT1, B0
         INTRCPT2, G00      12.657481   0.187984   67.333
For      SES slope, B1
         INTRCPT2, G10      2.390199   0.105719   22.609
-----
Approx.
d.f.    P-value
-----
159    0.000
7183   0.000
-----

```

The outcome variable is MATHACH  
 Final estimation of fixed effects

(with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	12.657481	0.187330	67.568	159	0.000
For SES slope, B1					
INTRCPT2, G10	2.390199	0.119309	20.034	7183	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, U0	2.18361	4.76815	159	1037.09077	0.000
level-1, R	6.08559	37.03440			

Statistics for current covariance components model

Deviance = 46643.331427  
 Number of estimated parameters = 2

Note that we now estimate two fixed effects – the intercept and the effect of student’s SES. The intercept  $\gamma_{00}$  is no longer the average math achievement – it is now math achievement for someone with all predictors equal to zero. In this case, it’s math achievement for someone with SES=0, but because the SES scale was designed to have a mean of 0, the intercept (12.66) is essentially the math achievement for someone with average SES. The effect of SES,  $\gamma_{10}$ , can be interpreted as follows: one unit increase in SES is associated with 2.39 unit increase in one’s math achievement. So math achievement for someone with SES being 1 unit above the mean would be:

$$12.66 + 2.39 = 15.05$$

Note that each  $\beta_{0j}$  is now the mean outcome for each group (i.e. school) adjusted for the differences among these groups in SES.

As we now accounted for some portion of the variance by controlling for SES, we can calculate the adjusted intra-class correlation:  $\rho = 4.76815 / (4.76815 + 37.03440) = .11406362$

The decrease in  $\rho$  from .18035673 to .11406362 reflects a reduction in the relative share of between-school variance when we control for student SES. But there is still significant variation across schools.

We could also calculate the proportion of variance explained at each level by comparing the current variance estimates to those in the null model. (This is the easiest method recommended by Bryk and Raudenbush; another method is suggested by Snijders and Bosker and described in Luke book, p.35-37)IO

$$(8.61431 - 4.76815) / 8.61431 = .44648498$$

$$(39.14831 - 37.03440) / 39.14831 = .05399748$$

So controlling for individuals' SES explained 45% of between-school variance, and 5% of within-school variance in math achievement. We could also calculate the total percentage of variance explained:

$$(39.14831+8.61431-4.76815-37.03440)/(39.14831+8.61431)= .12478524$$

So students' SES explained 12% of the total variance in math achievement.

Let's take this one step further. So far we assumed that an individual student's SES would have the same impact on his or her math achievement regardless of the school where that student is studying. Let's relax that assumption.

## Model 2. Model with random intercept and random slopes (one way ANCOVA with random intercept and slopes)

### LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij}) + r_{ij}$$

### LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

Here, level-1 slopes are allowed to vary across level-2 units. But we do not try to predict that variation – only describe it.

Now we have:

$\gamma_{00}$  is the average intercept across the level-2 units (grand mean of math achievement controlling for SES – i.e. the mean for someone with SES=0)

$\gamma_{10}$  is the average SES slope across the level-2 units (i.e. average effect of SES across schools)

$u_{0j}$  is the unique addition to the intercept associated with level-2 unit j (indicates how the intercept for school j differs from the grand mean)

$u_{1j}$  is the unique addition to the slope associated with level-2 unit j (indicates how the effect of SES in school j differs from the average effect of SES for all schools)

Note that:

$$\begin{pmatrix} u_{0j} \\ u_{1j} \end{pmatrix} \sim N \left( \mathbf{0}, \begin{pmatrix} \tau_{00} & \tau_{01} \\ \tau_{10} & \tau_{11} \end{pmatrix} \right)$$

Our tau matrix now contains the variance in the level-1 intercepts ( $\tau_{00}$ ), the variance in level-1 slopes ( $\tau_{11}$ ), as well as the covariance between level-1 intercepts and slopes ( $\tau_{01} = \tau_{10}$ ). Note that covariance value indicates how much intercepts and slopes covary: in our example (below), there is a negative correlation between intercepts and slopes. That is, the higher the intercept, the smaller the slope (i.e. if the school level of math achievement is high, the effect of SES in that school is smaller).

Sigma\_squared = 36.82835

Tau		
INTRCPT1,B0	4.82978	-0.15399
SES,B1	-0.15399	0.41828

Tau (as correlations)  
 INTRCPT1,B0 1.000 -0.108  
 SES,B1 -0.108 1.000

```
-----
Random level-1 coefficient   Reliability estimate
-----
INTRCPT1, B0                0.797
SES, B1                     0.179
-----
```

The value of the likelihood function at iteration 21 = -2.331928E+004

The outcome variable is MATHACH

Final estimation of fixed effects:

```
-----
Fixed Effect                Coefficient   Standard Error   T-ratio   Approx. d.f.   P-value
-----
For      INTRCPT1, B0
INTRCPT2, G00              12.664935    0.189874        66.702    159          0.000
For      SES slope, B1
INTRCPT2, G10              2.393878    0.118278        20.240    159          0.000
-----
```

The outcome variable is MATHACH

Final estimation of fixed effects  
 (with robust standard errors)

```
-----
Fixed Effect                Coefficient   Standard Error   T-ratio   Approx. d.f.   P-value
-----
For      INTRCPT1, B0
INTRCPT2, G00              12.664935    0.189251        66.921    159          0.000
For      SES slope, B1
INTRCPT2, G10              2.393878    0.117697        20.339    159          0.000
-----
```

Final estimation of variance components:

```
-----
Random Effect                Standard Deviation   Variance Component   df   Chi-square   P-value
-----
INTRCPT1, U0                 2.19768              4.82978              159   905.26472    0.000
SES slope, U1                0.64675              0.41828              159   216.21178    0.002
level-1, R                   6.06864              36.82835
-----
```

Statistics for current covariance components model

```
-----
Deviance                      = 46638.560929
Number of estimated parameters = 4
-----
```

Here, like in the previous model, the math achievement for someone with average SES (SES=0) is 12.66; each unit increase in SES is associated with 2.39 units increase in math achievement. But, examining variance components, we notice that there is a significant variation in slopes (p-value =.002) – this means that SES effects vary across schools, so 2.39 is the effect for an average school. Here, if we want to divide the unexplained variance into within-school and between-school, we need to take into account the covariance: level 1 component is simply 36.82835, but level 2 component is  $(4.82978+0.41828+2*-0.15399)= 4.94008$ .

Note that in addition to the average reliability of school means, we now also have an estimate of reliability for the effect of SES, and it is much lower: .179. It is normal that the reliability of slopes is much lower than that of intercepts. The precision of estimation of the intercept (which in this case is a school mean) depends only on the sample size within each school. The precision of estimation of the slope depends both on the sample size and on the variability of SES within that school. Schools that are homogeneous with respect to SES will exhibit slope estimation with poor precision. But the average reliability of the slopes is relatively low because the true slope variance across schools is much smaller than the variance of the true means.

Note that low reliabilities do not invalidate the HLM analysis, but very low reliabilities (typically < .10) often indicate that a random coefficient might be considered fixed (i.e., the same across groups) in subsequent analyses.

### Model 3. Means-as-outcomes model (a.k.a. Intercepts as outcomes)

#### LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + r_{ij}$$

#### LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j}$$

This model allows us to predict variation in the levels of math achievement using level-2 variables. If we would attempt to do this using regular OLS, we would be artificially inflating the sample size and pretend we have 7185 data points to evaluate the effect of type of school (Catholic vs public), when in fact it's only 160 schools. Aggregating the data to school level would be more acceptable, but we would not have any assessment of within-school variation. Note, however, that the sample size for level 2 becomes important as soon as you try to include predictors at this level!

Sigma\_squared = 39.15135

Tau  
INTRCPT1,B0 6.67771

Tau (as correlations)  
INTRCPT1,B0 1.000

```
-----
Random level-1 coefficient   Reliability estimate
-----
INTRCPT1, B0                0.877
-----
```

The value of the likelihood function at iteration 4 = -2.353915E+004

The outcome variable is MATHACH  
Final estimation of fixed effects:

```
-----
Fixed Effect                Coefficient   Standard      Approx.
                        Error          T-ratio      d.f.        P-value
-----
For      INTRCPT1, B0
INTRCPT2, G00                11.393043   0.292887     38.899      158      0.000
SECTOR, G01                   2.804889   0.439142      6.387      158      0.000
-----
```

-----  
 The outcome variable is MATHACH  
 Final estimation of fixed effects  
 (with robust standard errors)  
 -----

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.393043	0.292258	38.983	158	0.000
SECTOR, G01	2.804889	0.435823	6.436	158	0.000

-----  
 Final estimation of variance components:  
 -----

Random Effect		Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, level-1,	U0	2.58413	6.67771	158	1296.76559	0.000
	R	6.25710	39.15135			

-----  
 Statistics for current covariance components model  
 -----

Deviance = 47078.295826  
 Number of estimated parameters = 2

Here, we see a positive effect of Catholic schools on math achievement – the average achievement of Catholic schools is 2.8 units higher than for public schools. The intercept now is an average value for a public school student. There is, nevertheless, significant school-level variance remaining. As we did with earlier models, we can calculate the percentage of variance in math achievement explained by school type. Note that here we only explain level 2 variance – level 1 variance remained the same. For level 2 variance:

$$(8.61431 - 6.67771)/8.61431 = .22481197$$

So 22% of school-level variance in math achievement was explained by type of school.

#### **Model 4. Means as outcomes model with level 1 covariate**

As a next step, we can add level-1 covariates to this means-as-outcomes model. These level-1 variables can be added as fixed effects (i.e., assuming that the effects of these covariates are the same for all schools –that’s what we did in model 1) or as random effects (i.e., assuming that the effects of level 1 variables vary across schools – that’s what we did in model 2). We will right away opt for a more complex option, assuming that the effects of level 1 variable – SES – vary across schools.

**LEVEL 1 MODEL**

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij}) + r_{ij}$$

**LEVEL 2 MODEL**

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

Sigma\_squared = 36.79508

Tau		
INTRCPT1,B0	3.96459	0.71641
SES,B1	0.71641	0.44990

Tau (as correlations)

INTRCPT1,B0	1.000	0.536
SES,B1	0.536	1.000

Random level-1 coefficient	Reliability estimate
INTRCPT1, B0	0.765
SES, B1	0.189

The value of the likelihood function at iteration 21 = -2.330093E+004  
 The outcome variable is MATHACH

Final estimation of fixed effects:

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.476646	0.231587	49.557	158	0.000
SECTOR, G01	2.533835	0.344798	7.349	158	0.000
For SES slope, B1					
INTRCPT2, G10	2.385451	0.118329	20.160	159	0.000

The outcome variable is MATHACH

Final estimation of fixed effects  
 (with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.476646	0.225026	51.001	158	0.000
SECTOR, G01	2.533835	0.352411	7.190	158	0.000
For SES slope, B1					
INTRCPT2, G10	2.385451	0.119008	20.044	159	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, U0	1.99113	3.96459	158	766.83844	0.000

SES slope, U1	0.67075	0.44990	159	216.12223	0.002
level-1, R	6.06589	36.79508			

-----  
 Statistics for current covariance components model  
 -----

Deviance = 46601.861400  
 Number of estimated parameters = 4

Now the intercept is the value for average SES student in a public school: 11.48. The value for an average-SES Catholic school student is 2.53 units higher: 11.45+2.53=13.98

Further, one unit increase in SES is associated with 2.39 units increase in math score. But there is still significant variation across schools in intercepts, and there is also significant variation in SES slopes – so SES doesn't have the same effect across schools.

### Model 5. Intercepts and Slopes as outcomes (a.k.a. Cross-level Interactions model)

Next, we will try to explain this variation in SES effects across schools – we'll explore whether this variation can be attributed to the type of school – public vs Catholic.

#### LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij}) + r_{ij}$$

#### LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{SECTOR}_j) + u_{1j}$$

This type of model allows us to explain the variation in both intercepts and slopes. Sometimes, it's called cross-level interactions model because we make the effect of level-1 variables (SES) dependent upon the value of level-2 variables (in this case, SECTOR).

Sigma\_squared = 36.76311

Tau		
INTRCPT1,B0	3.83295	0.54112
SES,B1	0.54112	0.12988

Tau (as correlations)		
INTRCPT1,B0	1.000	0.767
SES,B1	0.767	1.000

-----  
 Random level-1 coefficient    Reliability estimate  
 -----

INTRCPT1, B0	0.759
SES, B1	0.064

-----  
 The value of the likelihood function at iteration 198 = -2.328373E+004

The outcome variable is MATHACH  
 Final estimation of fixed effects:

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					

	INTRCPT2, G00	11.750237	0.232241	50.595	158	0.000
	SECTOR, G01	2.128611	0.346651	6.141	158	0.000
For	SES slope, B1					
	INTRCPT2, G10	2.958798	0.145460	20.341	158	0.000
	SECTOR, G11	-1.313096	0.219062	-5.994	158	0.000

The outcome variable is MATHACH  
Final estimation of fixed effects  
(with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value	
For	INTRCPT1, B0					
	INTRCPT2, G00	11.750237	0.218675	53.734	158	0.000
	SECTOR, G01	2.128611	0.355697	5.984	158	0.000
For	SES slope, B1					
	INTRCPT2, G10	2.958798	0.144092	20.534	158	0.000
	SECTOR, G11	-1.313096	0.214271	-6.128	158	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, U0	1.95779	3.83295	158	756.04082	0.000
SES slope, U1	0.36039	0.12988	158	178.09113	0.131
level-1, R	6.06326	36.76311			

Statistics for current covariance components model

Deviance = 46567.464841  
Number of estimated parameters = 4

In terms of fixed effects, the difference between this model and the previous one is the introduction of the effect of SECTOR on SES, which can be interpreted as an interaction term between SECTOR and SES. That is, the effect of SES for public schools is 2.96 per one unit increase in SES; but for Catholic schools, the effect of SES is  $(2.96-1.31)=1.65$  per one unit increase in SES. So students' math scores are more sensitive to their SES in public schools than in Catholic schools.

We can also examine the amount of variance in SES slopes explained by SECTOR: the unconditional variance in SES slopes was 0.44990, and the variance in this model (controlling for SECTOR) is only 0.12988.

$$(0.44990-0.12988)/0.44990 = .71131363$$

So SECTOR explained 71% of between-school variance in effects of SES on math achievement. Also note that now that we controlled for sector, the variation in SES slopes across schools is no longer significant. Therefore, we could run this model as a model with nonrandomly varying slopes.

**Model 6. Model with Nonrandomly Varying Slopes.**

**LEVEL 1 MODEL**

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij}) + r_{ij}$$

**LEVEL 2 MODEL**

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{SECTOR}_j)$$

Sigma\_squared = 36.84019

Tau  
INTRCPT1,B0 3.69423

Tau (as correlations)  
INTRCPT1,B0 1.000

Random level-1 coefficient	Reliability estimate
INTRCPT1, B0	0.808

The value of the likelihood function at iteration 6 = -2.328616E+004

The outcome variable is MATHACH  
Final estimation of fixed effects:

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.797994	0.228514	51.629	158	0.000
SECTOR, G01	2.138170	0.341344	6.264	158	0.000
For SES slope, B1					
INTRCPT2, G10	2.951177	0.140609	20.989	7181	0.000
SECTOR, G11	-1.312849	0.211996	-6.193	7181	0.000

The outcome variable is MATHACH  
Final estimation of fixed effects  
(with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.797994	0.214976	54.880	158	0.000
SECTOR, G01	2.138170	0.350028	6.109	158	0.000
For SES slope, B1					
INTRCPT2, G10	2.951177	0.143779	20.526	7181	0.000
SECTOR, G11	-1.312849	0.212824	-6.169	7181	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
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INTRCPT1,	U0	1.92204	3.69423	158	837.19099	0.000
level-1,	R	6.06961	36.84019			

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Statistics for current covariance components model

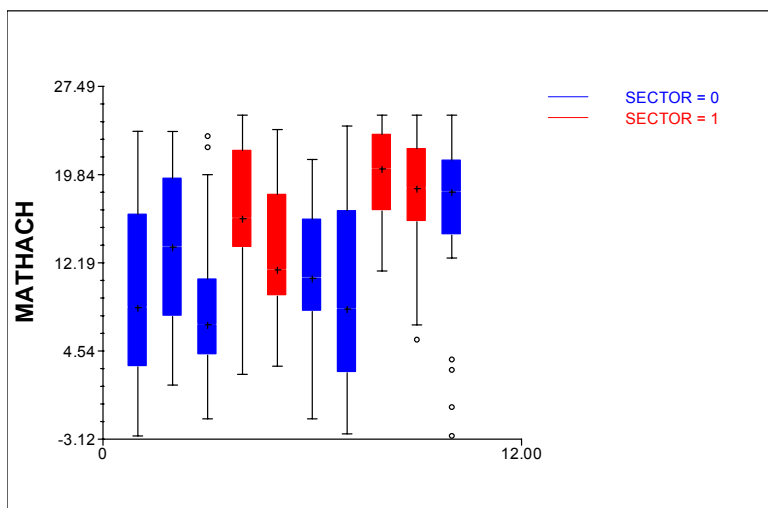
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Deviance = 46572.326387  
 Number of estimated parameters = 2

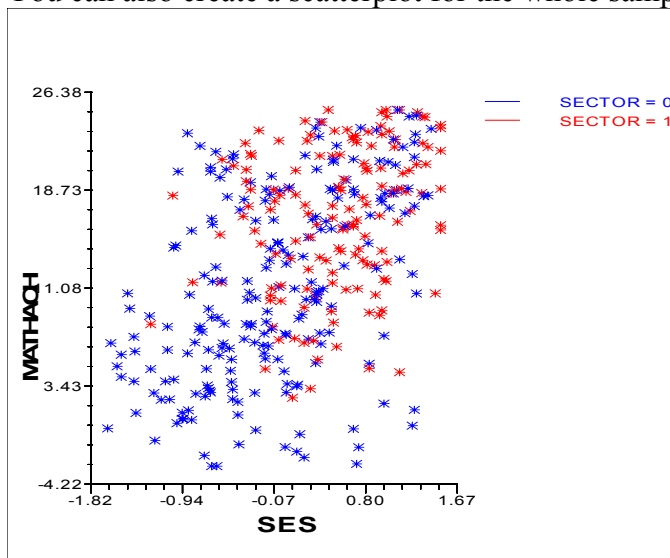
Note that we are able to model how sector shapes SES, but we do not allow any other variation in SES slopes because there is no significant variation beyond that accounted for by sector.

### Using graphs to examine the data

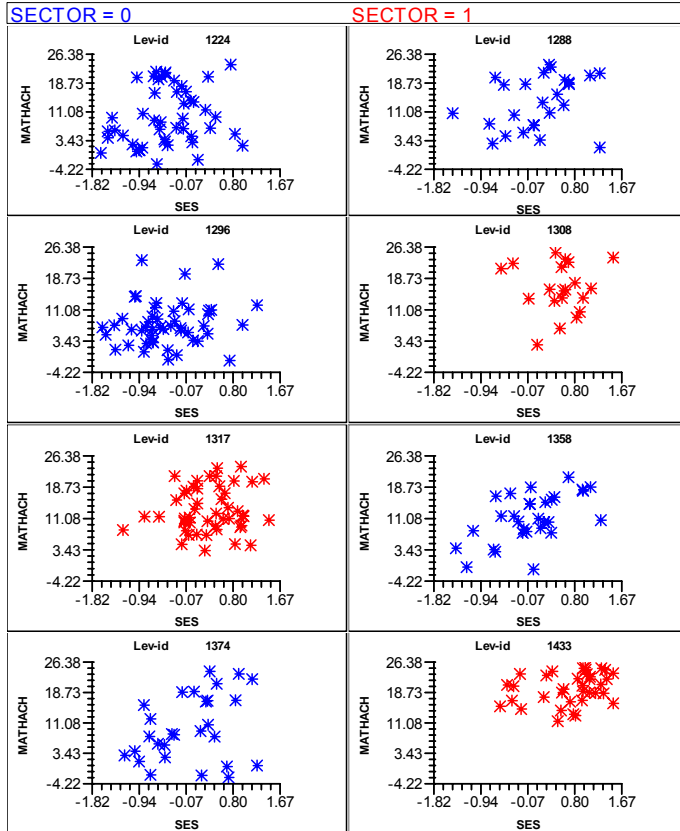
HLM has some limited graphing capabilities allowing you to examine the data before starting to build models. You can examine your data by creating boxplots for a variable, e.g., your dependent variable, by group. You can also mark these groups according to one level-2 variable:



You can also create a scatterplot for the whole sample by values of a level-2 variable:



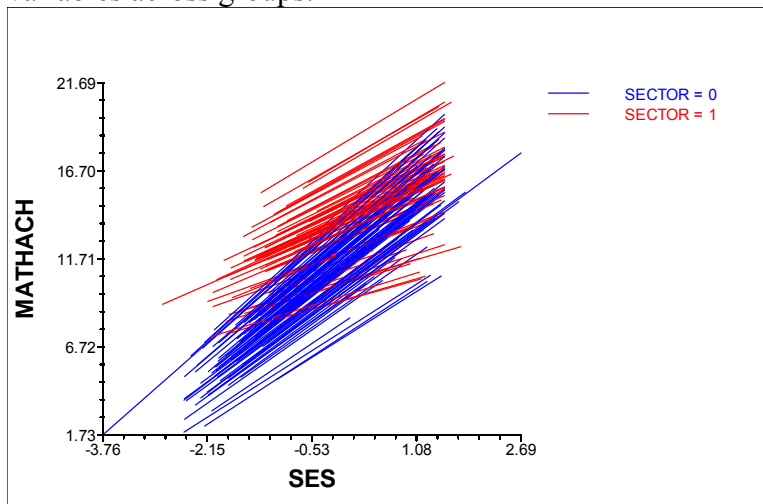
Or you can create scatterplots separating groups and colorcoding them by a level-2 variable:



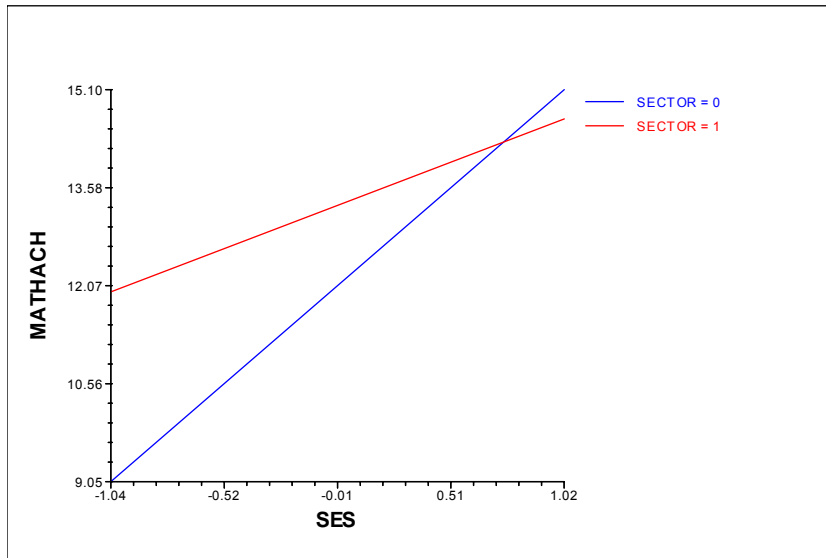
### Graphing Equations

Graphs can also be used to better illustrate and understand the models you estimate. These can greatly assist in interpreting the findings. HLM offers a range of such graphing options.

E.g., Graph Equations → Level 1 equation graphing. Here you can examine slopes for level-1 variables across groups:



Or you can graph the relationships based on the fixed effects in your last model using Graph Equations → Model graphs. Here, you have a range of options. For example, you can look at how level 1 slopes vary depending on values of level 2 variables (if you have a cross-level interaction in your model). You need to select a level-1 variable as you X, and level-2 variable as Z focus:



Or you can examine how predicted values vary by level of both level-1 and level-2 variables by selecting level-2 variable as your X, and level-1 as your Z-focus:

