

September 20, 2007
SC705: Advanced Statistics
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Class notes: HLM Model Building Strategies

Model Selection Strategy

To summarize, we saw that multilevel models can include 3 types of predictors:

- Level-1 predictors (e.g., student SES)
- Level-2 predictors (e.g., school SECTOR)
- Level-1 predictors aggregated to level 2 (e.g., MEANSES)

In addition, we have a number of choices:

- The intercept can be estimated as either fixed or random (typically random)
- The effects of level 1 predictors can be estimated as either fixed effects or random effects
- Level 2 predictors can be used to predict the intercept (i.e., as direct predictors of DV)
- Level 2 predictors can explain the variation in slopes of level 1 predictors (i.e., as cross-level interactions)

Because so many components are involved, it is best to proceed incrementally.

1. Start by fitting a fully unconditional model. Evaluate level 2 variance to see if HLM is necessary.
2. Estimate a model with random intercept and slopes using only level 1 variables (all slopes should be random effects). Evaluate slope variance and decide whether some slopes should be fixed.
3. Estimate means-as-outcomes with level 1 covariates model to select level 2 predictors of intercept.
4. For slopes with significant variance, use level 2 predictors to explain that variance (i.e., estimate an intercepts-and-slopes-as-outcomes model).
5. If the slope variance remaining after entering level 2 predictors is not statistically significant, estimate that slope as non-randomly varying.

When making decisions what variables to include and whether to estimate random or fixed effects, we can use hypothesis testing tools.

Using Hypothesis Testing to Build Models

HLM6 allows you to test various hypotheses which can be helpful when evaluating which variables and which random effects to include in your model. The basic idea behind hypothesis testing is to build a set of contrasts that would add up to zero under the null hypothesis, and then test the hypothesis that, combined, they are indeed zero.

1. Single parameter tests of significance.

Single parameter tests are presented in your regular HLM output; in practice, there is no need to run such tests in addition to the regular output, but for learning purposes, we will start with these. Suppose we want to test whether a specific coefficient (e.g. the SES slope for average SES

public schools (i.e., intercept for SES slope, gamma 10) is zero. The set of contrasts that we will specify for that will include 1 for G10 and 0 for everything else; therefore, we will test $H_0: G10=0$.

Results of General Linear Hypothesis Testing

	Coefficients	Contrast
For INTRCPT1, B0		
INTRCPT2, G00	12.096006	0.000
SECTOR, G01	1.226384	0.000
MEANSES, G02	5.333056	0.000
For SES slope, B1		
INTRCPT2, G10	2.937981	1.000
SECTOR, G11	-1.640954	0.000
MEANSES, G12	1.034427	0.000

Chi-square statistic = 396.102629
 Degrees of freedom = 1
 P-value = 0.000000

Here, we reject H_0 .

2. Multi-parameter tests of significance.

Here, we test the hypothesis that multiple coefficients are all equal to 0. Typically, we do that in order to decide whether they can be omitted from the model. This can either be coefficients for different variables (possibly related, e.g. sets of dummies), or coefficients for the same variable in different parts of the model. For example, for could test that all coefficients for SES slope are zero. That would mean testing a combined hypothesis:

$G10=0$

$G11=0$

$G12=0$

We can do that by selecting 1 for each of these coefficients when selecting contrasts (separate column for each):

Results of General Linear Hypothesis Testing

	Coefficients	Contrast		
For INTRCPT1, B0				
INTRCPT2, G00	12.095250	0.000	0.000	0.000
SECTOR, G01	1.224401	0.000	0.000	0.000
MEANSES, G02	5.336698	0.000	0.000	0.000
For SES slope, B1				
INTRCPT2, G10	2.935664	1.000	0.000	0.000
SECTOR, G11	-1.642102	0.000	1.000	0.000
MEANSES, G12	1.044120	0.000	0.000	1.000

Chi-square statistic = 516.728437
 Degrees of freedom = 3
 P-value = 0.000000

We reject H_0 ; the coefficients associated with SES slope are jointly significant. We can also test whether MEANSES is significant across equations:

Results of General Linear Hypothesis Testing

		Coefficients	Contrast	
For	INTRCPT1, B0			
	INTRCPT2, G00	12.095250	0.000	0.000
	SECTOR, G01	1.224401	0.000	0.000
	MEANSES, G02	5.336698	1.000	0.000
For	SES slope, B1			
	INTRCPT2, G10	2.935664	0.000	0.000
	SECTOR, G11	-1.642102	0.000	0.000
	MEANSES, G12	1.044120	0.000	1.000
Chi-square statistic = 257.389086				
Degrees of freedom = 2				
P-value = 0.000000				

3. Tests for equality of coefficients.

We can also test whether two or more coefficients are equal. This is typically used when we have a series of related dummy variables, and we want to combine some dummies. E.g., we could have students' racial/ethnic identification, with dummy variables representing African American, Mexican American, Puerto Rican, Asian American, etc., the omitted category being White. We could then wonder whether we could simplify that into one dichotomy, White vs ethnic minority. But to test whether the data support this simplification, we'd test whether coefficients for each ethnic group equal to each other (i.e., are all groups different from Whites in the same way?). We don't have sets of dummy variables in this dataset, so I will show how to do it with a less realistic example. Suppose I want to test if the effect of SECTOR equals the effect of MEANSES. Then I test:

$$G01=G02$$

$$G10=G11$$

To do this using contrasts, we redefine it as:

$$G01-G02=0$$

$$G10-G11=0$$

Results of General Linear Hypothesis Testing				
		Coefficients	Contrast	
For	INTRCPT1, B0			
	INTRCPT2, G00	12.096006	0.000	0.000
	SECTOR, G01	1.226384	-1.000	0.000
	MEANSES, G02	5.333056	1.000	0.000
For	SES slope, B1			
	INTRCPT2, G10	2.937981	0.000	0.000
	SECTOR, G11	-1.640954	0.000	1.000
	MEANSES, G12	1.034427	0.000	-1.000
Chi-square statistic = 83.658410				
Degrees of freedom = 2				
P-value = 0.000000				

Here, we reject H_0 , so we wouldn't be able to combine those variables (not that we really wanted to in this hypothetical example).

4. Tests for variance components

If we are interested in testing hypotheses about variance components or their combinations, we should utilize likelihood ratio tests based on deviance values. To use such test, we estimate two models, and calculate D0-D1. The resulting difference follows chi-square distribution with $df = \text{number of parameters for model 0} - \text{number of parameters for model 1}$.

Level-1 Model

$$Y = B0 + B1*(SES) + R$$

Level-2 Model

$$B0 = G00 + G01*(SECTOR) + G02*(MEANSES) + U0$$

$$B1 = G10 + G11*(SECTOR) + G12*(MEANSES) + U1$$

Deviance = 46501.875643

Number of estimated parameters = 4

Level-1 Model

$$Y = B0 + B1*(SES) + R$$

Level-2 Model

$$B0 = G00 + G01*(SECTOR) + G02*(MEANSES) + U0$$

$$B1 = G10 + G11*(SECTOR) + G12*(MEANSES)$$

Statistics for current covariance components model

Deviance = 46502.952743

Number of estimated parameters = 2

Variance-Covariance components test

Chi-square statistic = 1.07710

Number of degrees of freedom = 2

P-value = >.500

P-value indicates that there is no significant difference in model fit between these two models. Of course, we already knew that based on the output for variance components. But if we have multiple slopes that we think should be fixed rather than random, we can do such a test for more than one variance component simultaneously by comparing a model with random slopes to that where these slopes are fixed.

The issue of centering

You have already noticed that HLM6 asks you whether and how you'd like to center your predictors. Here, we will discuss the issues involved in making these decisions.

Level-1 predictors:

1. Natural metric (X):

You should only use the original metric if the value of 0 for a predictor is a meaningful value. When 0 is not meaningful, the estimate of the intercept will be arbitrary and may be estimated with poor precision. Lack of precision in HLM can be very problematic. First, because you are

estimating within-group intercepts, thus with possibly small N, the estimates may be quite unstable. Second, because you may be trying to model variation in these intercepts, your model will be affected by the unreliability of the estimates.

2. Grand-mean centering (X - grand mean):

This will address the problems with estimation of intercept in original metric. Because the 0 values will fall in the middle of the distribution of the predictors, the intercept estimates will be estimated with much more precision. The intercept is also interpretable. Specifically, it will represent the group-mean value for a person with a (grand) average on every predictor. The interpretation of the intercepts is now “adjusted group mean.” The interpretation of slopes does not change. E.g. our measure of SES is already grand-mean centered because it is a standardized scale. So we can interpret the fixed effect for the intercept as the average math achievement adjusted for SES – i.e., the average math achievement for someone with average SES.

LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{FEMALE}_{ij}) + \beta_{2j}(\text{SES}_{ij} - \overline{\text{SES}_{..}}) + r_{ij}$$

LEVEL 2 MODEL

$$\begin{aligned} \beta_{0j} &= \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j} \\ \beta_{1j} &= \gamma_{10} + \gamma_{11}(\text{SECTOR}_j) + u_{1j} \\ \beta_{2j} &= \gamma_{20} + \gamma_{21}(\text{SECTOR}_j) \end{aligned}$$

Level-1 Coefficients	Level-2 Predictors
INTRCPT1, B0	INTRCPT2, G00 SECTOR, G01
FEMALE slope, B1	INTRCPT2, G10 SECTOR, G11
# SES slope, B2	INTRCPT2, G20 SECTOR, G21

'#' - The residual parameter variance for this level-1 coefficient has been set to zero.

Summary of the model specified (in equation format)

Level-1 Model

$$Y = B0 + B1*(FEMALE) + B2*(SES) + R$$

Level-2 Model

$$\begin{aligned} B0 &= G00 + G01*(SECTOR) + U0 \\ B1 &= G10 + G11*(SECTOR) + U1 \\ B2 &= G20 + G21*(SECTOR) \end{aligned}$$

Sigma_squared = 36.45566

Tau

INTRCPT1,B0	4.20999	-1.19503
FEMALE,B1	-1.19503	1.10455

Tau (as correlations)
 INTRCPT1,B0 1.000 -0.554
 FEMALE,B1 -0.554 1.000

Random level-1 coefficient	Reliability estimate
INTRCPT1, B0	0.681
FEMALE, B1	0.236

Final estimation of fixed effects
 (with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	12.437528	0.253308	49.100	158	0.000
SECTOR, G01	2.084354	0.416840	5.000	158	0.000
For FEMALE slope, B1					
INTRCPT2, G10	-1.222739	0.220850	-5.537	158	0.000
SECTOR, G11	0.031679	0.401264	0.079	158	0.938
For SES slope, B2					
INTRCPT2, G20	2.919985	0.141351	20.658	7179	0.000
SECTOR, G21	-1.293137	0.207803	-6.223	7179	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, U0	2.05183	4.20999	121	347.17034	0.000
FEMALE slope, U1	1.05098	1.10455	121	153.20942	0.025
level-1, R	6.03785	36.45566			

Note that while it may seem inappropriate at first to center a dummy variable, in HLM it actually is quite useful. If uncentered, the intercept in a model with a dummy variable is the average value when the dummy variable is 0. If the dummy variable is centered, the intercept then becomes the mean adjusted for the proportion of cases with the dummy variable=1. For example, if the indicator for sex variable is centered around the grand mean, this centered predictor can take two values. If the subject is female, it will equal the proportion of male students in the sample. If the subject is male, it will equal to minus the proportion of female students in the sample. Zero on this variable becomes the average proportion of female students. The intercept again will be the adjusted group mean – in this case, it is adjusted for the difference among level-2 units in the percentage of female students.

LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{FEMALE}_{ij} - \overline{\text{FEMALE}_{..}}) + \beta_{2j}(\text{SES}_{ij} - \overline{\text{SES}_{..}}) + r_{ij}$$

LEVEL 2 MODEL

$$\begin{aligned} \beta_{0j} &= \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j} \\ \beta_{1j} &= \gamma_{10} + \gamma_{11}(\text{SECTOR}_j) + u_{1j} \\ \beta_{2j} &= \gamma_{20} + \gamma_{21}(\text{SECTOR}_j) \end{aligned}$$

	Level-1 Coefficients	Level-2 Predictors
	INTRCPT1, B0	INTRCPT2, G00 SECTOR, G01
%	FEMALE slope, B1	INTRCPT2, G10 SECTOR, G11
#	SES slope, B2	INTRCPT2, G20 SECTOR, G21

'#' - The residual parameter variance for this level-1 coefficient has been set to zero.

'%' - This level-1 predictor has been centered around its grand mean.

Level-1 Model

$$Y = B0 + B1*(FEMALE) + B2*(SES) + R$$

Level-2 Model

$$B0 = G00 + G01*(SECTOR) + U0$$

$$B1 = G10 + G11*(SECTOR) + U1$$

$$B2 = G20 + G21*(SECTOR)$$

Sigma_squared = 36.45669

Tau

INTRCPT1,B0	3.25698	-0.61167
FEMALE,B1	-0.61167	1.09435

Tau (as correlations)

INTRCPT1,B0	1.000	-0.324
FEMALE,B1	-0.324	1.000

Random level-1 coefficient	Reliability estimate
INTRCPT1, B0	0.774
FEMALE, B1	0.234

Final estimation of fixed effects (with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	11.791658	0.214281	55.029	158	0.000
SECTOR, G01	2.101157	0.333510	6.300	158	0.000
For FEMALE slope, B1					
INTRCPT2, G10	-1.222670	0.220851	-5.536	158	0.000
SECTOR, G11	0.032271	0.401209	0.080	158	0.936
For SES slope, B2					
INTRCPT2, G20	2.919869	0.141361	20.655	7179	0.000
SECTOR, G21	-1.292989	0.207806	-6.222	7179	0.000

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value

INTRCPT1,	U0	1.80471	3.25698	121	488.52692	0.000
FEMALE slope,	U1	1.04611	1.09435	121	153.19922	0.025
level-1,	R	6.03794	36.45669			

3. Group-mean centering (\bar{X} – group mean):

Predictors can also be centered around the mean value for the group to which they belong. The intercept can then be interpreted as the average outcome for each group. This allows interpretation of parameter estimates as person-level effects within each group (i.e. if you differ from your group's average by one unit, your math achievement will increase by \bar{X} units).

Again, we can group-mean center dummy variables as well. For females, we will get a value equal to the proportion of male students in school j ; for males, it will take the value equal to minus the proportion of females in that school. The fact that it is a dummy variable does not change the interpretation of the intercept when group mean-centering is employed.

LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{FEMALE}_{ij} - \overline{\text{FEMALE}}_{.j}) + \beta_{2j}(\text{SES}_{ij} - \overline{\text{SES}}_{.j}) + r_{ij}$$

LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{SECTOR}_j) + u_{1j}$$

$$\beta_{2j} = \gamma_{20} + \gamma_{21}(\text{SECTOR}_j)$$

The outcome variable is MATHACH
The model specified for the fixed effects was:

Level-1 Coefficients	Level-2 Predictors
INTRCPT1, B0	INTRCPT2, G00 SECTOR, G01
* FEMALE slope, B1	INTRCPT2, G10 SECTOR, G11
#* SES slope, B2	INTRCPT2, G20 SECTOR, G21

'#' - The residual parameter variance for this level-1 coefficient has been set to zero.

'*' - This level-1 predictor has been centered around its group mean.

Summary of the model specified (in equation format)

Level-1 Model

$$Y = B0 + B1*(\text{FEMALE}) + B2*(\text{SES}) + R$$

Level-2 Model

$$B0 = G00 + G01*(\text{SECTOR}) + U0$$

$$B1 = G10 + G11*(\text{SECTOR}) + U1$$

$$B2 = G20 + G21*(\text{SECTOR})$$

$$\text{Sigma_squared} = 36.45732$$

```

Tau
INTRCPT1,B0      6.75745      -0.63530
FEMALE,B1       -0.63530      0.82580

```

```

Tau (as correlations)
INTRCPT1,B0  1.000 -0.269
FEMALE,B1 -0.269  1.000

```

```

-----
Random level-1 coefficient   Reliability estimate
-----
INTRCPT1, B0                0.882
FEMALE, B1                  0.188
-----

```

Note: The reliability estimates reported above are based on only 123 of 160 units that had sufficient data for computation. Fixed effects and variance components are based on all the data.

The value of the likelihood function at iteration 31 = -2.330178E+004

The outcome variable is MATHACH
Final estimation of fixed effects
(with robust standard errors)

```

-----
Fixed Effect           Coefficient   Standard Error   T-ratio   Approx. d.f.   P-value
-----
For      INTRCPT1, B0
INTRCPT2, G00          11.393469    0.292627        38.935     158     0.000
SECTOR, G01            2.804207    0.436272         6.428     158     0.000
For      FEMALE slope, B1
INTRCPT2, G10         -1.224963    0.218270        -5.612     158     0.000
SECTOR, G11            0.421184    0.422651         0.997     158     0.321
For      SES slope, B2
INTRCPT2, G20          2.732981    0.156703        17.440     7179    0.000
SECTOR, G21          -1.310898    0.229605        -5.709     7179    0.000
-----

```

Final estimation of variance components:

```

-----
Random Effect           Standard Deviation   Variance Component   df   Chi-square   P-value
-----
INTRCPT1, U0            2.59951             6.75745             121   890.99031    0.000
FEMALE slope, U1       0.90873             0.82580             121   150.58868    0.035
level-1, R              6.03799             36.45732
-----

```

Note: The chi-square statistics reported above are based on only 123 of 160 units that had sufficient data for computation. Fixed effects and variance components are based on all the data.

Important:

Under grand-mean centering or no centering, the parameter estimates reflect a combination of (1) person-level effects and (2) compositional effects. But when we use a group-centered predictor, we only estimate the person-level effects.

In order not to discard the compositional effects with group-mean centering, level-2 variables should be created to represent the group mean values for each group-mean centered predictor. Because the group mean is effectively removed from the individual scores, the level-2 values

will be orthogonal to the level-1 values. Note that while HLM6 has an option for group-mean centering, it does not compute the group mean values of a predictor to be included as a level-2 variable – you have to do that in another statistical package and then import the data into HLM (more on this below).

E.g. we can use group mean centering for SES and using mean SES as a school level variable (here, MEANSES is already in the dataset):

LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij} - \overline{\text{SES}}_{..}) + r_{ij}$$

LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j) + \gamma_{02}(\text{MEANSES}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{SECTOR}_j) + \gamma_{12}(\text{MEANSES}_j) + u_{1j}$$

```
-----
Level-1
Coefficients
-----
          INTRCPT1, B0
*      SES slope, B1
Level-2
Predictors
-----
          INTRCPT2, G00
          SECTOR, G01
          MEANSES, G02
          INTRCPT2, G10
          SECTOR, G11
          MEANSES, G12

''' - This level-1 predictor has been centered around its group mean.
Summary of the model specified (in equation format)
-----

Level-1 Model

      Y = B0 + B1*(SES) + R

Level-2 Model
      B0 = G00 + G01*(SECTOR) + G02*(MEANSES) + U0
      B1 = G10 + G11*(SECTOR) + G12*(MEANSES) + U1

Sigma_squared =      36.70313

Tau
INTRCPT1,B0      2.37996      0.19058
SES,B1           0.19058      0.14892

Tau (as correlations)
INTRCPT1,B0      1.000      0.320
SES,B1           0.320      1.000

-----
Random level-1 coefficient      Reliability estimate
-----
INTRCPT1, B0                      0.733
SES, B1                            0.073
-----
```

Final estimation of fixed effects
(with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value
For INTRCPT1, B0					
INTRCPT2, G00	12.096006	0.173699	69.638	157	0.000
SECTOR, G01	1.226384	0.308484	3.976	157	0.000
MEANSES, G02	5.333056	0.334600	15.939	157	0.000
For SES slope, B1					
INTRCPT2, G10	2.937981	0.147620	19.902	157	0.000
SECTOR, G11	-1.640954	0.237401	-6.912	157	0.000
MEANSES, G12	1.034427	0.332785	3.108	157	0.003

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value
INTRCPT1, U0	1.54271	2.37996	157	605.29503	0.000
SES slope, U1	0.38590	0.14892	157	162.30867	0.369
level-1, R	6.05831	36.70313			

Here, the effects of SES turn out to be quite complex: For those who are in a public school whose SES is at their school's average and whose school itself is average in terms of its SES, the math achievement is 12.096. If you are in a Catholic school with such properties, it's $12.1+1.2=13.3$. But if your school's average SES is 1 unit higher than the average for all schools, then your math achievement increases by 5.33. Further, in addition to these school-level effects, your individual SES also plays a role – if you are in an average (in terms of SES) public school, one unit increase in your SES will raise your math score by 2.94. In a Catholic school, that effect would be $2.94-1.64=1.30$. But if you are in a public school and your school is 1 unit above an average school in its SES, then your personal SES impact (per one unit) would be $2.94+1.03=3.97$. For a Catholic school in that situation, that effect of SES would become $2.94-1.64+1.03=2.33$. Interestingly, personal SES seems to have stronger impact on math achievement in those schools that have relatively high school-level SES.

The choice between grand-mean centering and group-mean centering depends on your theoretical thinking about processes. If you think that the absolute values of level 1 variable matter, then use grand-mean centering. If you think that it is the relative position of the person with regards to their group's mean is what matters, then use group-centering.

Level-2 predictors:

Centering issues for level-2 predictors are essentially the same issues faced in any regression. If the value of 0 for a predictor is not meaningful, the intercept will not have a meaningful interpretation and the estimate may lack precision. When these conditions exist, centering is advisable. You can either use grand-mean centering (then the intercept will reflect the average group) or center around some constant (then the intercept will reflect a group with the value of the predictor equal to that constant). Note that HLM6 only has the grand-mean centering option for level-2 predictors – if you want to center around some other value, you would have to

generate such a centered variable in another statistical program and then import the data into HLM. Typically, however, grand-mean centering is just fine.

LEVEL 1 MODEL

$$\text{MATHACH}_{ij} = \beta_{0j} + \beta_{1j}(\text{SES}_{ij} - \overline{\text{SES}}_{..}) + r_{ij}$$

LEVEL 2 MODEL

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{SECTOR}_j - \overline{\text{SECTOR}}_{.}) + \gamma_{02}(\text{MEANSES}_j - \overline{\text{MEANSES}}_{.}) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{SECTOR}_j - \overline{\text{SECTOR}}_{.}) + \gamma_{12}(\text{MEANSES}_j - \overline{\text{MEANSES}}_{.}) + u_{1j}$$

	Level-1 Coefficients	Level-2 Predictors
	INTRCPT1, B0	INTRCPT2, G00
\$		SECTOR, G01
\$		MEANSES, G02
*	SES slope, B1	INTRCPT2, G10
\$		SECTOR, G11
\$		MEANSES, G12

'*' - This level-1 predictor has been centered around its group mean.
 '\$' - This level-2 predictor has been centered around its grand mean.

Summary of the model specified (in equation format)

Level-1 Model

$$Y = B0 + B1*(SES) + R$$

Level-2 Model

$$B0 = G00 + G01*(SECTOR) + G02*(MEANSES) + U0$$

$$B1 = G10 + G11*(SECTOR) + G12*(MEANSES) + U1$$

Sigma_squared = 36.70313

Tau

INTRCPT1,B0	2.37996	0.19058
SES,B1	0.19058	0.14892

Tau (as correlations)

INTRCPT1,B0	1.000	0.320
SES,B1	0.320	1.000

Random level-1 coefficient Reliability estimate

INTRCPT1, B0	0.733
SES, B1	0.073

Final estimation of fixed effects
 (with robust standard errors)

Fixed Effect	Coefficient	Standard Error	T-ratio	Approx. d.f.	P-value

For INTRCPT1, B0					
INTRCPT2, G00	12.631549	0.140082	90.173	157	0.000
SECTOR, G01	1.226384	0.308484	3.976	157	0.000
MEANSES, G02	5.333056	0.334600	15.939	157	0.000
For SES slope, B1					
INTRCPT2, G10	2.219870	0.108224	20.512	157	0.000
SECTOR, G11	-1.640954	0.237401	-6.912	157	0.000
MEANSES, G12	1.034427	0.332785	3.108	157	0.003

Final estimation of variance components:

Random Effect	Standard Deviation	Variance Component	df	Chi-square	P-value

INTRCPT1, U0	1.54271	2.37996	157	605.29503	0.000
SES slope, U1	0.38590	0.14892	157	162.30867	0.369
level-1, R	6.05831	36.70313			

Creating Aggregated Variables from Level 1 Data

In Stata:

You can either use level 1 data file or (if you already have it) a single file for two levels. First sort this file by school id:

```
. use "C:\Documents and Settings\SARKISIN\My Documents\hsb1.dta", clear
. sort id
```

Then use egen command to generate an aggregated variable:

```
. by id: egen meanses2=mean(ses)
```

If you have two separate files, you'll end up generating this variable in level 1 file, and then you'll have to create a combined file for two levels by merging the two files (both files have to be sorted by id!):

```
. sort id

. save "C:\Documents and Settings\SARKISIN\My Documents\hsb1.dta", replace
file C:\Documents and Settings\SARKISIN\My Documents\hsb1.dta saved

. use "C:\Documents and Settings\SARKISIN\My Documents\hsb2.dta", clear
. sort id

. save "C:\Documents and Settings\SARKISIN\My Documents\hsb2.dta", replace
file C:\Documents and Settings\SARKISIN\My Documents\hsb2.dta saved

. use "C:\Documents and Settings\SARKISIN\My Documents\hsb1.dta", clear

. merge id using "C:\Documents and Settings\SARKISIN\My Documents\hsb2.dta"
variable id does not uniquely identify observations in the master data
```

```
. tab _merge
```

_merge	Freq.	Percent	Cum.
3	7,185	100.00	100.00
Total	7,185	100.00	

```
. drop _merge
```

In SPSS:

Here you can use command Aggregate to generate a new aggregated variable in level-1 file or a merged file. If you have two separate files, you'll end up generating that variable in level 1 file, and you'll have to transfer it into level 2 (utilizing Merge function in SPSS).

```
AGGREGATE
 / BREAK = id
 / meanses2 = MEAN(ses).
```

Then merge:

```
MATCH FILES /FILE=*
 /TABLE='C:\ HSB2.SAV'
 /BY id.
EXECUTE.
```

Missing Data Note

HLM handles missing data at level 1 by deleting observations with missing data using listwise deletion at either the MDM creation stage or when the analysis is run (you can choose between these two options when creating the file). If deletion at the MDM creation stage is chosen, listwise deletion is performed based on the all level-1 variables selected for inclusion in the MDM file. If deletion at the analysis stage is chosen, listwise deletion is performed based on the variables included in the actual model to be run. Both types of deletion can be problematic if the data are not Missing Completely at Random (MCAR) – i.e. if the pattern of missing values depends on the value of either the dependent or the independent variables. If you have missing data and you suspect they are not MCAR (which is often the case), you should deal with them in another statistical program before transferring your data to HLM.

One of the best choices available if the data are not MCAR is multiple imputation. This procedure produces *M* "complete" data sets, and you can then apply HLM to these multiply-imputed data to produce appropriate estimates that incorporate the uncertainty resulting from imputation. The options for that can be found under Estimation Settings (Plausible Values or Multiple Imputation).

For level 2 (and level 3 if present), HLM assumes data files to be complete. If any of the higher level variables contain missing data, units with missing data will automatically be deleted when the MDM file is created. Again, you should deal with missing data before you start working with HLM. Also, note that if your input data are ASCII and your missing data are coded, for example, as 99, HLM will treat such missing data codes as valid data values, so you should be careful.