

**SC704: Topics in Multivariate Analysis**  
**Instructor: Natasha Sarkisian**  
Count Data Models

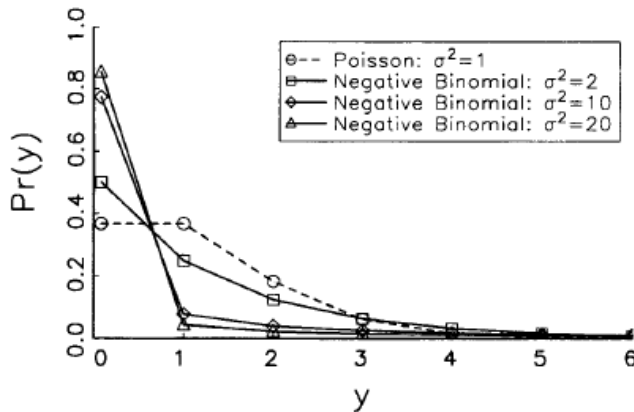
Negative Binomial Model

Using Poisson, we attempted to account for some sources of heterogeneity - but the model doesn't fit very well. Maybe we didn't take into account all sources of heterogeneity - could try additional variables. That's important to explore, but rarely helps. In practice, Poisson regression models rarely fits due to overdispersion.

There is another process that often creates overdispersion - it is known as contagion - violation of the assumption of the independence of events. This assumption is often unrealistic; e.g. if you have your first child, that increases your chances of having your second.

To better model overdispersion from this and other sources, we can use negative binomial model. It allows taking into account unobserved heterogeneity. To do so, it introduces an additional parameter - alpha, known as the dispersion parameter. Increasing alpha increases conditional variance of X. If alpha is zero, the model becomes regular Poisson model. Here's a comparison of Poisson and negative binomial distributions with different variances for mean count=1 and mean count=10:

Panel A:  $E(y)=1$



Panel B:  $E(y)=10$

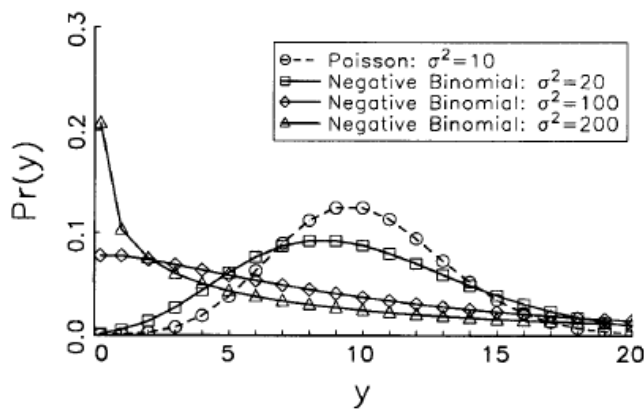
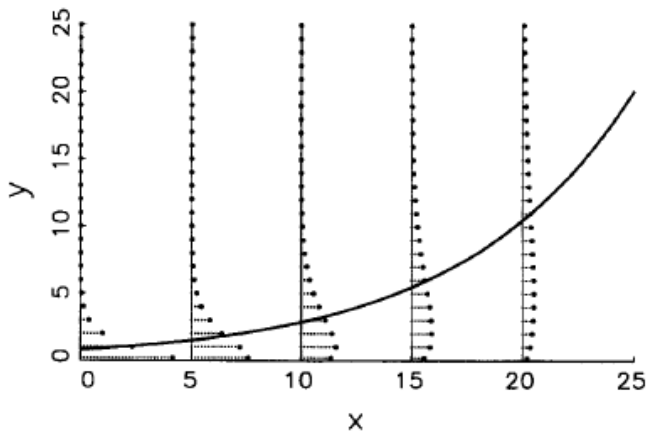


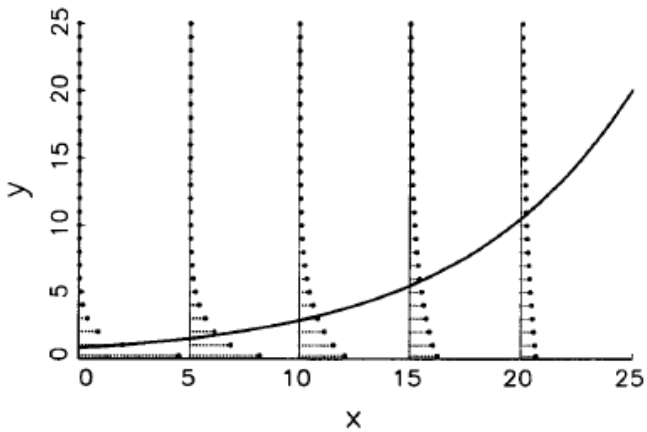
Figure 8.6. Comparisons of the Negative Binomial and Poisson Distributions

And here's an example of regression curves for negative binomial models:

Panel A: NBRM with  $\alpha=0.5$



Panel B: NBRM with  $\alpha=1.0$



**Figure 8.7.** Distribution of Counts for the Negative Binomial Regression Model

Now let's run NB model for our data:

```
. nbreg childs sex married sibs born educ
Negative binomial regression
Log likelihood = -4711.6789
Number of obs = 2745
LR chi2(5) = 380.47
Prob > chi2 = 0.0000
Pseudo R2 = 0.0388
```

childs	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
sex	.2086278	.0346569	6.02	0.000	.1407014 .2765542
married	.471206	.034682	13.59	0.000	.4032305 .5391816
sibs	.0397041	.0054244	7.32	0.000	.0290725 .0503358
born	-.2231164	.0616061	-3.62	0.000	-.3438622 -.1023706
educ	-.0616831	.0058316	-10.58	0.000	-.0731129 -.0502534
_cons	.9198597	.1211683	7.59	0.000	.6823743 1.157345
/lnalpha	-1.523939	.1086487			-1.736886 -1.310991
alpha	.2178522	.0236694			.1760678 .2695528

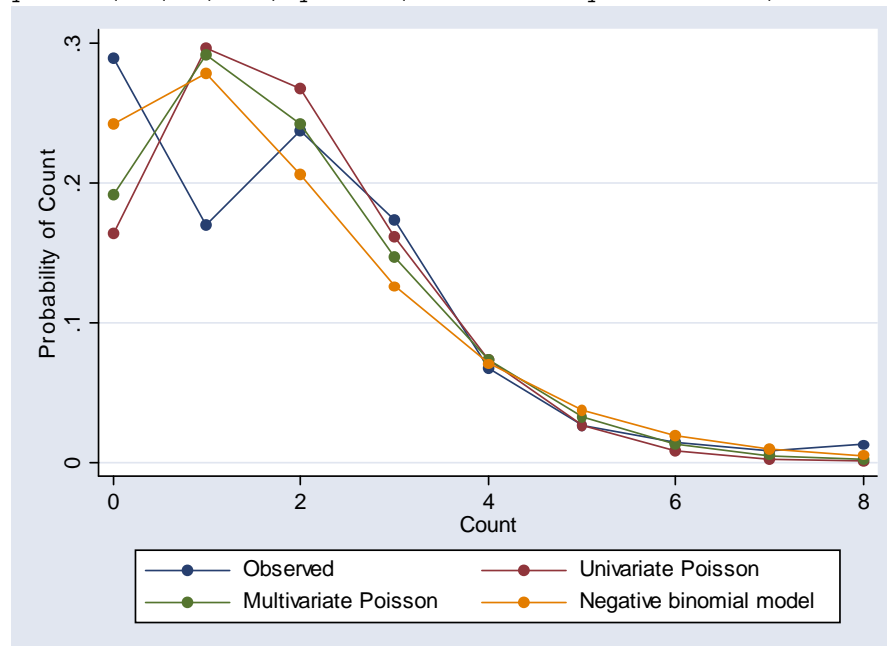
Likelihood-ratio test of alpha=0:  $\text{chibar2}(01) = 145.66$  Prob>= $\text{chibar2} = 0.000$

Interpretation of the results for negative binomial model is exactly the same as for Poisson model. But we have an extra line of output to interpret - the likelihood-ratio test. This allows us to see whether NB model should be used in place of regular Poisson. If probability is below the cutoff, it means that there is overdispersion (Alpha is not zero) and we should be using NB model rather than Poisson.

Now let's compare their performance graphically:

```
. prcounts nb, plot max(8)
(19 missing values generated)
. lab var nbpreq "Negative binomial model"

. gr twoway connected poisobeq poispreq prmpreq expopreq nbpreq poisval,
ylabel(0 (.1) .3) ytitle("Probability of Count")
```



The graph confirms the results of the test: NB model does better than regular multivariate Poisson. But it still underpredicts zeros and overpredicts ones. Unfortunately, the goodness of fit tests that are available after Poisson are not available after negative binomial. But the significance test for alpha tells us if Poisson performs better than negative binomial.

The interpretation tools for nbreg are the same as for poisson; we can get IRR and use prtab, prgen, prchange, and prvalue commands, as well as mfx command. We could also estimate this model with exposure.

As for diagnostics, everything is similar to Poisson, except for boxtid which doesn't work with nbreg. To obtain a GLM negative binomial model that's identical to the one estimated to nbreg, you need to specify the exact alpha to use - otherwise it uses the default value of 1 and the results differ. So here we use:

```
. glm childs sex married sibs born educ, family(nb .2178552)
```

Generalized linear models		No. of obs	=	2745
Optimization	: ML	Residual df	=	2739
		Scale parameter	=	1
Deviance	= 3284.463783	(1/df) Deviance	=	1.199147

Pearson = 2908.984543 (1/df) Pearson = 1.062061

Variance function:  $V(u) = u + (.2178552)u^2$  [Neg. Binomial]  
 Link function :  $g(u) = \ln(u)$  [Log]  
 AIC = 3.437289  
 Log likelihood = -4711.678905 BIC = -18401.67

childs	Coef.	OIM Std. Err.	z	P> z	[95% Conf. Interval]	
sex	.2086279	.0346384	6.02	0.000	.1407379	.2765179
married	.4712062	.0346364	13.60	0.000	.4033201	.5390924
sibs	.0397041	.0054238	7.32	0.000	.0290737	.0503346
born	-.2231165	.0616059	-3.62	0.000	-.3438618	-.1023712
educ	-.0616831	.0058316	-10.58	0.000	-.0731129	-.0502533
_cons	.9198593	.1211388	7.59	0.000	.6824317	1.157287

We can obtain residuals etc. after this.

In addition to regular nbreg where overdispersion is assumed to be constant, we can also use generalized negative binomial regression to model overdispersion:  
 . gnbreg childs sex married sibs born educ, lnalpha(sex married sibs born educ)

Generalized negative binomial regression      Number of obs = 2745  
 LR chi2(5) = 222.46  
 Prob > chi2 = 0.0000  
 Log likelihood = -4587.1261      Pseudo R2 = 0.0237

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
childs						
sex	.079685	.0354711	2.25	0.025	.0101628	.1492071
married	.3413691	.0387924	8.80	0.000	.2653374	.4174008
sibs	.0369471	.0047258	7.82	0.000	.0276847	.0462095
born	-.1967968	.0582151	-3.38	0.001	-.3108963	-.0826973
educ	-.0514978	.0056236	-9.16	0.000	-.0625199	-.0404758
_cons	1.085011	.1189463	9.12	0.000	.8518807	1.318142
lnalpha						
sex	-1.557369	.1884906	-8.26	0.000	-1.926804	-1.187934
married	-4.256861	.819715	-5.19	0.000	-5.863473	-2.650249
sibs	-.1051836	.0405024	-2.60	0.009	-.1845669	-.0258003
born	.1353893	.3910783	0.35	0.729	-.63111	.9018887
educ	.1619184	.0358938	4.51	0.000	.0915678	.232269
_cons	.3279141	.7155448	0.46	0.647	-1.074528	1.730356

Looks like overdispersion parameter varies by sex, marital status, number of siblings, and education, so the contagion process operates differently for different people.

### Zero-Inflated Count Data Models

The problem that our negative binomial model still has - underpredicting zeros, overpredicting ones -- is very common and sometimes this problem can be very severe when there are a lot of zeros in the distribution. Example - Sarkisian and Gerstel 2004 article. We can use zero-inflated count models to correct for that - they model two different processes. They assume two latent groups - one is capable of having positive counts, the other one is not - it will always have zero count. For example, some are capable of having children, and the number that they can have might vary, but others cannot have children and their count will always remain zero. But these two groups are latent - no information on actual fertility situation. We can also have zeros in the first group. We can distinguish structural zeros (this behavior is not in this person's repertoire at all) vs chance zeros (this behavior is in this person's repertoire, but did not occur during the specified period). E.g.: "How many times last week did you smoke marijuana?" Some zeros mean the person never smokes it; other zeros mean the person does smoke but did not smoke last week.

Therefore, this model is a two-step process - first, have to predict the membership in two groups - "always zero" and "not always zero" and second, predict the count in the "not always zero" group.

```
. zip childs sex married sibs born educ, inflate(sex married sibs born educ)
```

```
Zero-inflated poisson regression          Number of obs   =       2745
                                           Nonzero obs     =       1951
                                           Zero obs       =        794

Inflation model = logit                  LR chi2(5)      =       130.65
Log likelihood = -4524.192               Prob > chi2     =        0.0000
```

childs	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
childs						
sex	.0014908	.0320997	0.05	0.963	-.0614234	.0644405
married	.0307475	.0333411	0.92	0.356	-.0345999	.0960949
sibs	.0292838	.0045691	6.41	0.000	.0203286	.038239
born	-.1728303	.0563097	-3.07	0.002	-.2831953	-.0624654
educ	-.0382489	.0052824	-7.24	0.000	-.0486021	-.0278956
_cons	1.363043	.1094042	12.46	0.000	1.148615	1.577472
-----						
inflate						
sex	-1.267402	.1427508	-8.88	0.000	-1.547189	-.987616
married	-3.867796	.6722317	-5.75	0.000	-5.185346	-2.550246
sibs	-.0907598	.0284525	-3.19	0.001	-.1465256	-.034994
born	.3182067	.2733966	1.16	0.244	-.2176408	.8540542
educ	.1671403	.0267744	6.24	0.000	.1146635	.2196171
_cons	-.9103566	.5168716	-1.76	0.078	-1.923406	.102693

Note the inflate option we specified - we have to specify that option, it tells Stata what variables to use to predict the membership in "Always Zero" group. In this case, we used the same variables but we could have used a smaller subset of the variables or even different variables altogether. We'll return to interpreting this output. But let's prepare to graphically examine the fit:

```
. prcounts zip, plot max(8)
(19 missing values generated)
. lab var zipreq "ZIP"
```

```

. zinb childs sex married sibs born educ, inflate(sex married sibs born educ)
Zero-inflated negative binomial regression      Number of obs   =      2745
                                                Nonzero obs     =      1951
                                                Zero obs       =       794

Inflation model = logit                      LR chi2(5)      =     124.23
Log likelihood = -4522.91                    Prob > chi2     =      0.0000

```

	childs	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----							
childs							
	sex	.0060583	.0331917	0.18	0.855	-.0589961	.0711128
	married	.0346028	.0344018	1.01	0.314	-.0328234	.102029
	sibs	.0297016	.004743	6.26	0.000	.0204055	.0389977
	born	-.1730859	.0572733	-3.02	0.003	-.2853394	-.0608324
	educ	-.0384851	.0054302	-7.09	0.000	-.0491281	-.0278422
	_cons	1.347192	.1125643	11.97	0.000	1.12657	1.567814
-----							
inflate							
	sex	-1.290154	.1468538	-8.79	0.000	-1.577982	-1.002326
	married	-4.405718	1.215488	-3.62	0.000	-6.78803	-2.023406
	sibs	-.0911606	.02947	-3.09	0.002	-.1489207	-.0334006
	born	.3417874	.2818703	1.21	0.225	-.2106681	.894243
	educ	.1715742	.0277136	6.19	0.000	.1172565	.2258919
	_cons	-.9919407	.5360101	-1.85	0.064	-2.042501	.0586197
-----							
	/lnalpha	-3.718083	.6593754	-5.64	0.000	-5.010435	-2.425731
-----							
	alpha	.0242805	.0160099			.006668	.0884134
-----							

```

. prcounts zinb, plot max(8)
(19 missing values generated)
. lab var zinbpreq "ZINB"

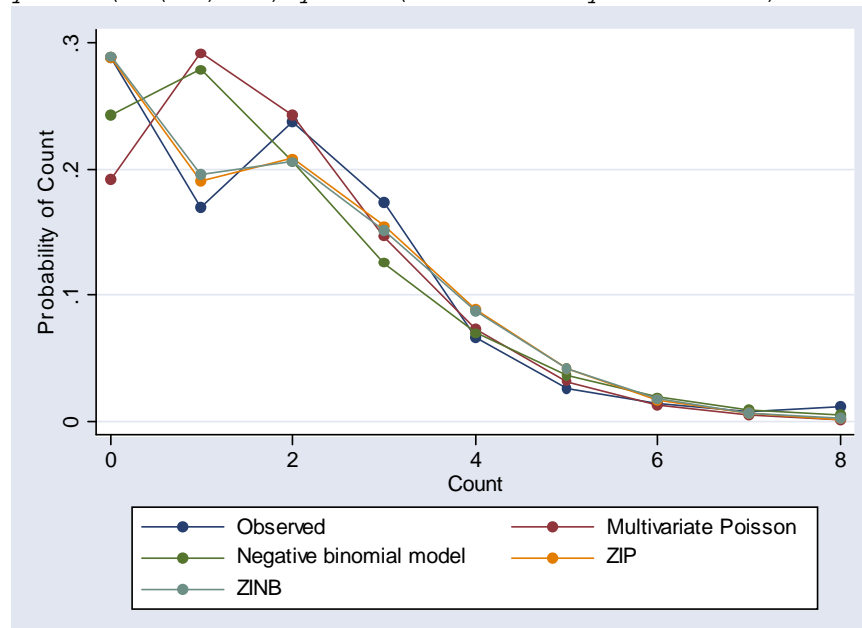
```

Before interpreting the results, let's figure out which model fits best.

```

. gr twoway connected poisobeq prmpreq nbpreq zipreq zinbpreq poisval,
ylabel(0 (.1) .3) ytitle("Probability of Count")

```



Both ZIP and ZINB approximate the observed distribution much better than regular Poisson and NB models. We could also plot deviations from observed counts rather than actual counts and get comparisons of fit:

```
. countfit childs sex married sibs born educ, inflate(sex married sibs born educ)
```

Variable		PRM	NBRM	ZIP	ZINB
childs					
respondents sex		1.216	1.232	1.001	1.006
		6.73	6.02	0.05	0.18
R is married		1.566	1.602	1.031	1.035
		15.54	13.59	0.92	1.01
number of brothers and sisters		1.039	1.041	1.030	1.030
		9.14	7.32	6.41	6.26
was r born in this country		0.802	0.800	0.841	0.841
		-4.23	-3.62	-3.07	-3.02
highest year of school completed		0.940	0.940	0.962	0.962
		-12.81	-10.58	-7.24	-7.09
Constant		2.598	2.509	3.908	3.847
		9.45	7.59	12.46	11.97
lnalpha					
Constant			0.218		0.024
			-14.03		-5.64
inflate					
respondents sex				0.282	0.275
				-8.88	-8.79
R is married				0.021	0.012
				-5.75	-3.62
number of brothers and sisters				0.913	0.913
				-3.19	-3.09
was r born in this country				1.375	1.407
				1.16	1.21
highest year of school completed				1.182	1.187
				6.24	6.19
Constant				0.402	0.371
				-1.76	-1.85
Statistics					
alpha			0.218		
N		2745	2745	2745	2745
ll		-4784.508	-4711.679	-4524.192	-4522.910
bic		9616.521	9478.781	9143.394	9148.749
aic		9581.016	9437.358	9072.383	9071.821

legend: b/t

Comparison of Mean Observed and Predicted Count

Model	Maximum Difference	At Value	Mean  Diff
PRM	-0.122	1	0.028
NBRM	-0.109	1	0.027
ZIP	0.030	2	0.012
ZINB	0.032	2	0.013

PRM: Predicted and actual probabilities

Count Actual Predicted |Diff| Pearson

0	0.289	0.192	0.097	135.055
1	0.170	0.292	0.122	139.312
2	0.238	0.242	0.005	0.231
3	0.174	0.147	0.027	13.674
4	0.067	0.073	0.006	1.361
5	0.026	0.032	0.006	3.069
6	0.015	0.013	0.002	0.526
7	0.008	0.005	0.003	5.097
8	0.012	0.002	0.011	163.156
9	0.000	0.001	0.001	1.924
-----				
Sum	1.000	1.000	0.278	463.405

NBRM: Predicted and actual probabilities

Count	Actual	Predicted	Diff	Pearson
-----				
0	0.289	0.242	0.047	24.952
1	0.170	0.279	0.109	116.103
2	0.238	0.206	0.032	13.512
3	0.174	0.126	0.048	50.004
4	0.067	0.070	0.003	0.315
5	0.026	0.037	0.011	8.820
6	0.015	0.019	0.005	3.010
7	0.008	0.010	0.002	0.867
8	0.012	0.005	0.007	30.214
9	0.000	0.003	0.003	7.016
-----				
Sum	1.000	0.997	0.265	254.813

ZIP: Predicted and actual probabilities

Count	Actual	Predicted	Diff	Pearson
-----				
0	0.289	0.288	0.001	0.014
1	0.170	0.191	0.021	6.403
2	0.238	0.208	0.030	11.561
3	0.174	0.155	0.019	6.512
4	0.067	0.089	0.021	14.210
5	0.026	0.042	0.016	16.286
6	0.015	0.017	0.003	1.083
7	0.008	0.006	0.002	1.298
8	0.012	0.002	0.010	135.546
9	0.000	0.001	0.001	1.886
-----				
Sum	1.000	1.000	0.124	194.798

ZINB: Predicted and actual probabilities

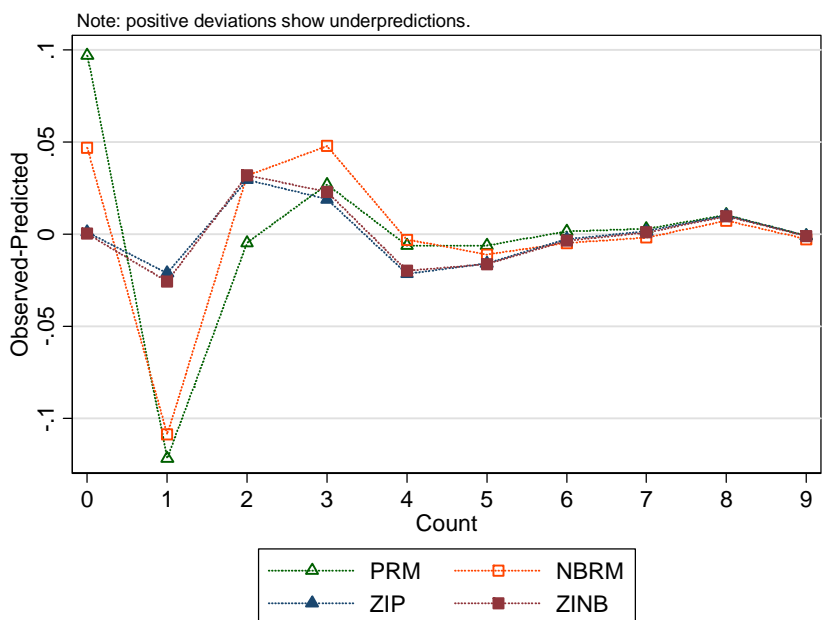
Count	Actual	Predicted	Diff	Pearson
-----				
0	0.289	0.289	0.000	0.001
1	0.170	0.196	0.026	9.202
2	0.238	0.206	0.032	13.730
3	0.174	0.151	0.023	9.695
4	0.067	0.087	0.020	12.320
5	0.026	0.042	0.016	16.787
6	0.015	0.018	0.003	1.855
7	0.008	0.007	0.001	0.389
8	0.012	0.003	0.010	104.052



9	0.000	0.001	0.001	2.445
Sum	1.000	1.000	0.132	170.477

Tests and Fit Statistics

PRM	BIC=-12117.116	AIC=	3.490	Prefer	Over	Evidence
vs NBRM	BIC=-12254.857	dif=	137.740	NBRM	PRM	Very strong
	AIC= 3.438	dif=	0.052	NBRM	PRM	
	LRX2= 145.658	prob=	0.000	NBRM	PRM	p=0.000
vs ZIP	BIC=-12590.244	dif=	473.127	ZIP	PRM	Very strong
	AIC= 3.305	dif=	0.185	ZIP	PRM	
	Vuong= 11.165	prob=	0.000	ZIP	PRM	p=0.000
vs ZINB	BIC=-12584.889	dif=	467.772	ZINB	PRM	Very strong
	AIC= 3.305	dif=	0.185	ZINB	PRM	
NBRM	BIC=-12254.857	AIC=	3.438	Prefer	Over	Evidence
vs ZIP	BIC=-12590.244	dif=	335.387	ZIP	NBRM	Very strong
	AIC= 3.305	dif=	0.133	ZIP	NBRM	
vs ZINB	BIC=-12584.889	dif=	330.032	ZINB	NBRM	Very strong
	AIC= 3.305	dif=	0.133	ZINB	NBRM	
	Vuong= 10.441	prob=	0.000	ZINB	NBRM	p=0.000
ZIP	BIC=-12590.244	AIC=	3.305	Prefer	Over	Evidence
vs ZINB	BIC=-12584.889	dif=	-5.355	ZIP	ZINB	Positive
	AIC= 3.305	dif=	0.000	ZINB	ZIP	
	LRX2= 2.563	prob=	0.055	ZINB	ZIP	p=0.000



So now let's interpret this final model:

```
. zip childsex married sibs born educ, inflate(sex married sibs born educ)
Zero-inflated poisson regression          Number of obs   =      2745
                                           Nonzero obs     =      1951
                                           Zero obs       =       794
Inflation model = logit                  LR chi2(5)      =     130.65
Log likelihood = -4524.192                Prob > chi2     =      0.0000
```

childsex	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
childsex						
sex	.0014908	.0320997	0.05	0.963	-.0614234	.064405
married	.0307475	.0333411	0.92	0.356	-.0345999	.0960949
sibs	.0292838	.0045691	6.41	0.000	.0203286	.038239
born	-.1728303	.0563097	-3.07	0.002	-.2831953	-.0624654
educ	-.0382489	.0052824	-7.24	0.000	-.0486021	-.0278956
_cons	1.363043	.1094042	12.46	0.000	1.148615	1.577472
-----						
inflate						
sex	-1.267402	.1427508	-8.88	0.000	-1.547189	-.987616
married	-3.867796	.6722317	-5.75	0.000	-5.185346	-2.550246
sibs	-.0907598	.0284525	-3.19	0.001	-.1465256	-.034994
born	.3182067	.2733966	1.16	0.244	-.2176408	.8540542
educ	.1671403	.0267744	6.24	0.000	.1146635	.2196171
_cons	-.9103566	.5168716	-1.76	0.078	-1.923406	.102693
-----						

The first set of coefficients is from the equation predicting counts for the "Not Always Zero" group. These show that number of siblings increases number of children and being foreign born and having more education decreases it. These coefficients can be interpreted the same way as regular Poisson coefficients.

The second set of coefficients is from the equation that predicts membership in "Always Zero" group. These can be interpreted as logit coefficients. Note that they predict zeros - so their sign will usually be the opposite to that of the coefficients in the upper half of the output. These show that women are less likely than men to be in "Always zero" group, married are less likely than single people to be in it, those with more siblings are also likely to be in it, and those with more education are more likely to be in "Always zero" group.

To be able to interpret the size of these effects, let's use listcoef:

```
. listcoef
zip (N=2745): Factor Change in Expected Count
Observed SD: 1.6887584
Count Equation: Factor Change in Expected Count for Those Not Always 0
```

childsex	b	z	P> z	e^b	e^bStdX	SDofX
-----						
sex	0.00149	0.046	0.963	1.0015	1.0007	0.4970
married	0.03075	0.922	0.356	1.0312	1.0154	0.4985
sibs	0.02928	6.409	0.000	1.0297	1.0919	3.0008
born	-0.17283	-3.069	0.002	0.8413	0.9512	0.2893
educ	-0.03825	-7.241	0.000	0.9625	0.8925	2.9741
-----						

Binary Equation: Factor Change in Odds of Always 0

Always0	b	z	P> z	e^b	e^bStdX	SDofX
sex	-1.26740	-8.878	0.000	0.2816	0.5326	0.4970
married	-3.86780	-5.754	0.000	0.0209	0.1454	0.4985
sibs	-0.09076	-3.190	0.001	0.9132	0.7616	3.0008
born	0.31821	1.164	0.244	1.3747	1.0964	0.2893
educ	0.16714	6.243	0.000	1.1819	1.6439	2.9741

Or better yet with percentages:

. listcoef, percent

zip (N=2745): Percentage Change in Expected Count

Observed SD: 1.6887584

Count Equation: Percentage Change in Expected Count for Those Not Always 0

childs	b	z	P> z	%	%StdX	SDofX
sex	0.00149	0.046	0.963	0.1	0.1	0.4970
married	0.03075	0.922	0.356	3.1	1.5	0.4985
sibs	0.02928	6.409	0.000	3.0	9.2	3.0008
born	-0.17283	-3.069	0.002	-15.9	-4.9	0.2893
educ	-0.03825	-7.241	0.000	-3.8	-10.8	2.9741

Binary Equation: Factor Change in Odds of Always 0

Always0	b	z	P> z	%	%StdX	SDofX
sex	-1.26740	-8.878	0.000	-71.8	-46.7	0.4970
married	-3.86780	-5.754	0.000	-97.9	-85.5	0.4985
sibs	-0.09076	-3.190	0.001	-8.7	-23.8	3.0008
born	0.31821	1.164	0.244	37.5	9.6	0.2893
educ	0.16714	6.243	0.000	18.2	64.4	2.9741

Each additional sibling increases one's count by 3%, each year of education decreases it by 3.8%, and being foreign born decreases it by 16%. At the same time, women's odds of having no kids (being in always zero group) are 71.8% lower than men's, and the odds for married to be in always zero group are 97.9% lower than for single people. Further, each additional sibling decreases one's odds of not having kids by 8.7% and each additional year of education increases those odds by 18.2%.

Further, as for regular Poisson we can interpret predicted rates and predicted probabilities. Predicted rates for native-born:

. prtab sex married, x(born=1)

zip: Predicted rates for childs

responden	married	
ts sex	0	1
male	1.0721	2.2151
female	1.6977	2.2531

base x values for count equation:

	sex	married	sibs	born	educ
x=	1.5555556	.45974499	3.6018215	1	13.358834

base z values for binary equation:

	sex	married	sibs	born	educ
z=	1.5555556	.45974499	3.6018215	1	13.358834

Note that we could have separately specified the values of independent variables for the two equations - we would only used that if we used different variables in the two equations.

For foreign-born:

```
. prtab sex married, x(born=2)
zip: Predicted rates for childs
```

```
-----
responde |      married
ns sex   |      0      1
-----+-----
      male | 0.7569  1.8487
      female | 1.3159  1.8912
-----
```

base x values for count equation:

```
      sex      married      sibs      born      educ
x=  1.5555556  .45974499  3.6018215      2  13.358834
```

base z values for binary equation:

```
      sex      married      sibs      born      educ
z=  1.5555556  .45974499  3.6018215      2  13.358834
```

We can also examine changes in predicted rates as well as marginal effects.

```
. prchange
```

```
zip: Changes in Predicted Rate for childs
```

```
      min->max      0->1      +1/2      +sd/2
sex      0.2339      0.5252      0.2212      0.1072
married  0.7951      0.7951      0.8680      0.3761
sibs     2.4221      0.0697      0.0740      0.2221
born    -0.3756     -0.4412     -0.4010     -0.1159
educ    -2.2847     -0.1419     -0.1047     -0.3117
```

```
exp(xb):  2.0117
```

base x values for count equation:

```
      sex married      sibs      born      educ
x=  1.55556  .459745  3.60182  1.09217  13.3588
sd(x)= .496995  .498468  3.00084  .289315  2.97411
```

base z values for binary equation:

```
      sex married      sibs      born      educ
z=  1.55556  .459745  3.60182  1.09217  13.3588
sd(z)= .496995  .498468  3.00084  .289315  2.97411
```

We interpret these results the same way as for regular Poisson model. Note that here prchange does not compute marginal effects. But we can obtain them using mfx compute (this calculation will take a long time - takes a while to calculate standard errors).

```
. mfx compute
```

```
Marginal effects after zip
```

```
      y = predicted number of events (predict)
      = 2.0116755
```

```
-----
variable |      dy/dx      Std. Err.      z      P>|z|      [      95% C.I.      ]      X
-----+-----
      sex | .2137696      .07513      2.85      0.004      .066517      .361022      1.55556
married* | .7950725      .06097     13.04      0.000      .675569      .914576      .459745
      sibs | .074003      .0096      7.71      0.000      .055192      .092814      3.60182
-----
```

born		-.4005967	.11142	-3.60	0.000	-.618976	-.182218	1.09217
educ		-.1047399	.01113	-9.41	0.000	-.126553	-.082927	13.3588

(\*) dy/dx is for discrete change of dummy variable from 0 to 1

Note that all marginal effects are significant - this is because some of the variables had significant coefficients in the count model, and others in "Always zero" model, and marginal effects combined the two to calculate the overall impact of each variable on the expected count. It is evaluated at the mean of each variable with other variables also held at their means; for dummy variables it is evaluated as discrete change in the predicted rate. Unfortunately, because our sex and born variables are not 0-1 variables, mfx compute does not realize they are dummy variables. Therefore, always try to code all dummies as 0-1. An example of using marginal effects can be found in Sarkisian and Gerstel 2004.

We can also examine predicted probabilities using prvalue and prgen. The only difference in using these is that now we will get two probabilities for zero: One is the total probability - either because one is in "Always Zero" group or because they just didn't have their first kid yet. The other one is probability of being in "Always zero" group only. Let's examine these:

```
. prvalue, x(married=0 sex=1 born=1)
zip: Predictions for childs
Predicted rate: 1.07
Predicted probabilities:
Pr(y=0|x,z): 0.6788 Pr(y=1|x): 0.1792
Pr(y=2|x): 0.0961 Pr(y=3|x): 0.0343
Pr(y=4|x): 0.0092 Pr(y=5|x): 0.0020
Pr(y=6|x): 0.0004 Pr(y=7|x): 0.0001
Pr(y=8|x): 0.0000 Pr(y=9|x): 0.0000
Pr(Always0|z): 0.5116
x values for count equation
      sex    married      sibs      born      educ
x=      1          0  3.6018215      1  13.358834
z values for binary equation
      sex    married      sibs      born      educ
z=      1          0  3.6018215      1  13.358834
```

These were predicted probabilities (and the predicted rate!) for average single native-born men. We can see that according to our model 68% of these men don't have kids and most of these men are in always zero group - the probability of being in that group is .51. So the remaining 17% we assume just didn't start having children yet. No let's look at married men:

```
. prvalue, x(married=1 sex=1 born=1)
zip: Predictions for childs
Predicted rate: 2.22
Predicted probabilities:
Pr(y=0|x,z): 0.1282 Pr(y=1|x): 0.2366
Pr(y=2|x): 0.2620 Pr(y=3|x): 0.1935
Pr(y=4|x): 0.1071 Pr(y=5|x): 0.0475
Pr(y=6|x): 0.0175 Pr(y=7|x): 0.0055
Pr(y=8|x): 0.0015 Pr(y=9|x): 0.0004
Pr(Always0|z): 0.0214
x values for count equation
      sex    married      sibs      born      educ
x=      1          1  3.6018215      1  13.358834
z values for binary equation
```

```

z=          sex      married      sibs      born      educ
          1          1  3.6018215      1  13.358834

```

Only 13% of these men are expected to have no kids, and only 2% of them are in always zero group - the remaining 11% just didn't start having kids yet. We can do a similar analysis for women - let's put their results next to each other:

```

. quietly prvalue, x(married=0 sex=2 born=1) save
. prvalue, x(married=1 sex=2 born=1) dif
zip: Change in Predictions for child
Predicted rate: 2.25      Saved: 1.7
      Difference: .555

```

Predicted probabilities:

	Current	Saved	Difference
Pr(y=0 x,z):	0.1106	0.3692	-0.2586
Pr(y=1 x):	0.2353	0.2401	-0.0048
Pr(y=2 x):	0.2651	0.2038	0.0613
Pr(y=3 x):	0.1991	0.1153	0.0838
Pr(y=4 x):	0.1121	0.0489	0.0632
Pr(y=5 x):	0.0505	0.0166	0.0339
Pr(y=6 x):	0.0190	0.0047	0.0143
Pr(y=7 x):	0.0061	0.0011	0.0050
Pr(y=8 x):	0.0017	0.0002	0.0015
Pr(y=9 x):	0.0004	0.0000	0.0004
Pr(Always0 z):	0.0061	0.2278	-0.2216

x values for count equation

	sex	married	sibs	born	educ
Current=	2	1	3.6018215	1	13.358834
Saved=	2	0	3.6018215	1	13.358834
Diff=	0	1	0	0	0

z values for binary equation

	sex	married	sibs	born	educ
Current=	2	1	3.6018215	1	13.358834
Saved=	2	0	3.6018215	1	13.358834
Diff=	0	1	0	0	0

According to our model, 36% of single women don't have kids and 23% never will, while only 11% of married women don't have kids and only 0.6% never will.

We can also use prgen to make graphs like we did for Poisson model - but here again we will have two sets of probabilities for zero counts -total probability of zero and probability of "Always zero." E.g., see Long and Freese p. 282.

We can also adjust our final, best-fitting model to exposure time:

```

. zip childsex married sibs born educ, inflate(sex married sibs born educ)
exposure(reprage)
(31 missing values generated)

```

```

Zero-inflated poisson regression      Number of obs   =      2734
                                       Nonzero obs     =      1946
                                       Zero obs        =       788
Inflation model = logit               LR chi2(5)      =     119.40
Log likelihood = -4334.455             Prob > chi2     =      0.0000

```

```

-----+-----
      childsex |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
childsex      |

```

sex	.0673734	.0319959	2.11	0.035	.0046625	.1300842
married	.0372361	.0329312	1.13	0.258	-.0273079	.10178
sibs	.0213414	.004529	4.71	0.000	.0124647	.0302181
born	-.099738	.0548672	-1.82	0.069	-.2072757	.0077996
educ	-.04122	.0051174	-8.05	0.000	-.0512498	-.0311901
_cons	-1.996286	.1081046	-18.47	0.000	-2.208167	-1.784405
reprage	(exposure)					
-----						
inflate						
sex	-1.258563	.1789565	-7.03	0.000	-1.609311	-.9078144
married	-7.69451	37.75966	-0.20	0.839	-81.70207	66.31305
sibs	-.0533748	.0340675	-1.57	0.117	-.1201459	.0133964
born	.3318979	.3383992	0.98	0.327	-.3313523	.9951481
educ	.1963433	.0342241	5.74	0.000	.1292652	.2634213
_cons	-1.914812	.6732486	-2.84	0.004	-3.234355	-.5952693

Note that the model changed - marriage that seemed so important is no longer significant! Looks like that was just function of age. Sex, siblings, and education predict the count, and sex and education predict the membership in always zero group.

Let's use fitstat to see whether this model with exposure performs better than the model without:

```
. quietly fitstat, save
. quietly zip childs sex married sibs born educ if reprage~=., inflate(sex
married sibs born educ)
```

Note: Here we limit the model without exposure only to those who don't miss data on reprage variable.

```
. fitstat, dif
```

Measures of Fit for zip of childs

	Current	Saved	Difference
Model:	zip	zip	
N:	2734	2734	0
Log-Lik Intercept Only	-4825.719	-4825.719	0.000
Log-Lik Full Model	-4509.577	-4334.455	-175.121
D	9019.153(2722)	8668.911(2722)	350.243(0)
LR	632.285(10)	982.528(10)	350.243(0)
Prob > LR	0.000	0.000	.
McFadden's R2	0.066	0.102	-0.036
McFadden's Adj R2	0.063	0.099	-0.036
ML (Cox-Snell) R2	0.206	0.302	-0.095
Cragg-Uhler(Nagelkerke) R2	0.213	0.311	-0.098
AIC	3.308	3.180	0.128
AIC*n	9043.153	8692.911	350.243
BIC	-12521.451	-12871.693	350.243
BIC'	-553.150	-903.393	350.243
BIC used by Stata	9114.116	8763.873	350.243
AIC used by Stata	9043.153	8692.911	350.243

Difference of 350.243 in BIC' provides very strong support for saved model.

Note: p-value for difference in LR is only valid if models are nested.

We can see very strong support for the model with exposure.

The issue of diagnostics for zero-inflated models:

Unfortunately, many tests and work-around solutions that worked for nbreg and poisson don't work for zip and zinb. One big problem is that zip and zinb cannot be modeled using GLM. We can still test for multicollinearity and use

robust option, but linearity diagnostics and those used to identify outliers and leverage points are not available here. One could test for those using regular poisson or nbreg and then see if suggested fixes (e.g., a transformation or omitted leverage points) appear to improve the corresponding zero-inflated model.

### Zero-truncated models

Sometimes we have count data that have no zeros at all, because we only start accumulating data once at least one count was observed. For example, the length of hospital stay cannot be 0 because we only start observing counts once a person is admitted. In such cases, zero-truncated models, implemented by ztp and ztnb commands, are useful. E.g. say we only have data on the number of children after the person has their first one:

```
. gen childs0=childs
(5 missing values generated)
. replace childs0=. if childs==0
(799 real changes made, 799 to missing)
. ztp childs0 sex married sibs born educ
Zero-truncated Poisson regression
```

	Number of obs	=	1951
	LR chi2(5)	=	168.39
	Prob > chi2	=	0.0000
	Pseudo R2	=	0.0262

Log likelihood = -3129.8812

childs0	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
sex	.0050533	.0341538	0.15	0.882	-.061887	.0719936
married	.0439347	.0344268	1.28	0.202	-.0235405	.11141
sibs	.0283134	.0047432	5.97	0.000	.019017	.0376098
born	-.1934924	.0631899	-3.06	0.002	-.3173423	-.0696426
educ	-.0403873	.0055964	-7.22	0.000	-.0513561	-.0294186
_cons	1.406071	.1183233	11.88	0.000	1.174161	1.63798

```
. ztnb childs0 sex married sibs born educ
Zero-truncated negative binomial regression
```

	Number of obs	=	1951
	LR chi2(5)	=	114.29
Dispersion = mean	Prob > chi2	=	0.0000
Log likelihood = -3128.9162	Pseudo R2	=	0.0179

childs0	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
sex	.0043327	.0352032	0.12	0.902	-.0646644	.0733297
married	.0440371	.0354945	1.24	0.215	-.0255309	.1136051
sibs	.0285975	.0049392	5.79	0.000	.0189169	.0382781
born	-.1951289	.0649357	-3.00	0.003	-.3224005	-.0678573
educ	-.0403866	.0057732	-7.00	0.000	-.0517018	-.0290714
_cons	1.398945	.1221116	11.46	0.000	1.15961	1.638279
/lnalpha	-3.811634	.7616972			-5.304533	-2.318735
alpha	.022112	.0168427			.004969	.098398

Likelihood-ratio test of alpha=0:  $\chi^2(1) = 1.93$  Prob>= $\chi^2 = 0.082$

Note that the results of these models look very similar to those from the count equations of zero-inflated Poisson and NB models.



Examples of count data models:

Van der Burg, Brigitte, Jacques Siegers, and Rudolf Winter-Ebmer. 1998. Gender and Promotion in the Academic Labour Market. *Labour*, 12: 701-713.

Questions to answer about the article:

1. What are the dependent and the independent variables in this analysis?
2. What is reported in Table 1? How can we interpret these results? How do the authors discuss these results in the text?
3. What is presented in Table 2? How can we interpret these results?
4. In addition to what the authors chose to present, how else could they have presented their results?
5. What measures of model fit and model diagnostics are presented? What diagnostics and potential problems did the authors not address?

Sarkisian, Natalia and Naomi Gerstel. 2004. "Explaining the Gender Gap in Help to Parents: The Importance of Employment." *Journal of Marriage and the Family*, 66: 431-451.

Questions to answer about the article:

1. What are the dependent and the independent variables in this analysis?
2. What is reported in Table 1? How can we interpret these results? How do the authors discuss these results in the text?
3. In addition to what the authors chose to present, how else could they have presented their results?
4. What measures of model fit and model diagnostics are presented? What diagnostics and potential problems did the authors not address?