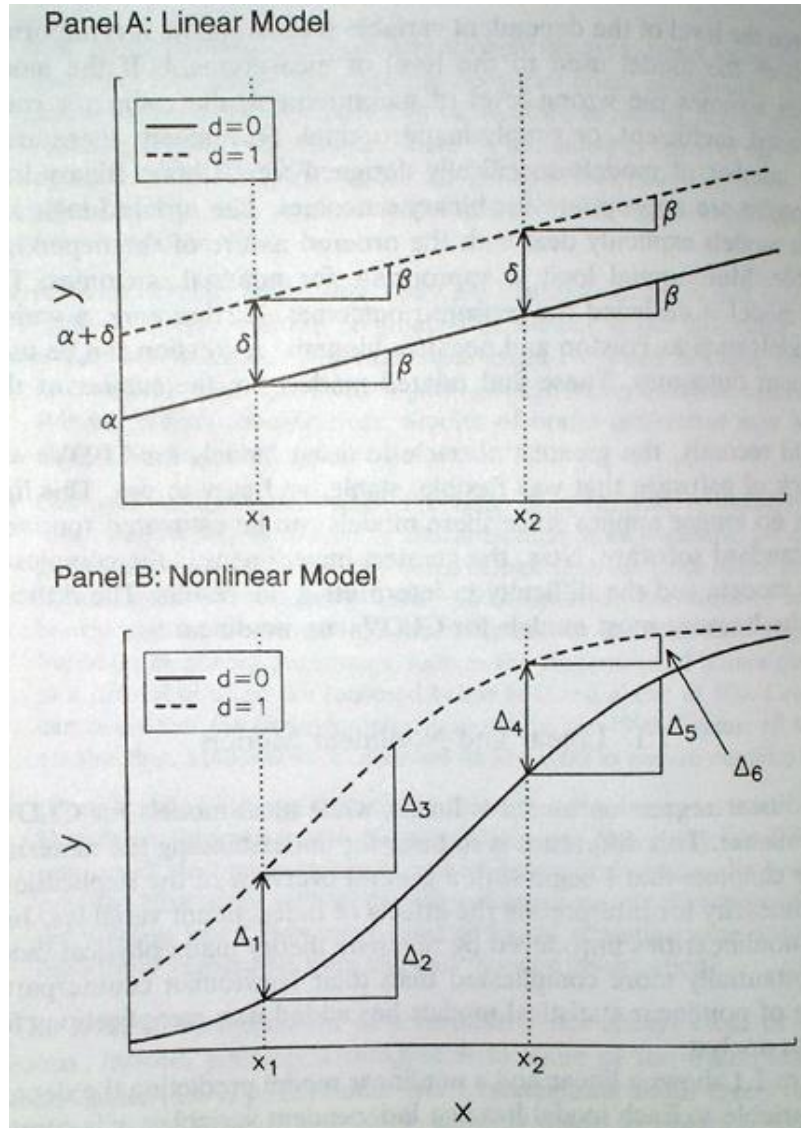


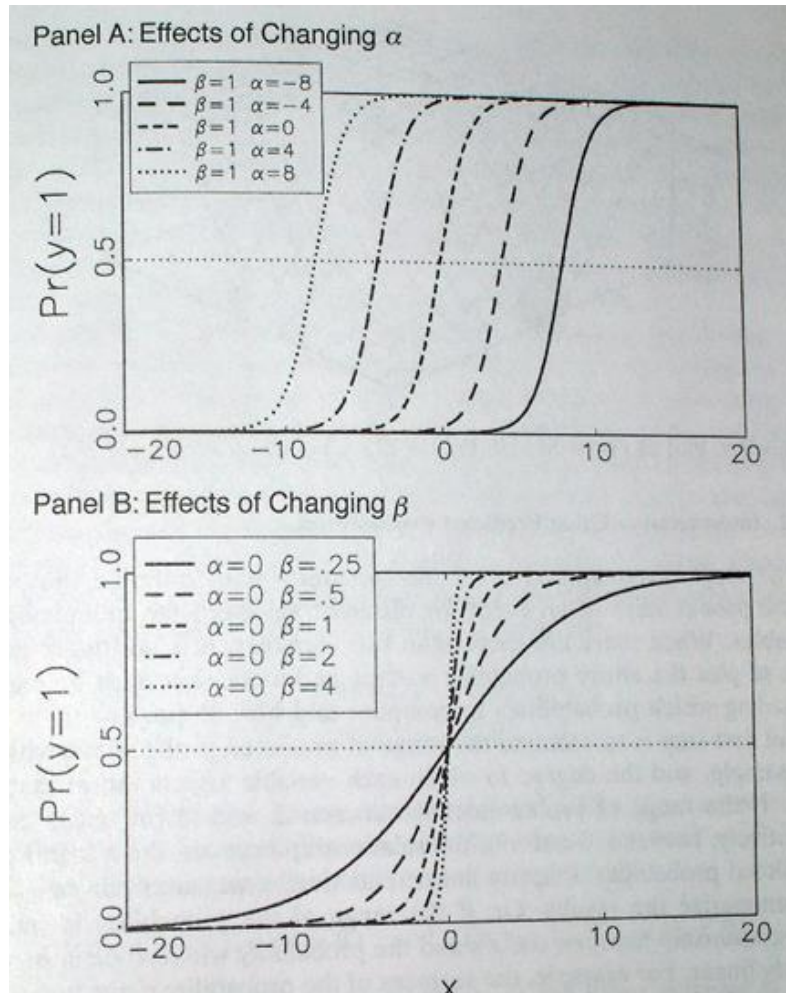
**Sociology 7704: Regression Models for Categorical Data**  
**Instructor: Natasha Sarkisian**

**Binary Logit: Interpretation**

As logistic regression models (whether binary, ordered, or multinomial) are nonlinear, they pose a challenge for interpretation. The increase in the dependent variable in a linear model is constant for all values of  $X$ . Not so for logit models – probability increases or decreases per unit change in  $X$  is nonconstant, as illustrated in this picture.



When interpreting logit regression coefficients, we can interpret only the sign and significance of the coefficients – cannot interpret the size. The following picture can give you an idea how the shape of the curve varies depending on the size of the coefficient, however. Note that, similarly to OLS regression, the constant determines the position of the curve along the  $X$  axis and the coefficient (beta) determines the slope.



Next, we'll examine various ways to interpret logistic regression results.

### 1. Coefficients and Odds Ratios

We'll use another model, focusing now on the probability of voting.

```
. codebook vote00
```

```
vote00
```

```
did r vote in 2000 election
```

```
-----
              type:  numeric (byte)
              label:  vote00

              range:  [1,4]
unique values:  4                               units:  1
                                              missing .:  14/2765
```

```
tabulation:  Freq.    Numeric  Label
              1780         1    voted
              822         2    did not vote
              138         3    ineligible
               11         4    refused to answer
               14          .
```

```
. gen vote=(vote00==1) if vote00<3
(163 missing values generated)
. gen married=(marital==1)
```

```
. logit vote age sex born married child educ
Iteration 0:   log likelihood = -1616.8899
Iteration 1:   log likelihood = -1365.9814
Iteration 2:   log likelihood = -1353.4091
Iteration 3:   log likelihood = -1353.2224
Iteration 4:   log likelihood = -1353.2224
Logistic regression
```

Number of obs	=	2590
LR chi2(6)	=	527.33
Prob > chi2	=	0.0000
Pseudo R2	=	0.1631

Log likelihood = -1353.2224

vote	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0466321	.003337	13.97	0.000	.0400917	.0531726
sex	.1094233	.09552	1.15	0.252	-.0777924	.296639
born	-.9673683	.1859278	-5.20	0.000	-1.33178	-.6029564
married	.4911099	.0983711	4.99	0.000	.2983062	.6839136
child	-.0391447	.0327343	-1.20	0.232	-.1033028	.0250133
educ	.2862839	.0197681	14.48	0.000	.2475391	.3250287
_cons	-4.352327	.3892601	-11.18	0.000	-5.115263	-3.589391

These are regular logit coefficients; so we can interpret the sign and significance but not the size of effects. So we can say that age increases the probability of voting but we can't say by how much – that's because a 1 year increase in age will not affect the probability the same way for a 30 year old and for a 40 year old.

To be able to interpret effect size, we turn to odds ratios. Note that odds ratios are only appropriate for logistic regression – they don't work for probit models.

Odds are ratios of two probabilities – probability of a positive outcome and a probability of a negative outcome (e.g. probability of voting divided by a probability of not voting). But since probabilities vary depending on values of X, such a ratio varies as well. What remains constant is the ratio of such odds – e.g. odds of voting for women divided by odds of voting for men will be the same number regardless of the values of other variables. Similarly, the odds ratio for age can be a ratio of the odds of voting for someone who is 31 y.o. to the odds of a 30 y.o. person, or of a 41 y.o. to a 40 y.o. person's odds – these will be the same regardless of what age values you pick, as long as they are one year apart. So let's examine the odds ratios.

```
. logit vote age sex born married child educ, or
Iteration 0:   log likelihood = -1616.8899
Iteration 1:   log likelihood = -1365.9814
Iteration 2:   log likelihood = -1353.4091
Iteration 3:   log likelihood = -1353.2224
Iteration 4:   log likelihood = -1353.2224
Logistic regression
```

Number of obs	=	2590
LR chi2(6)	=	527.33
Prob > chi2	=	0.0000
Pseudo R2	=	0.1631

Log likelihood = -1353.2224

vote	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
age	1.047736	.0034963	13.97	0.000	1.040906	1.054612
sex	1.115634	.1065654	1.15	0.252	.9251564	1.34533
born	.380082	.0706678	-5.20	0.000	.2640069	.5471915
married	1.634129	.160751	4.99	0.000	1.347574	1.981618
child	.9616115	.0314777	-1.20	0.232	.9018538	1.025329
educ	1.33147	.0263207	14.48	0.000	1.280869	1.38407

Another way to obtain odds ratios would be to use “logistic” command instead of “logit” – it automatically displays odds ratios instead of coefficients. But yet another, more convenient way is to use listcoef command (that’s one of the commands written by Scott Long that we downloaded as a part of spost package):

```
. listcoef
logit (N=2590): Factor change in odds
Odds of: 1 vs 0
```

	b	z	P> z	e^b	e^bStdX	SDofX
age	0.0466	13.974	0.000	1.048	2.230	17.195
sex	0.1094	1.146	0.252	1.116	1.056	0.497
born	-0.9674	-5.203	0.000	0.380	0.788	0.246
married	0.4911	4.992	0.000	1.634	1.278	0.499
childs	-0.0391	-1.196	0.232	0.962	0.936	1.676
educ	0.2863	14.482	0.000	1.331	2.311	2.926
constant	-4.3523	-11.181	0.000	.	.	.

The advantage of listcoef is that it reports regular coefficients, odds ratios, and standardized odds ratios in one table. Odds ratios are exponentiated logistic regression coefficients. They are sometimes called factor coefficients, because they are multiplicative coefficients. Odds ratios are equal to 1 if there is no effect, smaller than 1 if the effect is negative and larger than 1 if it is positive. So for example, the odds ratio for married indicates that the odds of voting for those who are married are 1.63 times higher than for those who are not married. And the odds ratio for education indicates that each additional year of education makes one’s odds of voting 1.33 times higher -- or, in other words, increases those odds by 33%. To get percent change directly, we can use percent option:

```
. listcoef, percent
logit (N=2590): Percentage Change in Odds
Odds of: 1 vs 0
```

vote	b	z	P> z	%	%StdX	SDofX
age	0.04663	13.974	0.000	4.8	123.0	17.1953
sex	0.10942	1.146	0.252	11.6	5.6	0.4972
born	-0.96737	-5.203	0.000	-62.0	-21.2	0.2457
married	0.49111	4.992	0.000	63.4	27.8	0.4990
childs	-0.03914	-1.196	0.232	-3.8	-6.4	1.6762
educ	0.28628	14.482	0.000	33.1	131.1	2.9257

Beware: if you would like to know what the increase would be per, say, 10 units increase in the independent variable – e.g. 10 years of education, you cannot simply multiple the odds ratio by 10! The coefficient, in fact, would be odds ratio to the power of 10. Or alternatively, you could take the regular logit coefficient, multiply it by 10 and then exponentiate it -- e.g., for education:

```
. di exp(0.28628*10)
17.510488
. di 1.3315^10
17.515063
```

Standardized odds ratios (presented under e^bStdX) are similar to regular odds ratios, but they display the change in the odds of voting per one standard deviation change in the independent variable. The last column in the table generated by listcoef shows what one standard deviation for each variable is. So for age the standardized odds ratio indicates that 17 years of age increase one’s odds of voting 2.23 times, or by 123%. Standardized odds ratios, like standardized coefficients in OLS, allow us to compare effect sizes across variables regardless of their measurement units. But,

beware of comparing negative and positive effects – odds ratios of 1.5 and .5 are not equivalent, even though the first one represents a 50% increase in odds and the second one represents a 50% decrease. This is because odds ratios cannot be below zero (there cannot be a decrease more than 100%), but they do not have an upper bound – i.e. can be infinitely high. In order to be able to compare positive and negative effects, we can reverse odds ratios and generate odds ratios for odds of not voting (rather than odds of voting).

```
. listcoef, reverse
logit (N=2590): Factor Change in Odds
Odds of: 0 vs 1
```

vote	b	z	P> z	e^b	e^bStdX	SDofX
age	0.04663	13.974	0.000	0.9544	0.4485	17.1953
sex	0.10942	1.146	0.252	0.8964	0.9470	0.4972
born	-0.96737	-5.203	0.000	2.6310	1.2682	0.2457
married	0.49111	4.992	0.000	0.6119	0.7826	0.4990
childs	-0.03914	-1.196	0.232	1.0399	1.0678	1.6762
educ	0.28628	14.482	0.000	0.7510	0.4328	2.9257

We can see for example that the odds ratio of 0.3801 for born is a negative effect corresponding in size to a positive odds ratio of 2.6310. Listcoef also has a help option that explains what's what :

```
. listcoef, reverse help
logit (N=2590): Factor Change in Odds
Odds of: 0 vs 1
```

vote	b	z	P> z	e^b	e^bStdX	SDofX
age	0.04663	13.974	0.000	0.9544	0.4485	17.1953
sex	0.10942	1.146	0.252	0.8964	0.9470	0.4972
born	-0.96737	-5.203	0.000	2.6310	1.2682	0.2457
married	0.49111	4.992	0.000	0.6119	0.7826	0.4990
childs	-0.03914	-1.196	0.232	1.0399	1.0678	1.6762
educ	0.28628	14.482	0.000	0.7510	0.4328	2.9257

```

b = raw coefficient
z = z-score for test of b=0
P>|z| = p-value for z-test
e^b = exp(b) = factor change in odds for unit increase in X
e^bStdX = exp(b*SD of X) = change in odds for SD increase in X
SDofX = standard deviation of X

```

When a set of dummies is used, we might be interested in all kinds of pairwise comparisons; to get odds ratios for those, we use pwcompare command:

```
. logit vote age sex born i.marital childs educ, or
Iteration 0: log likelihood = -1616.8899
Iteration 1: log likelihood = -1361.6039
Iteration 2: log likelihood = -1352.4837
Iteration 3: log likelihood = -1352.4548
Iteration 4: log likelihood = -1352.4548
Logistic regression
```

vote	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
age	1.048782	.0040525	12.33	0.000	1.040869 1.056755
sex	1.11771	.1080131	1.15	0.250	.924849 1.350789
born	.3761262	.0701482	-5.24	0.000	.2609655 .5421061
marital					

```

Number of obs = 2590
LR chi2(9) = 528.87
Prob > chi2 = 0.0000
Pseudo R2 = 0.1635
Log likelihood = -1352.4548

```

```

widowed | .6014296 .125745 -2.43 0.015 .3992255 .9060482
divorced | .5493787 .0741513 -4.44 0.000 .4216796 .7157496
separated | .6970315 .1716079 -1.47 0.143 .4302175 1.129319
never married | .6503118 .0840993 -3.33 0.001 .5047112 .8379156
|
childs | .9655952 .0325389 -1.04 0.299 .9038806 1.031523
educ | 1.333732 .0265289 14.48 0.000 1.282737 1.386754
_cons | .0196952 .0081736 -9.46 0.000 .0087319 .0444234
-----

```

```

. pwcompare marital
Pairwise comparisons of marginal linear predictions
Margins : asbalanced

```

```

-----
| Contrast Std. Err. Unadjusted [95% Conf. Interval]
-----+-----
vote |
marital |
widowed vs married | -.5084458 .2090768 -.9182288 -.0986628
divorced vs married | -.5989672 .1349731 -.8635096 -.3344249
separated vs married | -.3609247 .2461983 -.8434645 .121615
never married vs married | -.4303034 .1293215 -.6837689 -.1768379
divorced vs widowed | -.0905214 .2213725 -.5244036 .3433607
separated vs widowed | .1475211 .3044299 -.4491506 .7441927
never married vs widowed | .0781424 .2412223 -.3946447 .5509295
separated vs divorced | .2380425 .2618905 -.2752534 .7513384
never married vs divorced | .1686638 .1560947 -.1372761 .4746038
never married vs separated | -.0693787 .2594929 -.5779754 .4392181
-----

```

And to get actual odds ratios:

```

. pwcompare marital, eform
Pairwise comparisons of marginal linear predictions
Margins : asbalanced

```

```

-----
| exp(b) Std. Err. Unadjusted [95% Conf. Interval]
-----+-----
vote |
marital |
widowed vs married | .6014296 .125745 .3992255 .9060482
divorced vs married | .5493787 .0741513 .4216796 .7157496
separated vs married | .6970315 .1716079 .4302175 1.129319
never married vs married | .6503118 .0840993 .5047112 .8379156
divorced vs widowed | .9134548 .2022138 .5919083 1.409677
separated vs widowed | 1.158958 .3528214 .63817 2.104742
never married vs widowed | 1.081277 .2608281 .6739194 1.734865
separated vs divorced | 1.268763 .332277 .7593797 2.119835
never married vs divorced | 1.183722 .1847727 .8717295 1.607377
never married vs separated | .9329733 .2421 .5610331 1.551494
-----

```

A side note: something that can be helpful when doing hypothesis testing for groups of dummies (instead of using acc option in test or lrtest):

```

. testparm i.marital
( 1) [vote]2.marital = 0
( 2) [vote]3.marital = 0
( 3) [vote]4.marital = 0
( 4) [vote]5.marital = 0

      chi2( 4) =    26.50
      Prob > chi2 =    0.0000

```

## 2. Predicted Probabilities

In addition to regular coefficients and odds ratios, we also should examine predicted probabilities – both for the actual observations in our data and for strategically selected hypothetical cases.

Predicted probabilities are always calculated for a specific set of independent variables' values.

One thing we can calculate is predicted probabilities for the actual data that we have – for each case, we take the values of all independent variables and plug it into the equation:

```
. predict prob
(option p assumed; Pr(vote))
(26 missing values generated)

. sum prob if e(sample)
Variable |      Obs      Mean   Std. Dev.      Min      Max
-----+-----
prob |      2590   .6833977   .204702   .0205784   .9926677
```

Mean of predicted probabilities represents the average proportion in the sample:

```
. sum vote if e(sample)
Variable |      Obs      Mean   Std. Dev.      Min      Max
-----+-----
vote |      2590   .6833977   .4652406         0         1
```

These are predicted probabilities for the actual cases in our dataset. It can be useful, however, to calculate predicted probabilities for hypothetical sets of values – some interesting combinations that we could compare and contrast.

```
. margins, atmeans
```

```
Adjusted predictions      Number of obs   =      2590
Model VCE      : OIM

Expression      : Pr(vote), predict()
at
   age          =    46.93591 (mean)
   sex          =    1.553282 (mean)
   born         =    1.064479 (mean)
   married      =    .4675676 (mean)
   child5       =    1.838996 (mean)
   educ         =   13.39459 (mean)

-----+-----
          |      Delta-method
          |      Margin   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
   _cons |   .7249026   .0100274    72.29   0.000   .7052494   .7445559
-----+-----
```

This calculates a predicted probability for a case with all values set at the mean. So an “average” person has 72.5% chance of voting. We can also see what these averages are. If we do not specify atmeans (and do not specify values for each variable), the margins command calculates average predicted probability across the observations we have in the dataset.

Clearly, for some variables, averages don't make sense – e.g., we don't want to use averages for dummy variables; rather, we'd want to specify what values to use. Here is an example of specifying values:

```
. margins, at(age=30 born=1 sex=2 married=0) atmeans
```

```
Adjusted predictions      Number of obs   =      2590
Model VCE      : OIM
```

```

Expression : Pr(vote), predict()
at          : age          =          30
              sex          =           2
              born         =           1
              married       =           0
              childss       =    1.838996 (mean)
              educ          =    13.39459 (mean)

```

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]	
_cons	.5151914	.0219222	23.50	0.000	.4722246	.5581581

This is the predicted value for someone who is 30, native born, female, and unmarried (and has average number of children and average education). Note that if you have a set of dummy variables, you can just specify the category number, e.g., if you are using i.marital, you can write (marital=2) in the at option.

We can also use margins command to compare predictions at different values:

```

. margins, at(married=0 married=1) atmeans
Adjusted predictions              Number of obs   =          2590

```

```

Model VCE      : OIM
Expression     : Pr(vote), predict()
1._at          : age          =    46.93591 (mean)
                  sex          =    1.553282 (mean)
                  born         =    1.064479 (mean)
                  married       =           0
                  childss       =    1.838996 (mean)
                  educ          =    13.39459 (mean)

```

```

2._at          : age          =    46.93591 (mean)
                  sex          =    1.553282 (mean)
                  born         =    1.064479 (mean)
                  married       =           1
                  childss       =    1.838996 (mean)
                  educ          =    13.39459 (mean)

```

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]	
_at						
1	.6768395	.0143948	47.02	0.000	.6486262	.7050528
2	.7738877	.0131271	58.95	0.000	.748159	.7996164

```

. margins, at(age=(30(10)70)) atmeans

```

```

Adjusted predictions              Number of obs   =          2590
Model VCE      : OIM

```

```

Expression     : Pr(vote), predict()

```

```

1._at          : age          =          30
                  sex          =    1.553282 (mean)
                  born         =    1.064479 (mean)
                  married       =    .4675676 (mean)
                  childss       =    1.838996 (mean)
                  educ          =    13.39459 (mean)

```



```

2._at      : age          =          40
              sex          =    1.553282 (mean)
              born         =    1.064479 (mean)
              married      =    .4675676 (mean)
              childss      =    1.838996 (mean)
              educ         =    13.39459 (mean)

3._at      : age          =          50
              sex          =    1.553282 (mean)
              born         =    1.064479 (mean)
              married      =    .4675676 (mean)
              childss      =    1.838996 (mean)
              educ         =    13.39459 (mean)

4._at      : age          =          60
              sex          =    1.553282 (mean)
              born         =    1.064479 (mean)
              married      =    .4675676 (mean)
              childss      =    1.838996 (mean)
              educ         =    13.39459 (mean)

5._at      : age          =          70
              sex          =    1.553282 (mean)
              born         =    1.064479 (mean)
              married      =    .4675676 (mean)
              childss      =    1.838996 (mean)
              educ         =    13.39459 (mean)

```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
_at						
1		.5446694	.0160415	33.95	0.000	.5132286 .5761101
2		.6559903	.0111333	58.92	0.000	.6341694 .6778113
3		.752464	.01005	74.87	0.000	.7327664 .7721617
4		.8289379	.0106262	78.01	0.000	.8081108 .8497649
5		.8853845	.0104219	84.95	0.000	.864958 .9058111

### To have a more compact legend:

```
. margins, at(age=(30(10)70) married=(0 1)) atmeans noatlegend
```

```
Adjusted predictions      Number of obs      =      2590
Model VCE      : OIM
```

```
Expression      : Pr(vote), predict()
```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
_at						
1		.4873847	.0196449	24.81	0.000	.4488815 .525888
2		.6084111	.0200184	30.39	0.000	.5691757 .6476464
3		.6024896	.0157359	38.29	0.000	.5716478 .6333313
4		.7123775	.0151151	47.13	0.000	.6827525 .7420025
5		.7072717	.0141615	49.94	0.000	.6795157 .7350278
6		.7979096	.0125434	63.61	0.000	.773325 .8224942
7		.7938829	.0139495	56.91	0.000	.7665424 .8212234
8		.8629015	.0111527	77.37	0.000	.8410427 .8847604
9		.8599425	.0132394	64.95	0.000	.8339938 .8858911

```

10 | .9093663 .0097348 93.41 0.000 .8902865 .9284462
-----

```

```
. mlistat
```

```
at() values held constant
```

```

sex      born      childs      educ
-----
1.55     1.06     1.84     13.4

```

```
at() values vary
```

```

_at |      age      married
-----+-----
1 |      30          0
2 |      30          1
3 |      40          0
4 |      40          1
5 |      50          0
6 |      50          1
7 |      60          0
8 |      60          1
9 |      70          0
10 |     70          1

```

We could also separate groups and do predictions separately (note that group-based means are used for each group, so it is different from using that variable within “at” option).

```
. margins, over(married) at(age=(30(10)70) ) atmeans noatlegend
```

```

Adjusted predictions      Number of obs      =      2590
Model VCE      : OIM

```

```

Expression      : Pr(vote), predict()
over            : married

```

```

-----
|              Delta-method
|              Margin      Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
_at#married |
1 0 |      .4787915      .0187124     25.59   0.000      .4421158      .5154673
1 1 |      .6177066      .0203227     30.39   0.000      .5778749      .6575383
2 0 |      .5942195      .0151977     39.10   0.000      .5644325      .6240064
2 1 |      .7203395      .0149981     48.03   0.000      .6909437      .7497353
3 0 |      .7000965      .0141623     49.43   0.000      .672339      .727854
3 1 |      .8041548      .0121038     66.44   0.000      .7804318      .8278778
4 0 |      .7881948      .0143163     55.06   0.000      .7601354      .8162543
4 1 |      .8674719      .0105976     81.86   0.000      .846701      .8882428
5 0 |      .8557462      .0137381     62.29   0.000      .82882      .8826724
5 1 |      .9125447      .0092091     99.09   0.000      .8944952      .9305942
-----

```

```
. mlistat
```

```
at() values vary
```

```

_at |      age      sex      born      married      childs      educ
-----+-----
1 |      30     1.59     1.05          0     1.53     13.2
2 |      30     1.51     1.08          1     2.2     13.7
3 |      40     1.59     1.05          0     1.53     13.2

```

4	40	1.51	1.08	1	2.2	13.7
5	50	1.59	1.05	0	1.53	13.2
6	50	1.51	1.08	1	2.2	13.7
7	60	1.59	1.05	0	1.53	13.2
8	60	1.51	1.08	1	2.2	13.7
9	70	1.59	1.05	0	1.53	13.2
10	70	1.51	1.08	1	2.2	13.7

Margins command also permits us to transform our predictions and get p-values and CI for transformed version:

```
. margins, at(married=(0 1)) atmeans noatlegend expression(1-predict(pr))
```

```
Adjusted predictions      Number of obs   =      2590
Model VCE      : OIM
```

```
Expression      : 1-predict(pr)
```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
	_at					
	1	.3231605	.0143948	22.45	0.000	.2949472 .3513738
	2	.2261123	.0131271	17.22	0.000	.2003836 .251841

Or to test if predicted probability is different from, say, 0.5:

```
. margins, at(married=(0 1)) atmeans noatlegend expression(predict(pr)-.5)
```

```
Adjusted predictions      Number of obs   =      2590
Model VCE      : OIM
```

```
Expression      : predict(pr)-.5
```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
	_at					
	1	.1768395	.0143948	12.28	0.000	.1486262 .2050528
	2	.2738877	.0131271	20.86	0.000	.248159 .2996164

We can also use mtable to obtain values of predicted probabilities for various combinations of categorical variables – but note that we need to specify what values to use for all other variables – e.g., in this case, all other variables are set at the mean.

```
. qui logit vote age sex born married child educ
. mtable, at(born=(0 1) married=(0 1)) atmeans
```

```
Expression: Pr(vote), predict()
```

		born	married	Pr(y)
	1	0	0	0.854
	2	0	1	0.906
	3	1	0	0.690
	4	1	1	0.785

Specified values of covariates

	age	sex	childs	educ
Current	46.9	1.55	1.84	13.4

This allows us to see that the effect of one variable depends on the level of the other – for native born individuals, marriage increases chances of voting by 9.5%, but for the foreign born, marriage increases these chances by 12.2%. We can also get confidence intervals for predictions, as well as some other statistics:

```
. mtable, at(born=(0 1) married=(0 1)) atmeans statistics(ci)
```

Expression: Pr(vote), predict()

	born	married	Pr(y)	ll	ul
1	0	0	0.854	0.804	0.905
2	0	1	0.906	0.869	0.942
3	1	0	0.690	0.662	0.718
4	1	1	0.785	0.759	0.810

Specified values of covariates

	age	sex	childs	educ
Current	46.9	1.55	1.84	13.4

```
. mtable, at(born=(0 1) married=(0 1)) atmeans statistics(all)
```

Expression: Pr(vote), predict()

	born	married	Pr(y)	se	z	p
1	0	0	0.854	0.026	33.196	0.000
2	0	1	0.906	0.019	48.426	0.000
3	1	0	0.690	0.014	48.515	0.000
4	1	1	0.785	0.013	60.182	0.000

	ll	ul
1	0.804	0.905
2	0.869	0.942
3	0.662	0.718
4	0.759	0.810

Specified values of covariates

	age	sex	childs	educ
Current	46.9	1.55	1.84	13.4

You may also find an older command, prtab, useful (but note that it is not compatible with the new way to specifying dummies using i. – only works with xi: prefix in that case):

```
. prtab born married, rest(mean)
```

logit: Predicted probabilities of positive outcome for vote

```
-----
was r      |
born in    |
this       | married
country    |      0      1
```

```

-----+-----
      yes | 0.6903  0.7846
      no | 0.4587  0.5806
-----+-----
      age      sex      born      married      childs      educ
x=  46.935907  1.5532819  1.0644788  .46756757  1.8389961  13.394595

```

With `mtable`, the best way to do predictions by group is to use `over` option:

```
. mtable, at(born=(0 1) married=(0 1)) atmeans over(sex)
```

Expression: `Pr(vote), predict()`

		age	sex	born	married	childs	educ
1		46.2	1	0	0	1.68	13.4
2		47.5	2	0	0	1.96	13.4
3		46.2	1	0	1	1.68	13.4
4		47.5	2	0	1	1.96	13.4
5		46.2	1	1	0	1.68	13.4
6		47.5	2	1	0	1.96	13.4
7		46.2	1	1	1	1.68	13.4
8		47.5	2	1	1	1.96	13.4

		Pr(y)
1		0.843
2		0.863
3		0.898
4		0.911
5		0.672
6		0.705
7		0.770
8		0.796

Specified values where `.n` indicates no values specified with `at()`

		No at()
Current		.n

Note that it only makes sense to create such tables of predicted probabilities for variables that have significant effects – otherwise, you’ll see no differences.

Further, we can use `marginsplot` after `margin` to graph probabilities for certain sets of values. This is useful with continuous variables, as it allows us to see how predicted probability changes across values of one variable (given that the rest of them are set at some specific values).

For example, we can plot four curves that show how probability of voting changes by age for an average person who has 10, 12, 16, or 20 years of education.

```
. margins, at(age=(20(10)80) educ=(10 12 16 20)) atmeans noatlegend
```

```
Adjusted predictions      Number of obs   =      2590
Model VCE      : OIM
```

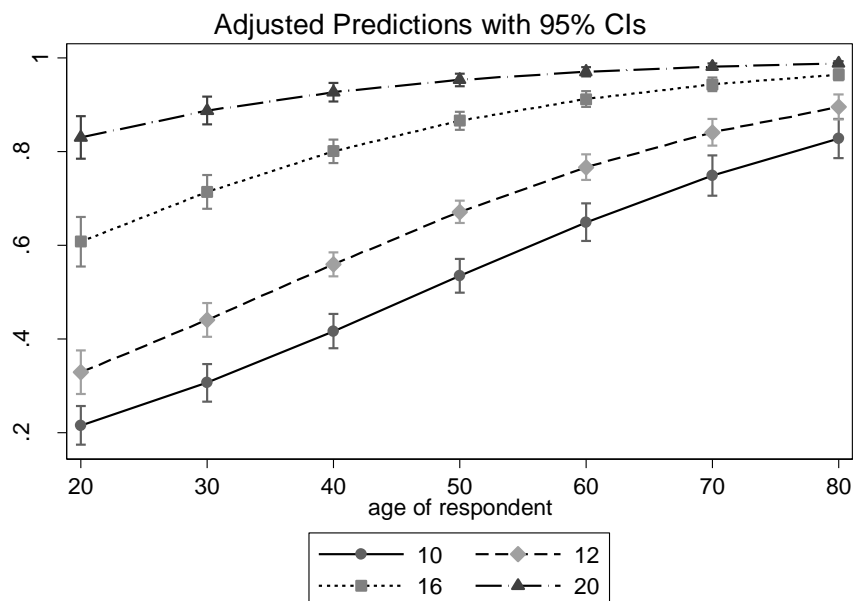
Expression : `Pr(vote), predict()`

		Delta-method			
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]

	_at					
1		.2160915	.0211483	10.22	0.000	.1746416
2		.3290183	.0237367	13.86	0.000	.2824951
3		.6080911	.0271769	22.38	0.000	.5548254
4		.8307875	.0231461	35.89	0.000	.785422
5		.3074013	.0204778	15.01	0.000	.2672656
6		.4411898	.0186146	23.70	0.000	.4047058
7		.7141425	.0183676	38.88	0.000	.6781426
8		.8877053	.0151673	58.53	0.000	.8579778
9		.4167808	.0186739	22.32	0.000	.3801807
10		.5597036	.0131027	42.72	0.000	.5340229
11		.8008927	.0125282	63.93	0.000	.7763378
12		.9271563	.010058	92.18	0.000	.9074431
13		.5350154	.0185145	28.90	0.000	.4987277
14		.6717814	.0119539	56.20	0.000	.6483522
15		.8662472	.0098556	87.89	0.000	.8469306
16		.953474	.0068999	138.19	0.000	.9399504
17		.6494415	.020568	31.58	0.000	.6091288
18		.7671963	.0138781	55.28	0.000	.7399956
19		.9124937	.0084824	107.58	0.000	.8958686
20		.970585	.004878	198.97	0.000	.9610244
21		.7489234	.0220661	33.94	0.000	.7056747
22		.8414212	.0146909	57.27	0.000	.8126275
23		.9437876	.0071808	131.43	0.000	.9297136
24		.981525	.0034965	280.72	0.000	.974672
25		.8276656	.0213316	38.80	0.000	.7858566
26		.8952132	.0136561	65.55	0.000	.8684477
27		.9643278	.0057912	166.52	0.000	.9529772
28		.9884446	.0025063	394.38	0.000	.9835323

```
. marginsplot
```

Variables that uniquely identify margins: age educ



If there are interactions or nonlinearities that required that you entered a variable more than once (e.g. X and X squared), you can also this marginplots to graph that.

```
. logit vote i.sex##c.age educ i.born i.marital child
```

```
Iteration 0:  log likelihood = -1616.8899
Iteration 1:  log likelihood = -1361.3117
Iteration 2:  log likelihood = -1352.2041
Iteration 3:  log likelihood = -1352.1752
Iteration 4:  log likelihood = -1352.1752
```

```
Logistic regression                                Number of obs   =       2590
                                                    LR chi2(10)    =       529.43
                                                    Prob > chi2    =       0.0000
Log likelihood = -1352.1752                      Pseudo R2      =       0.1637
```

vote	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
sex						
female	-.0844048	.2788219	-0.30	0.762	-.6308857	.4620761
age	.0451964	.0050282	8.99	0.000	.0353414	.0550514
sex#c.age						
female	.0045923	.006136	0.75	0.454	-.007434	.0166185
educ	.2877763	.0198892	14.47	0.000	.2487942	.3267584
born						
no	-.9707724	.1867578	-5.20	0.000	-1.336811	-.6047339
marital						
widowed	-.5480377	.2157987	-2.54	0.011	-.9709953	-.1250801
divorced	-.6021702	.13507	-4.46	0.000	-.8669025	-.3374379
separated	-.3569101	.2463735	-1.45	0.147	-.8397932	.125973
never mar..	-.4341406	.1294304	-3.35	0.001	-.6878196	-.1804616
childs	-.0334493	.0337876	-0.99	0.322	-.0996717	.0327732
_cons	-4.68753	.3754022	-12.49	0.000	-5.423305	-3.951756

```
. margins, at(age=(20(10)80) sex=(1 2 )) atmeans noatlegend
```

```
Adjusted predictions                                Number of obs   =       2590
Model VCE      : OIM
```

```
Expression      : Pr(vote), predict()
```

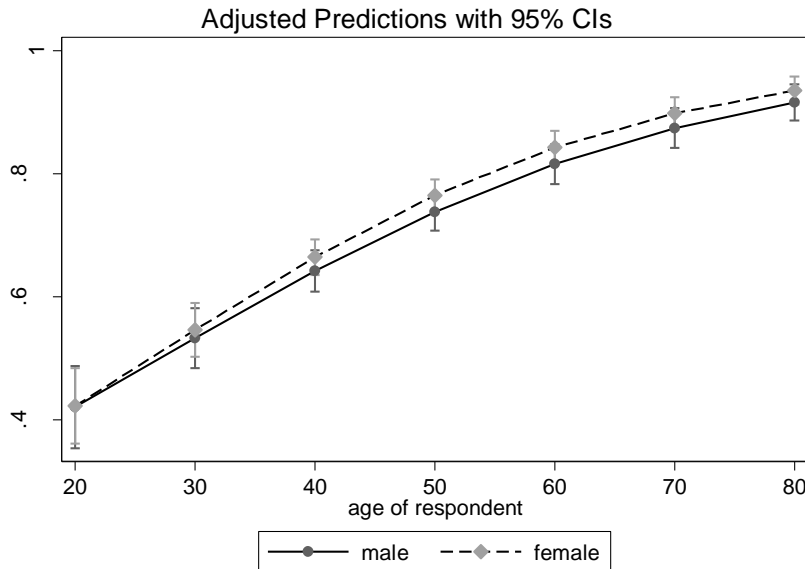
	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]	
_at						
1	.4208618	.0339115	12.41	0.000	.3543965	.4873271
2	.5331333	.0248281	21.47	0.000	.4844712	.5817954
3	.6421463	.0171949	37.35	0.000	.6084449	.6758478
4	.7382043	.0153451	48.11	0.000	.7081285	.7682802
5	.8158712	.016502	49.44	0.000	.7835279	.8482145
6	.8744164	.0166122	52.64	0.000	.8418571	.9069757
7	.9162574	.0151062	60.65	0.000	.8866497	.945865
8	.4226765	.0312803	13.51	0.000	.3613681	.4839848
9	.5463891	.022257	24.55	0.000	.5027662	.5900119
10	.6646261	.0147728	44.99	0.000	.6356719	.6935803
11	.7652831	.0132203	57.89	0.000	.7393717	.7911945
12	.8428718	.0139768	60.31	0.000	.8154779	.8702658

13		.8982235	.0134213	66.93	0.000	.8719183	.9245288
14		.935567	.011543	81.05	0.000	.9129432	.9581908

---

```
. marginsplot
```

Variables that uniquely identify margins: age sex



If you want to be able to format these graphs in your own ways, you can save predictions from margins into variables using mgen command:

```
. mgen, at(educ=(10 12 16 20) sex=(1 2) ) atmeans stub(edsex_)
```

Predictions from: margins, at(educ=(10 12 16 20) sex=(1 2)) atmeans predict(pr)

Variable	Obs	Unique	Mean	Min	Max	Label
edsex_pr1	8	8	.7320941	.4793427	.9470915	pr(y=1) from margins
edsex_ll1	8	8	.699113	.429009	.9269149	95% lower limit
edsex_ul1	8	8	.7650752	.5296764	.9672681	95% upper limit
edsex_educ	8	4	14.5	10	20	highest year of school co...
edsex_sex	8	2	1.5	1	2	respondents sex

---

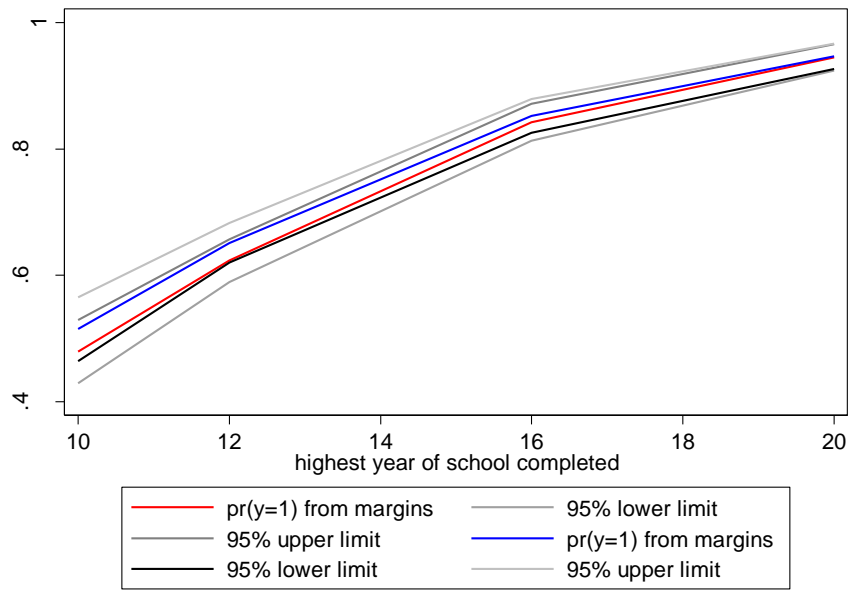
Specified values of covariates

	2.	2.	3.	4.	5.	
age	born	marital	marital	marital	marital	childs
46.93591	.0644788	.0926641	.1617761	.0351351	.2428571	1.838996

---

```
. graph twoway (line edsex_pr1 edsex_ll1 edsex_ul1 edsex_educ if edsex_sex==1,
> sort lcolor(red)) (line edsex_pr1 edsex_ll1 edsex_ul1 edsex_educ if edsex
> _sex==2, sort lcolor(blue))
```

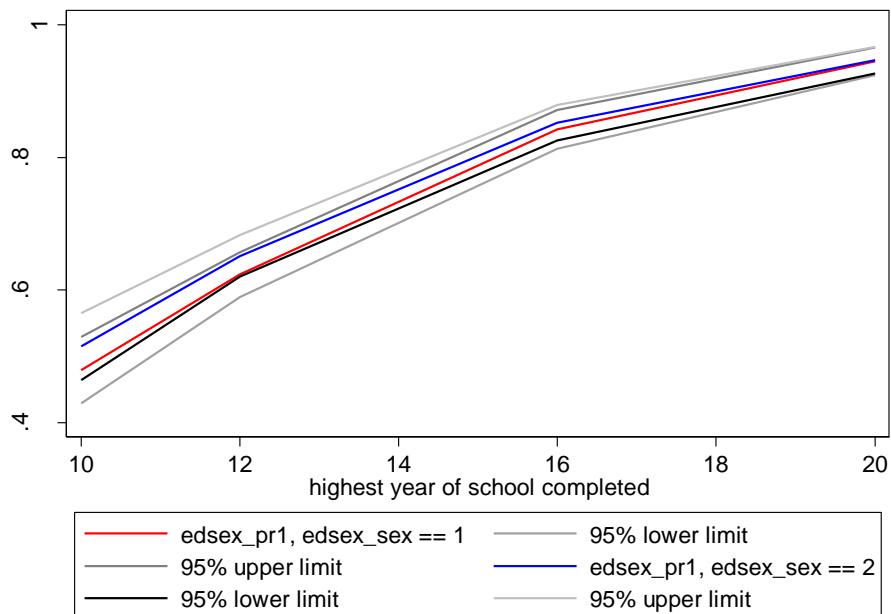




```
. separate edsex_pr1, by(edsex_sex)
```

variable name	storage type	display format	value label	variable label
edsex_pr11	float	%9.0g		edsex_pr1, edsex_sex == 1
edsex_pr12	float	%9.0g		edsex_pr1, edsex_sex == 2

```
. graph twoway (line edsex_pr11 edsex_ll1 edsex_ul1 edsex_educ if edsex_sex==1
> , sort lcolor(red)) (line edsex_pr12 edsex_ll1 edsex_ul1 edsex_educ if eds
> ex_sex==2, sort lcolor(blue))
```



### 3. Changes in Predicted Probabilities

Another way to interpret logistic regression results is using changes in predicted probabilities. These are changes in probability of the outcome as one variable changes, holding all other variables constant at certain values. There are two ways to measure such changes – discrete change and marginal effect.

#### A. Discrete change

Discrete change is a change in predicted probabilities corresponding to a given change in the independent variable. To obtain these, we calculate two probabilities and then calculate the difference between them. For example:

```
. mtable, at(sex=1) atmeans rowname(sex=1) statistics(ci)
```

Expression: Pr(vote), predict()

	Pr (y)	ll	ul
sex=1	0.713	0.684	0.742

Specified values of covariates

	age	sex	2. born	2. marital	3. marital	4. marital
Current	46.9	1	.0645	.0927	.162	.0351

	5. marital	childs	educ
Current	.243	1.84	13.4

```
. mtable, at(sex=2) atmeans rowname(sex=2) statistics(ci) below
```

Expression: Pr(vote), predict()

	Pr (y)	ll	ul
sex=1	0.713	0.684	0.742
sex=2	0.735	0.710	0.761

Specified values of covariates

	age	sex	2. born	2. marital	3. marital	4. marital
Set 1	46.9	1	.0645	.0927	.162	.0351
Current	46.9	2	.0645	.0927	.162	.0351

	5. marital	childs	educ
Set 1	.243	1.84	13.4
Current	.243	1.84	13.4

```
. mtable, dydx(sex) atmeans rowname(sex=2 - sex=1) statistics(ci) below brief
```

Expression: Pr(vote), predict()

	Pr (y)	ll	ul
sex=1	0.713	0.684	0.742

```

      sex=2 |      0.735      0.710      0.761
sex=2 - sex=1 |      0.022     -0.016      0.060

```

We can also calculate a bunch of predictions and then conduct pairwise comparisons and get significance tests for them using `margins` (need `post` option in `margins`):

```

. margins, at(sex=(1 2) marital=(1(1)5)) atmeans post
Expression: Pr(vote), predict()

```

```

      |      sex      marital      Pr(y)
-----+-----
      1 |      1          1      0.763
      2 |      1          2      0.660
      3 |      1          3      0.639
      4 |      1          4      0.692
      5 |      1          5      0.677
      6 |      2          1      0.783
      7 |      2          2      0.684
      8 |      2          3      0.665
      9 |      2          4      0.715
     10 |      2          5      0.701

```

Specified values of covariates

```

      |      2.
      |      age      born      childs      educ
-----+-----
Current |      46.9      .0645      1.84      13.4

```

```

. mat list e(b)

```

```

e(b) [1,10]

```

```

      1.      2.      3.      4.      5.      6.
      _at_ _at_ _at_ _at_ _at_ _at_
y1      .76342597 .65995848 .63936008 .69224379 .67726949 .7829323

      7.      8.      9.      10.
      _at_ _at_ _at_ _at_
y1      .68447002 .66460183 .71543157 .70109834

```

```

. mlincom 1 - 6

```

```

      |      lincom      pvalue      ll      ul
-----+-----
      1 |      -0.020      0.250      -0.053      0.014

```

But there are commands that make it easier to do.

```

. logit vote age i.sex i.born i.marital childs educ

```

```

Iteration 0:   log likelihood = -1616.8899

```

```

Iteration 1:   log likelihood = -1361.6039

```

```

Iteration 2:   log likelihood = -1352.4837

```

```

Iteration 3:   log likelihood = -1352.4548

```

```

Iteration 4:   log likelihood = -1352.4548

```

Logistic regression

Number of obs = 2590

LR chi2(9) = 528.87

Prob > chi2 = 0.0000

Pseudo R2 = 0.1635

Log likelihood = -1352.4548

```

      vote |      Coef.      Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+-----
      age |      .0476294      .003864      12.33      0.000      .0400561      .0552027
      sex |
female |      .1112819      .0966378      1.15      0.250      -.0781248      .3006886

```

born							
no		-.9778304	.1865018	-5.24	0.000	-1.343367	-.6122936
marital							
widowed		-.5084458	.2090768	-2.43	0.015	-.9182288	-.0986628
divorced		-.5989672	.1349731	-4.44	0.000	-.8635096	-.3344249
separated		-.3609247	.2461983	-1.47	0.143	-.8434645	.121615
never mar..		-.4303034	.1293215	-3.33	0.001	-.6837689	-.1768379
childs		-.0350106	.0336983	-1.04	0.299	-.101058	.0310368
educ		.2879809	.0198907	14.48	0.000	.2489958	.3269661
_cons		-4.793928	.3483981	-13.76	0.000	-5.476775	-4.11108

```
. mchange
logit: Changes in Pr(y) | Number of obs = 2590
Expression: Pr(vote), predict(pr)
```

		Change	p-value
age			
	+1	0.008	0.000
	+SD	0.125	0.000
	Marginal	0.008	0.000
sex			
	female vs male	0.019	0.250
born			
	no vs yes	-0.185	0.000
marital			
	widowed vs married	-0.089	0.019
	divorced vs married	-0.106	0.000
	separated vs married	-0.062	0.160
	never married vs married	-0.075	0.001
	divorced vs widowed	-0.017	0.682
	separated vs widowed	0.027	0.626
	never married vs widowed	0.014	0.746
	separated vs divorced	0.044	0.355
	never married vs divorced	0.031	0.278
	never married vs separated	-0.013	0.788
childs			
	+1	-0.006	0.301
	+SD	-0.010	0.302
	Marginal	-0.006	0.298
educ			
	+1	0.048	0.000
	+SD	0.128	0.000
	Marginal	0.050	0.000

```
Average predictions
| 0 1
-----+-----
Pr(y|base) | 0.317 0.683
```

Here we can see how probability changes when we go up by 1 unit (on average) and when we go up by 1 SD. For dichotomies, it is the difference between two categories. If values of independent variables are specified, predictions are computed at these values. For variables whose values are not specified, changes are averaged across observed values (i.e., margins' asobserved option).

Compare:

```
. mchange, atmeans
logit: Changes in Pr(y) | Number of obs = 2590
Expression: Pr(vote), predict(pr)
```

		Change	p-value
-----+-----			
age			
	+1	0.009	0.000
	+SD	0.132	0.000
	Marginal	0.009	0.000
sex			
	female vs male	0.022	0.251
born			
	no vs yes	-0.224	0.000
marital			
	widowed vs married	-0.101	0.024
	divorced vs married	-0.121	0.000
	separated vs married	-0.069	0.171
	never married vs married	-0.084	0.001
	divorced vs widowed	-0.020	0.680
	separated vs widowed	0.032	0.626
	never married vs widowed	0.017	0.747
	separated vs divorced	0.052	0.350
	never married vs divorced	0.037	0.280
	never married vs separated	-0.015	0.787
childs			
	+1	-0.007	0.303
	+SD	-0.012	0.305
	Marginal	-0.007	0.299
educ			
	+1	0.054	0.000
	+SD	0.134	0.000
	Marginal	0.057	0.000

Predictions at base value

	0	1
-----+-----		
Pr(y base)	0.274	0.726

Base values of regressors

		2.	2.	2.	3.	4.
	age	sex	born	marital	marital	marital
-----+-----						
at	46.9	.553	.0645	.0927	.162	.0351
	5.					
	marital	childs	educ			
-----+-----						
at	.243	1.84	13.4			

1: Estimates with margins option atmeans.

We can also request more change units by using amount option or delta option, as well as more stats; we can also limit this investigation to certain variables:

```
. mchange, amount(all)
logit: Changes in Pr(y) | Number of obs = 2590
Expression: Pr(vote), predict(pr)
```

		Change	p-value
-----+-----			
age			
	0 to 1	0.008	0.000
	+1	0.008	0.000
	+SD	0.125	0.000
	Range	0.505	0.000
	Marginal	0.008	0.000

sex			
	female vs male	0.019	0.250
born			
	no vs yes	-0.185	0.000
marital			
	widowed vs married	-0.089	0.019
	divorced vs married	-0.106	0.000
	separated vs married	-0.062	0.160
	never married vs married	-0.075	0.001
	divorced vs widowed	-0.017	0.682
	separated vs widowed	0.027	0.626
	never married vs widowed	0.014	0.746
	separated vs divorced	0.044	0.355
	never married vs divorced	0.031	0.278
	never married vs separated	-0.013	0.788
childs			
	0 to 1	-0.006	0.291
	+1	-0.006	0.301
	+SD	-0.010	0.302
	Range	-0.050	0.305
	Marginal	-0.006	0.298
educ			
	0 to 1	0.020	0.000
	+1	0.048	0.000
	+SD	0.128	0.000
	Range	0.858	0.000
	Marginal	0.050	0.000

Average predictions

	0	1
Pr(y base)	0.317	0.683

. mchange educ, delta(5) statistics(all)

logit: Changes in Pr(y) | Number of obs = 2590

Expression: Pr(vote), predict(pr)

	Change	p-value	LL	UL	z-value
educ					
+1	0.048	0.000	0.043	0.054	17.498
+delta	0.195	0.000	0.177	0.212	22.334
Marginal	0.050	0.000	0.044	0.056	16.917

	Std Err	From	To
educ			
+1	0.003	0.683	0.732
+delta	0.009	0.683	0.878
Marginal	0.003	.z	.z

Average predictions

	0	1
Pr(y base)	0.317	0.683

1: Delta equals 5.

We can get these changes for a more limited range than min to max:

```
. mchange, amount(range) trim(5)
logit: Changes in Pr(y) | Number of obs = 2590
Expression: Pr(vote), predict(pr)
```

		Change	p-value
age			
	5% to 95%	0.428	0.000
sex			
	female vs male	0.019	0.250
born			
	no vs yes	-0.185	0.000
marital			
	widowed vs married	-0.089	0.019
	divorced vs married	-0.106	0.000
	separated vs married	-0.062	0.160
	never married vs married	-0.075	0.001
	divorced vs widowed	-0.017	0.682
	separated vs widowed	0.027	0.626
	never married vs widowed	0.014	0.746
	separated vs divorced	0.044	0.355
	never married vs divorced	0.031	0.278
	never married vs separated	-0.013	0.788
childs			
	5% to 95%	-0.031	0.300
educ			
	5% to 95%	0.448	0.000

```
Average predictions
| 0 1
-----+-----
Pr(y|base) | 0.317 0.683
```

```
. centile educ, centile(0 5 95 100)
```

Variable	Obs	Percentile	Centile	-- Binom. Interp. -- [95% Conf. Interval]	
educ	2753	0	0	0	0*
		5	8	8	9
		95	18	18	18
		100	20	20	20*

\* Lower (upper) confidence limit held at minimum (maximum) of sample

People often conclude that two groups are different if confidence intervals do not overlap – but that is usually too conservative. Looking at discrete changes with a confidence interval is more informative. Note that if you have linked variables – variables with squared or cubed terms, or with interactions – you should use factor variable notation (as in the example below), and then the commands will keep track of that for you when generating predictions.

```
. logit vote childs i.sex i.born##c.educ i.marital age
Iteration 0: log likelihood = -1616.8899
Iteration 1: log likelihood = -1357.6437
Iteration 2: log likelihood = -1347.2285
Iteration 3: log likelihood = -1347.1953
Iteration 4: log likelihood = -1347.1953
Logistic regression
```

Number of obs	=	2590
LR chi2(10)	=	539.39
Prob > chi2	=	0.0000
Pseudo R2	=	0.1668

```
Log likelihood = -1347.1953
```

vote	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
childs	-.0340252	.0337392	-1.01	0.313	-.1001528	.0321024
sex						
female	.1040851	.0968511	1.07	0.283	-.0857396	.2939098
born						
no	1.40973	.7190502	1.96	0.050	.000417	2.819042
educ	.310949	.0214335	14.51	0.000	.26894	.352958
born#c.educ						
no	-.177672	.0521626	-3.41	0.001	-.2799088	-.0754351
marital						
widowed	-.4717709	.2097889	-2.25	0.025	-.8829496	-.0605922
divorced	-.5976599	.1353399	-4.42	0.000	-.8629213	-.3323985
separated	-.3473378	.2469307	-1.41	0.160	-.8313131	.1366374
never married	-.4409393	.1298164	-3.40	0.001	-.6953748	-.1865039
age	.0476348	.0038642	12.33	0.000	.0400611	.0552086
_cons	-5.086571	.3639937	-13.97	0.000	-5.799985	-4.373156

```
. mgen, atmeans at(educ=(0(2)20) born=1) stub(nb_)
Predictions from: margins, atmeans at(educ=(0(2)20) born=1) predict(pr)
Variable   Obs Unique      Mean      Min      Max  Label
-----
nb_pr1      11      11  .4991813  .0424739  .9570376  pr(y=1) from margins
nb_ll1      11      11  .4690912  .0204639  .9437191  95% lower limit
nb_ul1      11      11  .5292713  .0644839  .970356   95% upper limit
nb_educ     11      11      10         0      20  highest year of school com...
```

Specified values of covariates

	2.		2.	3.	4.	5.
childs	sex	born	marital	marital	marital	marital
1.838996	.5532819	1	.0926641	.1617761	.0351351	.2428571

age

-----

46.93591

```
. mgen, atmeans at(educ=(0(2)20) born=2) stub(fb_)
Predictions from: margins, atmeans at(educ=(0(2)20) born=2) predict(pr)
Variable   Obs Unique      Mean      Min      Max  Label
-----
fb_pr1      11      11  .4209988  .1537178  .7230831  pr(y=1) from margins
fb_ll1      11      11  .2880444  -.0171945  .5822673  95% lower limit
fb_ul1      11      11  .5539532  .3246301  .8638988  95% upper limit
fb_educ     11      11      10         0      20  highest year of school co...
```

Specified values of covariates

	2.		2.	3.	4.	5.
childs	sex	born	marital	marital	marital	marital
1.838996	.5532819	2	.0926641	.1617761	.0351351	.2428571

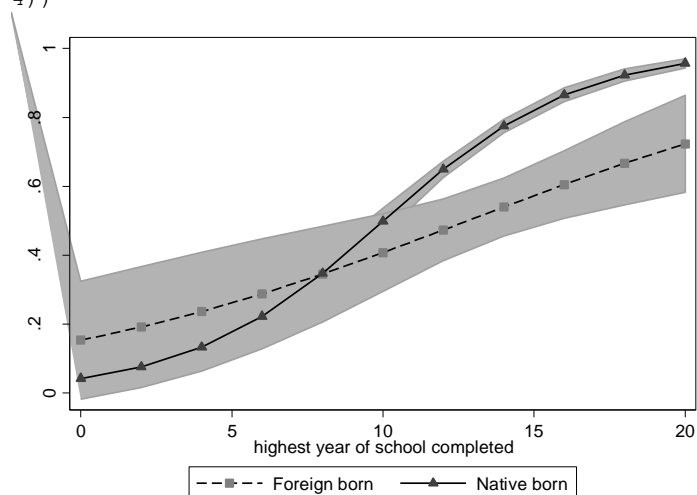
age

-----

46.93591



```
. lab var fb_pr1 "Foreign born"
. lab var nb_pr1 "Native born"
. graph twoway (rarea nb_ul1 nb_ll1 nb_educ, col(gs10)) (rarea fb_ul1 fb_ll1 fb_educ,
color(gs10)) (connected fb_pr1 nb_pr1 fb_educ, lpattern(dash solid)), legend(order(3
4))
```



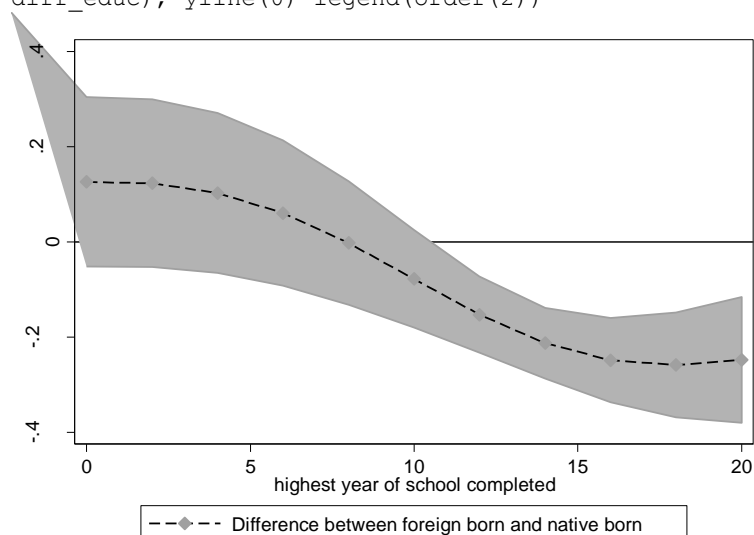
```
. mgen, dydx(born) at(educ=(0(2)20)) stub(diff_)
```

Predictions from: margins, dydx(born) at(educ=(0(2)20)) predict(pr)

Variable	Obs	Unique	Mean	Min	Max	Label
diff_d_pr1	11	11	-.0717905	-.2587097	.1261362	d_pr(y=1) from margins
diff_ll1	11	11	-.1983488	-.3805256	-.0519222	95% lower limit
diff_ul1	11	11	.0547677	-.1604816	.3041946	95% upper limit
diff_educ	11	11	10	0	20	highest year of school completed

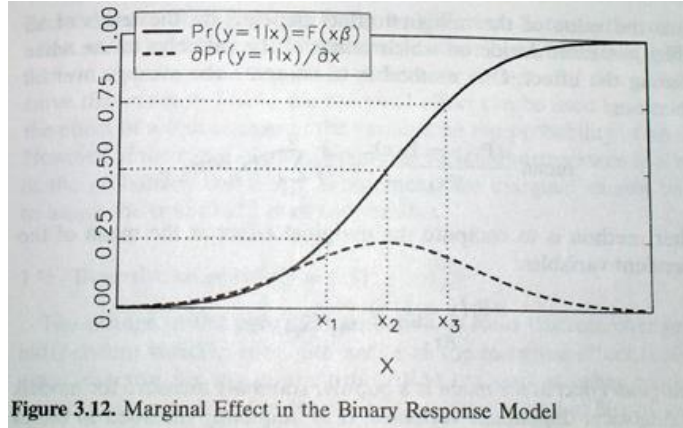
```
. lab var diff_d_pr1 "Difference between foreign born and native born"
```

```
. graph twoway (rarea diff_ul1 diff_ll1 diff_educ, col(gs10)) (connected diff_d_pr1
diff_educ), yline(0) legend(order(2))
```

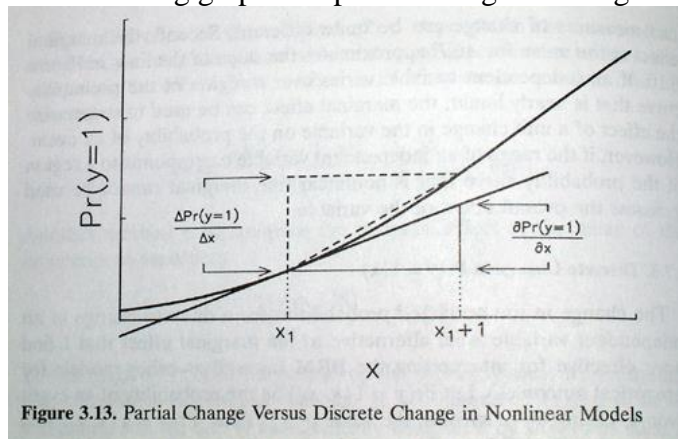


### B. Marginal effects.

One thing that we saw in the mchange output above but did not discuss yet is marginal effects – these are partial derivatives, slopes of probability curve at a certain set of values of independent variables. Marginal effects, of course, vary along X; they are the largest at the value of X that corresponds to  $P(Y=1|X)=.5$  – this can be seen in the graph.



The following graph compares a marginal change and a discrete change at a specific point:



Marginal effects are inappropriate for binary independent variables; that's why discrete changes are reported for those instead.

There are three ways that marginal effects are usually estimated:

1. Marginal effects at the mean (MEM)
2. Marginal effects at representative values (MER)
3. Average marginal effects (AME) (marginal effects are estimated at all values and then averaged out)

```
. logit vote age i.sex i.born i.marital childs educ
```

```
Iteration 0:  log likelihood = -1616.8899
Iteration 1:  log likelihood = -1361.6039
Iteration 2:  log likelihood = -1352.4837
Iteration 3:  log likelihood = -1352.4548
Iteration 4:  log likelihood = -1352.4548
```

Logistic regression

Number of obs = 2590

Log likelihood = -1352.4548

LR chi2(9)	=	528.87
Prob > chi2	=	0.0000
Pseudo R2	=	0.1635

vote	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0476294	.003864	12.33	0.000	.0400561	.0552027
sex						
female	.1112819	.0966378	1.15	0.250	-.0781248	.3006886
born						
no	-.9778304	.1865018	-5.24	0.000	-1.343367	-.6122936
marital						
widowed	-.5084458	.2090768	-2.43	0.015	-.9182288	-.0986628
divorced	-.5989672	.1349731	-4.44	0.000	-.8635096	-.3344249
separated	-.3609247	.2461983	-1.47	0.143	-.8434645	.121615
never mar..	-.4303034	.1293215	-3.33	0.001	-.6837689	-.1768379
childs	-.0350106	.0336983	-1.04	0.299	-.101058	.0310368
educ	.2879809	.0198907	14.48	0.000	.2489958	.3269661
_cons	-4.793928	.3483981	-13.76	0.000	-5.476775	-4.11108

### Average marginal effects (AME):

. mchange

logit: Changes in Pr(y) | Number of obs = 2590

Expression: Pr(vote), predict(pr)

	Change	p-value
age		
+1	0.008	0.000
+SD	0.125	0.000
Marginal	0.008	0.000
sex		
female vs male	0.019	0.250
born		
no vs yes	-0.185	0.000
marital		
widowed vs married	-0.089	0.019
divorced vs married	-0.106	0.000
separated vs married	-0.062	0.160
never married vs married	-0.075	0.001
divorced vs widowed	-0.017	0.682
separated vs widowed	0.027	0.626
never married vs widowed	0.014	0.746
separated vs divorced	0.044	0.355
never married vs divorced	0.031	0.278
never married vs separated	-0.013	0.788
childs		
+1	-0.006	0.301
+SD	-0.010	0.302
Marginal	-0.006	0.298
educ		
+1	0.048	0.000
+SD	0.128	0.000
Marginal	0.050	0.000

Average predictions

	0	1
Pr(y base)	0.317	0.683

In addition to mchange, we can also obtain marginal effects with dydx option in margins:

```
. margins, dydx(*)
```

Average marginal effects                      Number of obs     =        2590  
Model VCE        : OIM

Expression     : Pr(vote), predict()  
dy/dx w.r.t.   : age 2.sex 2.born 2.marital 3.marital 4.marital 5.marital  
                 childs educ

		Delta-method				[95% Conf. Interval]	
		dy/dx	Std. Err.	z	P> z		
age		.0083074	.0006053	13.72	0.000	.007121	.0094937
sex							
female		.0194592	.016928	1.15	0.250	-.0137191	.0526375
born							
no		-.1851289	.0364786	-5.07	0.000	-.2566257	-.1136321
marital							
widowed		-.0892473	.0380707	-2.34	0.019	-.1638646	-.0146301
divorced		-.1062677	.0244728	-4.34	0.000	-.1542335	-.0583019
separated		-.0621571	.044188	-1.41	0.160	-.148764	.0244498
never mar..		-.0747909	.0231535	-3.23	0.001	-.1201708	-.0294109
childs		-.0061064	.0058731	-1.04	0.298	-.0176175	.0054047
educ		.0502287	.0029691	16.92	0.000	.0444093	.0560481

Note: dy/dx for factor levels is the discrete change from the base level.

### Marginal effects at the mean (MEM):

```
. mchange, atmeans
```

logit: Changes in Pr(y) | Number of obs = 2590

Expression: Pr(vote), predict(pr)

		Change	p-value
age			
	+1	0.009	0.000
	+SD	0.132	0.000
	Marginal	0.009	0.000
sex			
	female vs male	0.022	0.251
born			
	no vs yes	-0.224	0.000
marital			
	widowed vs married	-0.101	0.024

divorced vs married		-0.121	0.000
separated vs married		-0.069	0.171
never married vs married		-0.084	0.001
divorced vs widowed		-0.020	0.680
separated vs widowed		0.032	0.626
never married vs widowed		0.017	0.747
separated vs divorced		0.052	0.350
never married vs divorced		0.037	0.280
never married vs separated		-0.015	0.787
childs			
	+1	-0.007	0.303
	+SD	-0.012	0.305
	Marginal	-0.007	0.299
educ			
	+1	0.054	0.000
	+SD	0.134	0.000
	Marginal	0.057	0.000

Predictions at base value

	0	1
Pr(y base)	0.274	0.726

Base values of regressors

	age	sex	born	marital	marital	marital
at	46.9	.553	.0645	.0927	.162	.0351

	marital	childs	educ
at	.243	1.84	13.4

1: Estimates with margins option atmeans.

We can also get them centered at means (the default option shows mean+1):

```
. mchange, atmeans centered
```

logit: Changes in Pr(y) | Number of obs = 2590

Expression: Pr(vote), predict(pr)

	Change	p-value
age		
+1 centered	0.009	0.000
+SD centered	0.162	0.000
Marginal	0.009	0.000
sex		
female vs male	0.022	0.251
born		
no vs yes	-0.224	0.000
marital		
widowed vs married	-0.101	0.024
divorced vs married	-0.121	0.000
separated vs married	-0.069	0.171
never married vs married	-0.084	0.001
divorced vs widowed	-0.020	0.680

Predictions at base valueBase values of regressors

```
1: Estimates with margins option at means.
```

```
. margins, atmeans
```

```

Expression      : Pr(vote), predict()
at              : age           = 46.93591 (mean)
                1.sex          = .4467181 (mean)
                2.sex          = .5532819 (mean)
                1.born         = .9355212 (mean)
                2.born         = .0644788 (mean)
                1.marital      = .4675676 (mean)
                2.marital      = .0926641 (mean)
                3.marital      = .1617761 (mean)
                4.marital      = .0351351 (mean)
                5.marital      = .2428571 (mean)
                childs         = 1.838996 (mean)
                educ           = 13.39459 (mean)

```

30

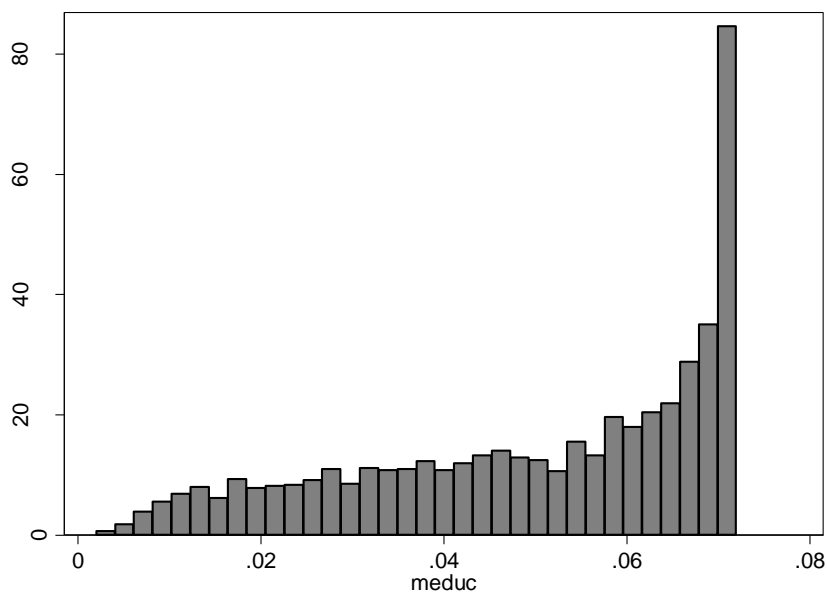
```
. di .7255038*(1-.7255038)* .2879809
.05735083
```

Histogram of marginal effects can help us better understand whether MEM or AME better represent what is going on in our sample:

```
. predict double prhat if e(sample)
(option pr assumed; Pr(vote))
(175 missing values generated)

. gen double meduc=prhat*(1-prhat) *_b[educ]
(175 missing values generated)

. histogram meduc
(bin=34, start=.00199118, width=.00205894)
```



### Marginal effects at representative values (MER):

```
. mchange educ, at(educ=12)

logit: Changes in Pr(y) | Number of obs = 2590

Expression: Pr(vote), predict(pr)
```

	Change	p-value
educ		
+1	0.057	0.000
+SD	0.154	0.000
Marginal	0.059	0.000

Average predictions

	0	1
Pr(y base)	0.382	0.618

Base values of regressors

		educ
-----+-----		
at		12

```
. mchange educ, at(educ=16)
```

```
logit: Changes in Pr(y) | Number of obs = 2590
```

```
Expression: Pr(vote), predict(pr)
```

		Change	p-value
-----+-----			
educ			
+1		0.036	0.000
+SD		0.090	0.000
Marginal		0.039	0.000

```
Average predictions
```

		0	1
-----+-----			
Pr(y base)		0.182	0.818

```
Base values of regressors
```

		educ
-----+-----		
at		16

```
. mchange educ, at(educ=10)
```

```
logit: Changes in Pr(y) | Number of obs = 2590
```

```
Expression: Pr(vote), predict(pr)
```

		Change	p-value
-----+-----			
educ			
+1		0.061	0.000
+SD		0.174	0.000
Marginal		0.061	0.000

```
Average predictions
```

		0	1
-----+-----			
Pr(y base)		0.503	0.497

```
Base values of regressors
```

		educ
-----+-----		
at		10