

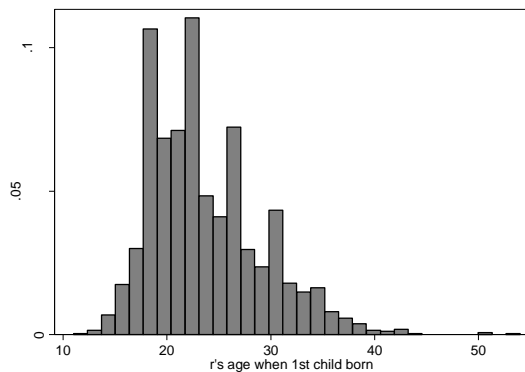
**Sociology 7704: Regression Models for Categorical Data**  
**Instructor: Natasha Sarkisian**

**Preliminary Data Screening**

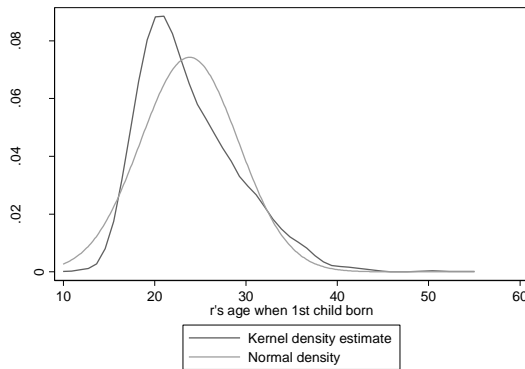
**A. Examining Univariate Normality**

Normality of each of the variables used in your model is not required, but it can often help us prevent further problems (especially heteroscedasticity and multivariate normality violations). Normality of the dependent variable is especially influential. We can examine the distribution graphically:

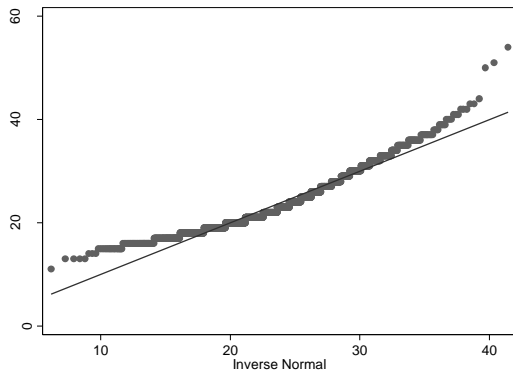
```
. histogram agekdbrn, normal  
(bin=34, start=18, width=2.0882353)
```



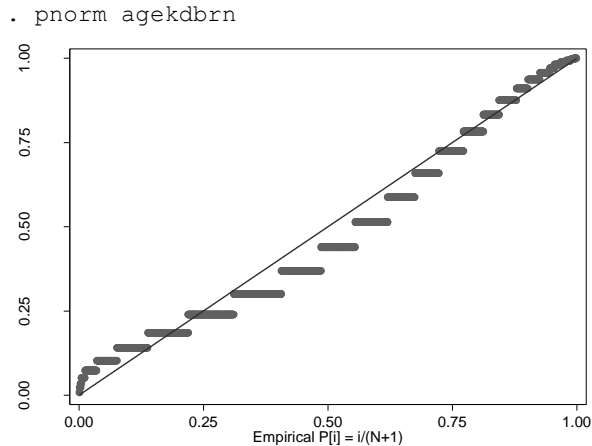
```
. kdensity age, normal
```



```
. qnorm agekdbrn
```



This is a quantile-normal (Q-Q) plot. It plots the quantiles of a variable against the quantiles of a normal distribution. In a perfectly normal distribution, all observations would be on the line, so the closest they are to being on the line, the closer the distribution to being normal. Any large deviations from the straight line indicate problems with normality. Note: this plot has nothing to do with linearity!



This is a standardized normal probability (P-P) plot, it is more sensitive to non-normality in the middle range of data, while qnorm is sensitive to non-normality near the tails.

We can also formally evaluate the distribution of a variable -- i.e., test the hypothesis of normality (with separate tests for skewness and kurtosis) using sktest:

```
. sktest age
```

| Skewness/Kurtosis tests for Normality |              |              |             |                 |
|---------------------------------------|--------------|--------------|-------------|-----------------|
| Variable                              | Pr(Skewness) | Pr(Kurtosis) | adj chi2(2) | joint Prob>chi2 |
| age                                   | 0.000        | 0.000        | .           | 0.0000          |

Here, the dot instead of chi-square value indicates that it's a very large number. This test is very sensitive to sample size, however – with large sample sizes, even small deviations from normality can be identified as statistically significant. But in this case, the graphs also confirmed this conclusion. Next, we'll consider transformations to bring this variable closer to normal. To search for transformations, we can use ladder command:

```
. ladder agekdbnr
```

| Transformation    | formula          | chi2(2) | P(chi2) |
|-------------------|------------------|---------|---------|
| cubic             | agekdbnr^3       | .       | 0.000   |
| square            | agekdbnr^2       | .       | 0.000   |
| identity          | agekdbnr         | .       | 0.000   |
| square-root       | sqrt(agekdbnr)   | .       | 0.000   |
| log               | log(agekdbnr)    | 32.49   | 0.000   |
| reciprocal root   | 1/sqrt(agekdbnr) | 8.57    | 0.014   |
| reciprocal        | 1/agekdbnr       | 14.84   | 0.001   |
| reciprocal square | 1/(agekdbnr^2)   | .       | 0.000   |
| reciprocal cubic  | 1/(agekdbnr^3)   | .       | 0.000   |

Ladder allows you to search for normalizing transformation – the larger the P value, the closer to normal. Typically, square roots, log, and inverse (1/x) transformations normalize right (positive)

skew. Inverse (reciprocal) transforms are “stronger” than logarithmic, which are “stronger” than square roots. For negative skews, we can use square or cubic transformation.

In this output, again, dots instead of chi2 indicate very large numbers. If there is a dot instead of P as well, it means that this specific transformation is not possible because of zeros or negative values. If zeros or negative values preclude a transformation that you think might help, the typical practice is to first add a constant that would get rid of such values (e.g., if you only have zeros but no negative values, you can add 1), and then perform a transformation. In this case, it appears that 1/square root brings the distribution closer to normal.

Note that just as sktest, in large samples the ladder command tests are rather sensitive to non-normalities – often it can be useful to take a random subsample and run ladder command on them to identify the best transformation. (But make sure the sample is not too small; keep it around 150-200 observations.)

```
. ladder age
```

| Transformation    | formula      | chi2 (2) | P (chi2) |
|-------------------|--------------|----------|----------|
| cubic             | age^3        | .        | 0.000    |
| square            | age^2        | .        | 0.000    |
| identity          | age          | .        | 0.000    |
| square-root       | sqrt (age)   | .        | 0.000    |
| log               | log (age)    | .        | 0.000    |
| reciprocal root   | 1/sqrt (age) | .        | 0.000    |
| reciprocal        | 1/age        | .        | 0.000    |
| reciprocal square | 1/ (age^2)   | .        | 0.000    |
| reciprocal cubic  | 1/ (age^3)   | .        | 0.000    |

It’s not normal and none of the transformations seem to help. If your sample size is large, everything will be significantly different from normal, so you should either rely on graphical examination (gladder) or randomly select a subsample of your dataset and do this type of analysis for that subsample. We can use sample command to take a 5% random sample from the data. We first “preserve” the dataset so that we can bring the rest of observations back after we are done with ladder, and then sample:

```
. preserve
```

```
. sample 5
(2627 observations deleted)
```

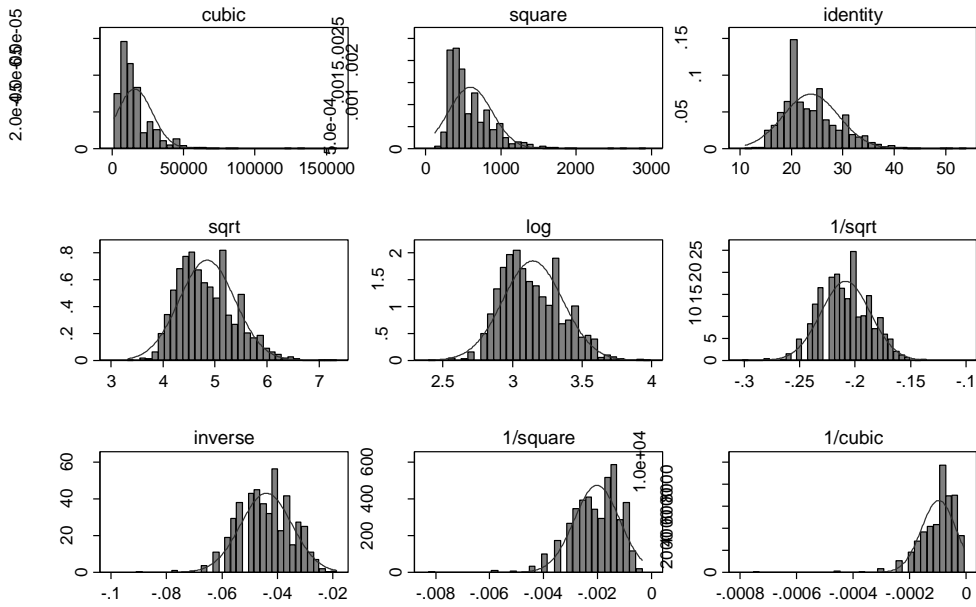
```
. ladder age
```

| Transformation    | formula      | chi2 (2) | P (chi2) |
|-------------------|--------------|----------|----------|
| cubic             | age^3        | 40.17    | 0.000    |
| square            | age^2        | 25.53    | 0.000    |
| identity          | age          | 10.53    | 0.005    |
| square-root       | sqrt (age)   | 6.81     | 0.033    |
| log               | log (age)    | 5.99     | 0.050    |
| reciprocal root   | 1/sqrt (age) | 4.78     | 0.091    |
| reciprocal        | 1/age        | 8.23     | 0.016    |
| reciprocal square | 1/ (age^2)   | 32.80    | 0.000    |
| reciprocal cubic  | 1/ (age^3)   | 63.69    | 0.000    |

Note that now it's much more clear which transformations bring this variable the closest to normal.

`. restore`  
 Restore command restores our original dataset (as it was when we ran preserve).  
 Let's examine transformations for agekdbrn graphically as well:

`. gladder agekdbrn`

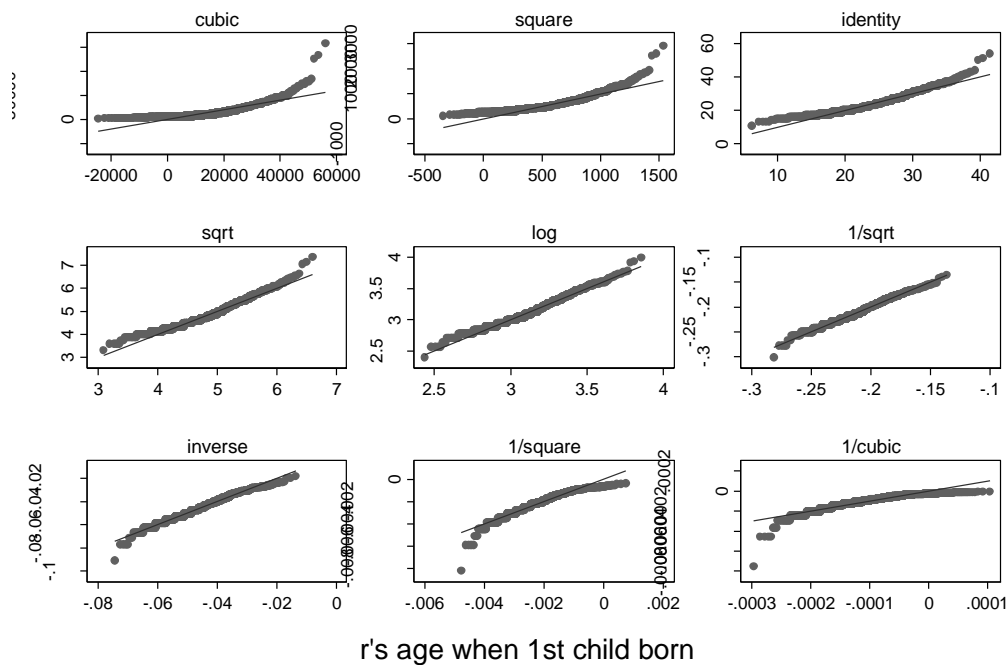


r's age when 1st child born

Histograms by transformation

Same using quantile-normal plots:

`. qladder agekdbrn`



Quantile-Normal plots by transformation

Let's attempt to use this transformation in our regression model:

```
. gen agekdbrnrr=1/(sqrt(agekdbrn))
(810 missing values generated)
```

```
. reg agekdbrnrr educ born sex mapres80 age
```

| Source   | SS         | df   | MS         | Number of obs = | 1089   |
|----------|------------|------|------------|-----------------|--------|
| Model    | .107910937 | 5    | .021582187 | F( 5, 1083) =   | 54.00  |
| Residual | .432834805 | 1083 | .000399663 | Prob > F =      | 0.0000 |
| Total    | .540745743 | 1088 | .000497009 | R-squared =     | 0.1996 |
|          |            |      |            | Adj R-squared = | 0.1959 |
|          |            |      |            | Root MSE =      | .01999 |

|  | agekdbrnrr | Coef.     | Std. Err. | t      | P> t  | [95% Conf. Interval] |
|--|------------|-----------|-----------|--------|-------|----------------------|
|  | educ       | -.0026108 | .0002316  | -11.27 | 0.000 | -.0030652 - .0021564 |
|  | born       | -.0075379 | .0023762  | -3.17  | 0.002 | -.0122004 - .0028755 |
|  | sex        | .0098921  | .0012561  | 7.88   | 0.000 | .0074274 .0123568    |
|  | mapres80   | -.0001494 | .000049   | -3.05  | 0.002 | -.0002455 - .0000533 |
|  | age        | -.0002532 | .0000409  | -6.19  | 0.000 | -.0003336 - .0001729 |
|  | _cons      | .2535923  | .0051683  | 49.07  | 0.000 | .2434514 .2637332    |

Overall, transformations should be used sparsely - always consider ease of model interpretation as well. Here, our transformation made interpretation more complicated. It is also important to check that we did not introduce any nonlinearities by this transformation – we'll deal with that issue soon.

If a variable contains zero or negative values, you need to add a constant to it before looking for transformations (such that all values of the variable become larger than zero). For example:

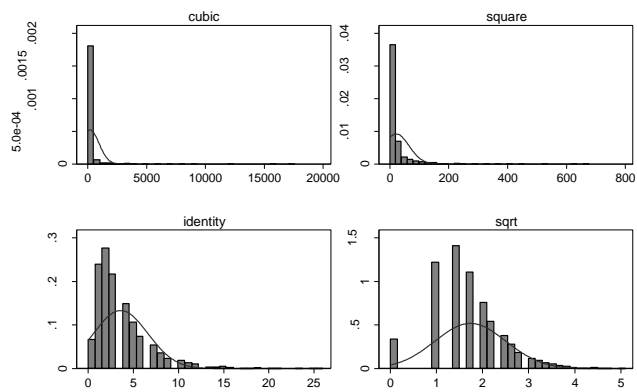
```
. sum sibs
```

| Variable | Obs  | Mean     | Std. Dev. | Min | Max |
|----------|------|----------|-----------|-----|-----|
| sibs     | 2756 | 3.599419 | 2.997262  | 0   | 26  |

. ladder sibs

| Transformation  | formula      | chi2(2) | P(chi2) |
|-----------------|--------------|---------|---------|
| cubic           | sibs^3       | .       | .       |
| square          | sibs^2       | .       | .       |
| identity        | sibs         | .       | 0.000   |
| square root     | sqrt(sibs)   | 64.41   | 0.000   |
| log             | log(sibs)    | .       | .       |
| 1/(square root) | 1/sqrt(sibs) | .       | .       |
| inverse         | 1/sibs       | .       | .       |
| 1/square        | 1/(sibs^2)   | .       | .       |
| 1/cubic         | 1/(sibs^3)   | .       | .       |

. gladder sibs



number of brothers and sisters

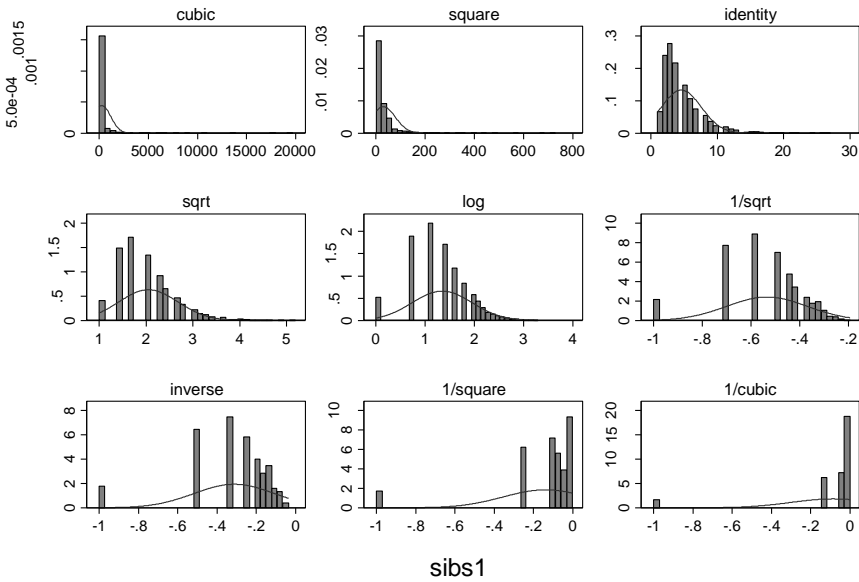
Histograms by transformation

. gen sibs1=sibs+1  
(9 missing values generated)

. ladder sibs1

| Transformation  | formula       | chi2(2) | P(chi2) |
|-----------------|---------------|---------|---------|
| cubic           | sibs1^3       | .       | .       |
| square          | sibs1^2       | .       | .       |
| identity        | sibs1         | .       | 0.000   |
| square root     | sqrt(sibs1)   | .       | 0.000   |
| log             | log(sibs1)    | 0.48    | 0.787   |
| 1/(square root) | 1/sqrt(sibs1) | .       | 0.000   |
| inverse         | 1/sibs1       | .       | 0.000   |
| 1/square        | 1/(sibs1^2)   | .       | .       |
| 1/cubic         | 1/(sibs1^3)   | .       | .       |

. gladder sibs1



Histograms by transformation

If a variable is negatively skewed, you might have an easier time finding a transformation for it after reversing it. For this example, we will generate a scale of happiness that's the reverse of an unhappiness scale and examine both distributions:

```
. tab1 happy7 satjob7 satfam7
```

-> tabulation of happy7

| happy or unhappy on the whole | Freq. | Percent | Cum.   |
|-------------------------------|-------|---------|--------|
| completely happy              | 141   | 12.16   | 12.16  |
| very happy                    | 510   | 43.97   | 56.12  |
| fairly happy                  | 391   | 33.71   | 89.83  |
| neither happy nor unhappy     | 69    | 5.95    | 95.78  |
| fairly unhappy                | 32    | 2.76    | 98.53  |
| very unhappy                  | 16    | 1.38    | 99.91  |
| completely unhappy            | 1     | 0.09    | 100.00 |
| Total                         | 1,160 | 100.00  |        |

-> tabulation of satjob7

| job satisfaction in general        | Freq. | Percent | Cum.   |
|------------------------------------|-------|---------|--------|
| completely satisfied               | 127   | 15.49   | 15.49  |
| very satisfied                     | 289   | 35.24   | 50.73  |
| fairly satisfied                   | 264   | 32.20   | 82.93  |
| neither satisfied nor dissatisfied | 53    | 6.46    | 89.39  |
| fairly dissatisfied                | 47    | 5.73    | 95.12  |
| very dissatisfied                  | 29    | 3.54    | 98.66  |
| completely dissatisfied            | 11    | 1.34    | 100.00 |
| Total                              | 820   | 100.00  |        |

-> tabulation of satfam7

| family satisfaction in general | Freq. | Percent | Cum. |
|--------------------------------|-------|---------|------|
|--------------------------------|-------|---------|------|

|                                    |       |        |        |
|------------------------------------|-------|--------|--------|
| completely satisfied               | 265   | 23.08  | 23.08  |
| very satisfied                     | 467   | 40.68  | 63.76  |
| fairly satisfied                   | 286   | 24.91  | 88.68  |
| neither satisfied nor dissatisfied | 70    | 6.10   | 94.77  |
| fairly dissatisfied                | 31    | 2.70   | 97.47  |
| very dissatisfied                  | 20    | 1.74   | 99.22  |
| completely dissatisfied            | 9     | 0.78   | 100.00 |
| Total                              | 1,148 | 100.00 |        |

```
. alpha happy7 satjob7 satfam7
```

```
Test scale = mean(unstandardized items)
```

```
Average interitem covariance: .525359
Number of items in the scale: 3
Scale reliability coefficient: 0.6732
```

```
. egen unhappiness=rowmean(happy7 satjob7 satfam7)
(1600 missing values generated)
```

```
. sum unhappiness
```

| Variable    | Obs  | Mean     | Std. Dev. | Min | Max |
|-------------|------|----------|-----------|-----|-----|
| unhappiness | 1165 | 2.469814 | .9298462  | 1   | 7   |

To reverse the scale, we add its maximum and its minimum and subtract the original scale from that:

```
. gen happiness=r(max)+r(min)-unhappiness
(1600 missing values generated)
```

```
. sum happiness
```

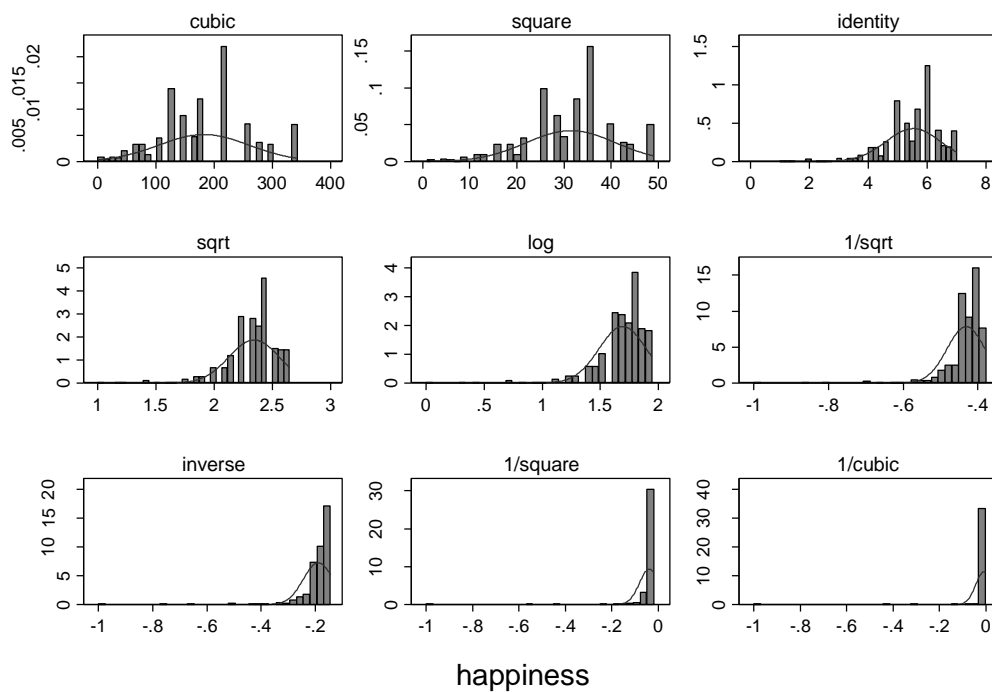
| Variable  | Obs  | Mean     | Std. Dev. | Min | Max |
|-----------|------|----------|-----------|-----|-----|
| happiness | 1165 | 5.530186 | .9298462  | 1   | 7   |

```
. ladder happiness
```

| Transformation  | formula          | chi2 (2) | P(chi2) |
|-----------------|------------------|----------|---------|
| cubic           | happin~s^3       | 11.17    | 0.004   |
| square          | happin~s^2       | 15.55    | 0.000   |
| identity        | happin~s         | .        | 0.000   |
| square root     | sqrt(happin~s)   | .        | 0.000   |
| log             | log(happin~s)    | .        | 0.000   |
| 1/(square root) | 1/sqrt(happin~s) | .        | 0.000   |
| inverse         | 1/happin~s       | .        | .       |
| 1/square        | 1/(happin~s^2)   | .        | .       |
| 1/cubic         | 1/(happin~s^3)   | .        | .       |

```
. gladder happiness
```



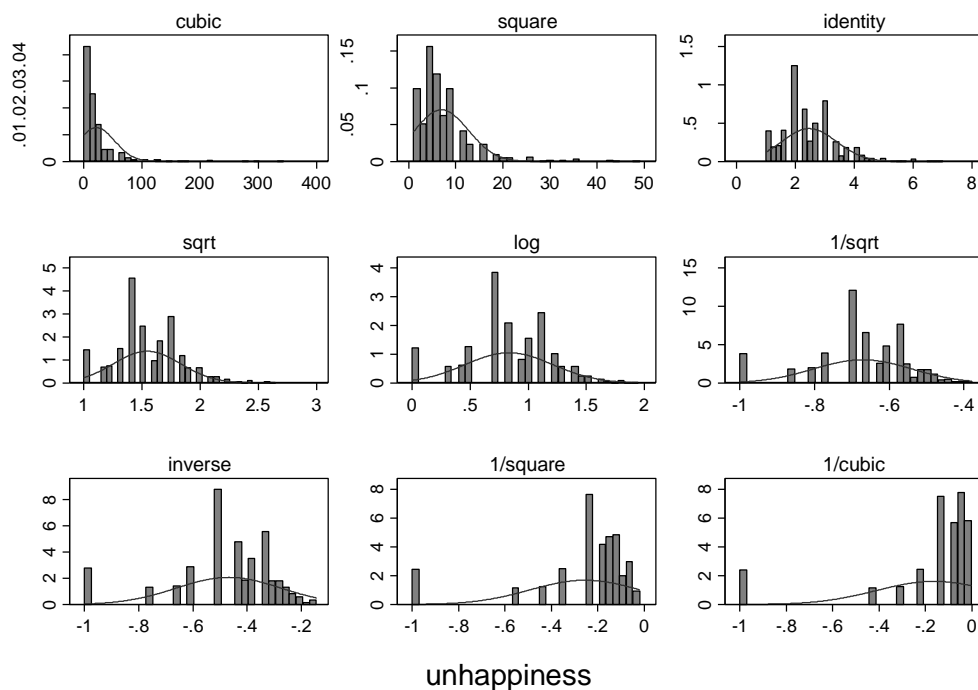


Histograms by transformation

. ladder unhappiness

| Transformation  | formula                 | chi2(2) | P(chi2) |
|-----------------|-------------------------|---------|---------|
| cubic           | $unhapp\sim s^3$        | .       | 0.000   |
| square          | $unhapp\sim s^2$        | .       | 0.000   |
| identity        | $unhapp\sim s$          | .       | 0.000   |
| square root     | $\sqrt{unhapp\sim s}$   | 27.32   | 0.000   |
| log             | $\log(unhapp\sim s)$    | 13.42   | 0.001   |
| 1/(square root) | $1/\sqrt{unhapp\sim s}$ | .       | 0.000   |
| inverse         | $1/unhapp\sim s$        | .       | 0.000   |
| 1/square        | $1/(unhapp\sim s^2)$    | .       | 0.000   |
| 1/cubic         | $1/(unhapp\sim s^3)$    | .       | 0.000   |

. gladder unhappiness



Histograms by transformation

We might want to use log, but if we want the interpretation to be about happiness, we will reverse it again after transforming:

```
. gen unhappylog=log(unhappiness)
(1600 missing values generated)
```

```
. sum unhappylog
```

| Variable   | Obs  | Mean     | Std. Dev. | Min | Max     |
|------------|------|----------|-----------|-----|---------|
| unhappylog | 1165 | .8345584 | .3790659  | 0   | 1.94591 |

```
. gen unhappylogr=r(max)+r(min)-unhappylog
(1600 missing values generated)
```

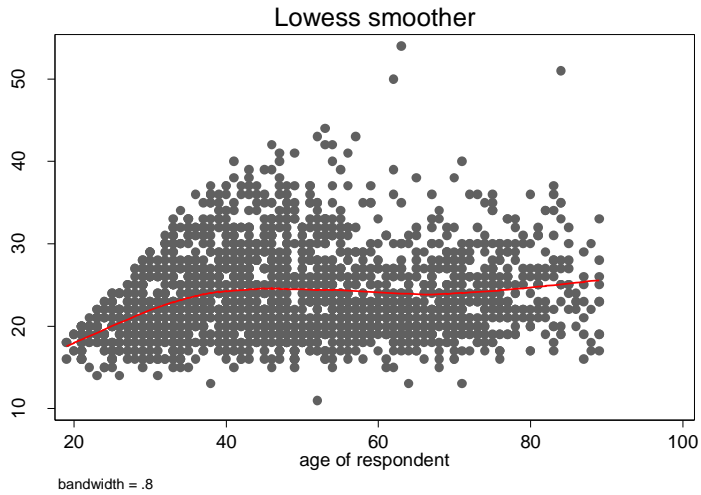
```
. sum unhappylogr
```

| Variable    | Obs  | Mean     | Std. Dev. | Min | Max     |
|-------------|------|----------|-----------|-----|---------|
| unhappylogr | 1165 | 1.111352 | .3790659  | 0   | 1.94591 |

## B. Examining bivariate linearity

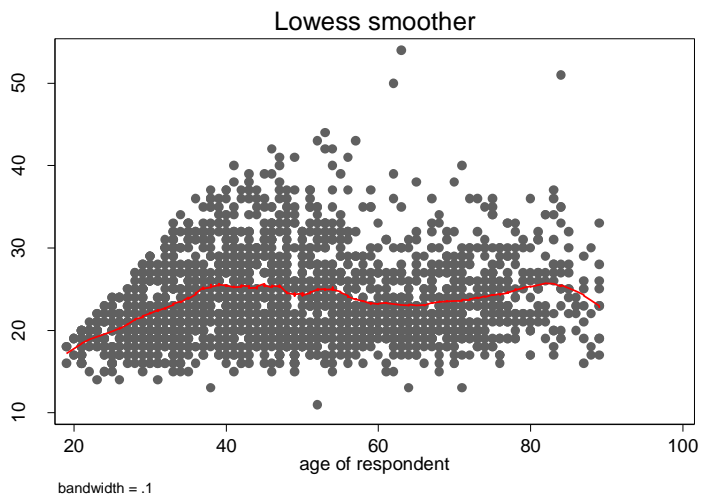
Before you run a regression, it's a good idea to examine your variables one at a time as indicated before, but we should also examine the relationship of each independent variable to the dependent to assess its linearity. A good tool for such an examination is *lowess* – i.e., a scatterplot with a locally weighted regression line going through it (here, it is based on means, but we can also do it using medians):

```
. lowess agekdbrn age
```



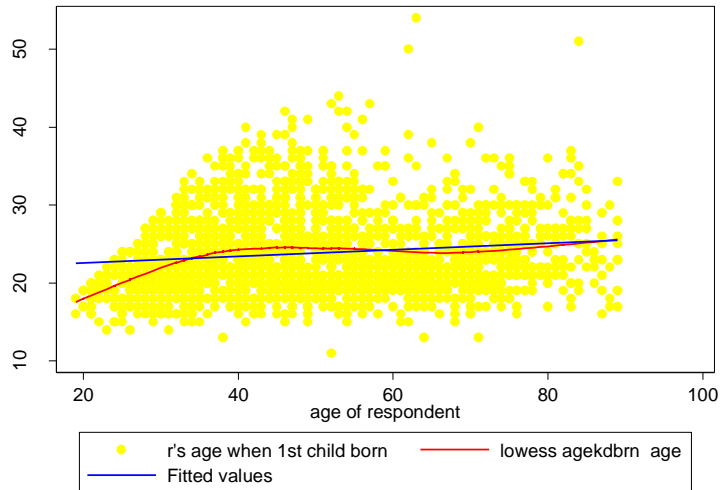
We can change bandwidth to make the curve less smooth (decrease the number) or smoother (increase the number):

```
. lowess agekdbnr age, bwidth(.1)
```



We can also add a regression line to see the difference better:

```
. scatter agekdbnr age, mcolor(yellow) || lowess agekdbnr age, lcolor(red) || lfit agekdbnr age, lcolor(blue)
```

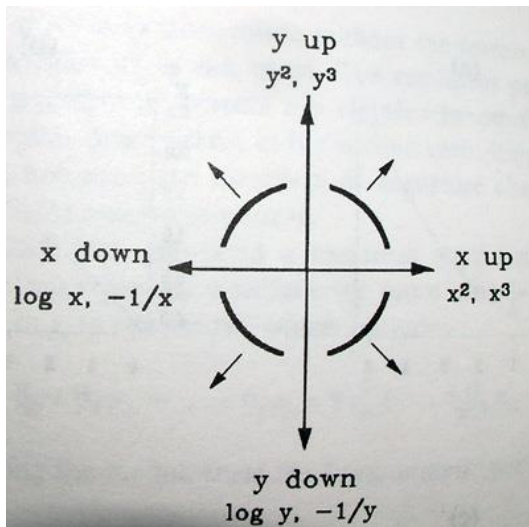


Based on lowess plots, we conclude that the relationship between age and agekdbrn is not linear and we need to address that.

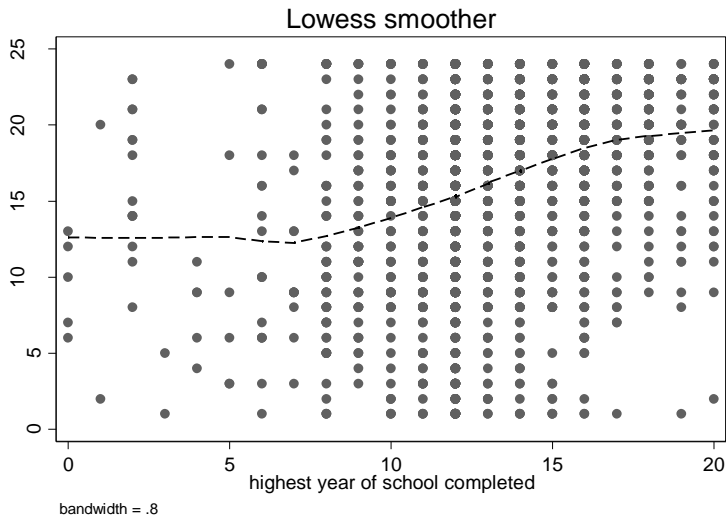
**Remedies for nonlinearity problems:**

When we find a nonlinear relationship, we usually try to find a transformation to linearize it, although sometimes we may choose to break up the corresponding independent variable into a series of dummies instead.

1. Monotone nonlinear relationship. Power transformations can be used to linearize relationships if strong monotone nonlinearities are found. The following chart gives suggestions for transformations when the curve looks a certain way:



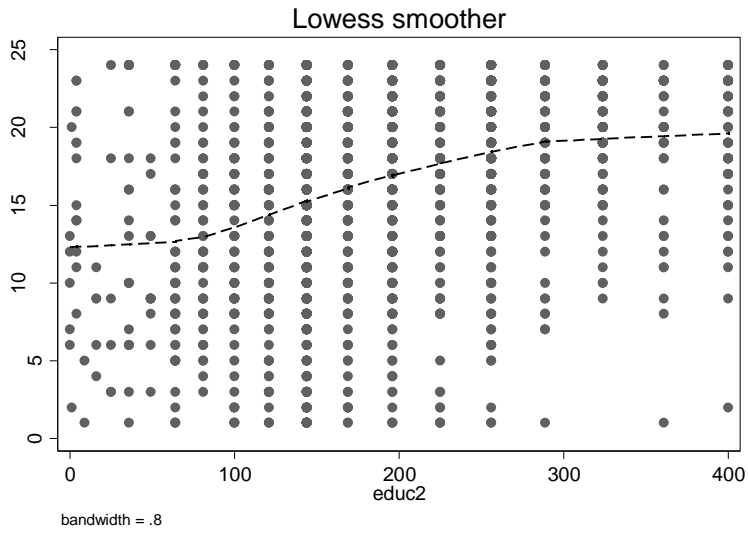
. lowess income98 educ



Either a square of X (educ) or a log of Y (income) should fix this.

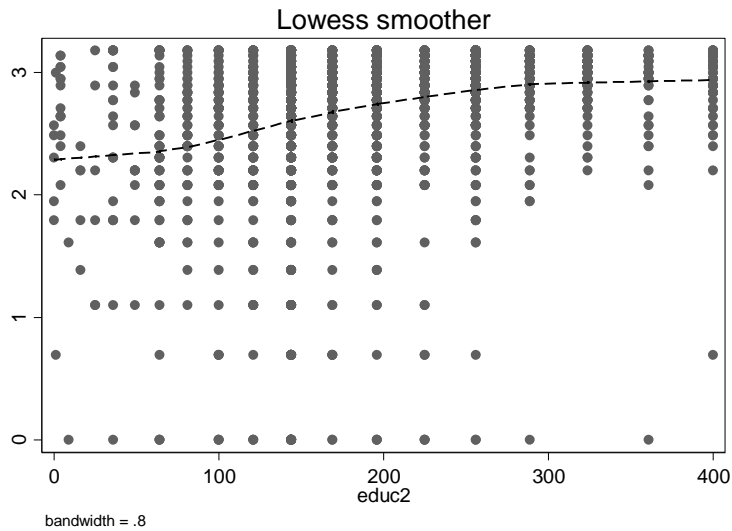
```
. gen educ2=educ^2
(12 missing values generated)
```

```
. lowess income98 educ2
```



```
. gen income98lg=log(income98)
(121 missing values generated)
```

```
. lowess income98lg educ2
```



**2. Nonmonotone relationship.** For non-monotone relationships (e.g. parabola or cubic), use polynomial functions of the variable, e.g. age and age squared, etc. The pictures above for age and agekdbrn relationship would suggest that we might want to add a cubic term for age as well as a squared term. It is important, however, to attempt to maintain simplicity and interpretability of the results when doing transformations. So let's try squared term. We want to enter both age and age squared into our regression model. But using age and age squared in the model at the same time will create multicollinearity because the two variables have a strong relationship—to avoid that, we have to mean-center age prior to generating a square and a cube. That is, whenever we plan to use more than a single term for the same variable in our regression model, always mean-center (i.e., if you just plan to use age squared without age, like we did for educ in the example above, then you don't need to mean center, but if we wanted to use both educ and educ2, we'd have to mean-center educ and only then generate educ2).

For example, without mean-centering:

```
. gen age2=age^2
(14 missing values generated)

. reg agekdbrn educ born sex mapres80 age age2
```

| Source   | SS         | df        | MS         | Number of obs = 1089 |                      |           |
|----------|------------|-----------|------------|----------------------|----------------------|-----------|
| Model    | 6138.53315 | 6         | 1023.08886 | F( 6, 1082)          | =                    | 44.22     |
| Residual | 25034.1298 | 1082      | 23.1369037 | Prob > F             | =                    | 0.0000    |
| -----    |            |           |            | R-squared            | =                    | 0.1969    |
| Total    | 31172.663  | 1088      | 28.6513447 | Adj R-squared        | =                    | 0.1925    |
| -----    |            |           |            | Root MSE             | =                    | 4.8101    |
| agekdbrn | Coef.      | Std. Err. | t          | P> t                 | [95% Conf. Interval] |           |
| educ     | .5678949   | .0569661  | 9.97       | 0.000                | .4561184             | .6796713  |
| born     | 1.567736   | .5723843  | 2.74       | 0.006                | .4446266             | 2.690844  |
| sex      | -2.140989  | .3028244  | -7.07      | 0.000                | -2.735179            | -1.546799 |
| mapres80 | .0332034   | .0117896  | 2.82       | 0.005                | .0100704             | .0563364  |
| age      | .2808181   | .055909   | 5.02       | 0.000                | .1711158             | .3905203  |
| age2     | -.0022448  | .0005551  | -4.04      | 0.000                | -.003334             | -.0011556 |
| _cons    | 8.92424    | 1.643755  | 5.43       | 0.000                | 5.698932             | 12.14955  |

```

. reg agekdbrn educ born sex mapres80 age age2, beta
Source |          SS          df          MS          Number of obs =      1089
-----+-----+-----+-----+-----+-----+-----+-----
Model |    6138.53315         6    1023.08886          F( 6, 1082) =     44.22
Residual |   25034.1298    1082    23.1369037          Prob > F      =     0.0000
-----+-----+-----+-----+-----+-----+-----
Total |   31172.663    1088    28.6513447          R-squared      =     0.1969
                                          Adj R-squared  =     0.1925
                                          Root MSE     =     4.8101

-----+-----+-----+-----+-----+-----+-----
agekdbrn |          Coef.      Std. Err.      t      P>|t|          Beta
-----+-----+-----+-----+-----+-----+-----
educ |    .5678949      .0569661      9.97   0.000      .2884756
born |    1.567736      .5723843      2.74   0.006      .0751117
sex |   -2.140989      .3028244     -7.07   0.000     -.1937892
mapres80 |   .0332034      .0117896      2.82   0.005      .080348
age |    .2808181      .055909      5.02   0.000      .790523
age2 |   -.0022448      .0005551     -4.04   0.000     -.637722
_cons |    8.92424      1.643755      5.43   0.000      .

```

Note that age and age2 have high betas with opposite signs -- that's one indication of multicollinearity. Often when high degree of multicollinearity is present, we would also observe high standard errors. In fact, when reading published research using OLS, pay attention to standard errors -- if they are high relative to the size of the coefficient itself, it's a reason for a concern about possible multicollinearity. Let's check our suspicion using VIFs (Variance Inflation Factors):

```

. vif
Variable |          VIF          1/VIF
-----+-----+-----+-----
age2 |    33.51      0.029845
age |    33.37      0.029963
educ |     1.13      0.886374
mapres80 |     1.10      0.911906
born |     1.01      0.986930
sex |     1.01      0.987914
-----+-----+-----+-----
Mean VIF |    11.86

```

Indeed, high degree of multicollinearity. But luckily, we can avoid it. When including variables that are generated using other variables already in the model (as in this case, or when we want to enter a product of two variables to model an interaction term), we should first mean-center the variable (only if it is continuous; don't mean-center dichotomous variables!). That's how we'd do it in this case:

```

. sum age
Variable |          Obs          Mean      Std. Dev.      Min      Max
-----+-----+-----+-----+-----+-----
age |    2751      46.28281      17.37049      18      89

. gen agemean=age-r(mean)
(14 missing values generated)
. gen agemean2=agemean^2
(14 missing values generated)

. reg agekdbrn educ born sex mapres80 agemean agemean2, beta
Source |          SS          df          MS          Number of obs =      1089
-----+-----+-----+-----+-----+-----+-----
Model |    6138.53316         6    1023.08886          F( 6, 1082) =     44.22
Residual |   25034.1298    1082    23.1369037          Prob > F      =     0.0000
-----+-----+-----+-----+-----+-----+-----
Total |   31172.663    1088    28.6513447          R-squared      =     0.1969
                                          Adj R-squared  =     0.1925
                                          Root MSE     =     4.8101

-----+-----+-----+-----+-----+-----+-----
agekdbrn |          Coef.      Std. Err.      t      P>|t|          Beta
-----+-----+-----+-----+-----+-----+-----

```

```
-----+-----
```

|          |  |           |          |       |       |           |
|----------|--|-----------|----------|-------|-------|-----------|
| educ     |  | .5678949  | .0569661 | 9.97  | 0.000 | .2884756  |
| born     |  | 1.567736  | .5723843 | 2.74  | 0.006 | .0751117  |
| sex      |  | -2.140989 | .3028244 | -7.07 | 0.000 | -.1937892 |
| mapres80 |  | .0332034  | .0117896 | 2.82  | 0.005 | .080348   |
| agemean  |  | .0730284  | .0105054 | 6.95  | 0.000 | .2055801  |
| agemean2 |  | -.0022448 | .0005551 | -4.04 | 0.000 | -.1209343 |
| _cons    |  | 17.11274  | 1.126117 | 15.20 | 0.000 | .         |

```
-----+-----
```

```
. vif
```

| Variable | VIF  | 1/VIF    |
|----------|------|----------|
| agemean2 | 1.20 | 0.829918 |
| agemean  | 1.18 | 0.848643 |
| educ     | 1.13 | 0.886374 |
| mapres80 | 1.10 | 0.911906 |
| born     | 1.01 | 0.986930 |
| sex      | 1.01 | 0.987914 |

```
-----+-----
```

|          |  |      |
|----------|--|------|
| Mean VIF |  | 1.11 |
|----------|--|------|

We can see that the multicollinearity problem has been solved. We also note that the squared term is significant. To better understand what this means substantively, we'll generate a graph:

```
. adjust educ born sex mapres80 if e(sample), gen(pred1)
```

```
-----+-----
```

Dependent variable: agekdbrn      Command: regress  
Created variable: pred1  
Variables left as is: age, age2  
Covariates set to mean: educ = 13.316804, born = 1.0707071, sex = 1.6244261, mapres80 = 39.440773

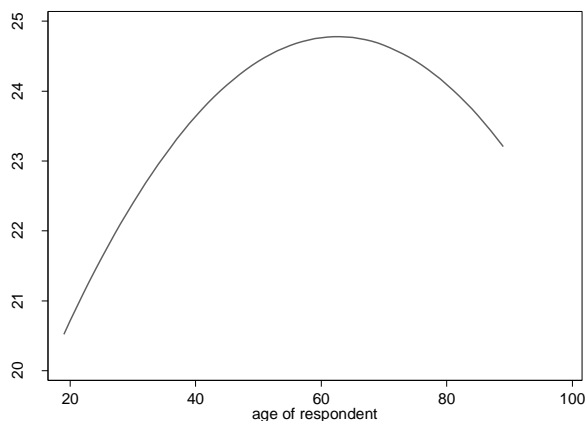
```
-----+-----
```

| All |  | xb      |
|-----|--|---------|
|     |  | 23.6648 |

```
-----+-----
```

Key: xb = Linear Prediction

```
. line pred1 age, sort
```



This doesn't quite replicate what we saw on lowess plot, so the relationship of age and agekdbrn is likely still misspecified. Let's try cube:

```
. gen agemean3=agemean^3  

(14 missing values generated)
```



```

. reg agekdbnr educ born sex mapres80 agemean agemean2 agemean3
-----+-----
Source |      SS      df      MS              Number of obs =    1089
-----+-----
Model |  7554.31674    7  1079.18811             F( 7, 1081) =    49.39
Residual | 23618.3463  1081  21.8486089             Prob > F      =    0.0000
-----+-----
Total | 31172.663  1088  28.6513447             R-squared     =    0.2423
                                           Adj R-squared =    0.2374
                                           Root MSE    =    4.6742
-----+-----
agekdbnr |      Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----
educ |    .581195   .055382    10.49  0.000   .4725265   .6898634
born |    1.292907  .5572673    2.32  0.021   .1994591   2.386355
sex |   -2.117214  .2942876   -7.19  0.000  -2.694654  -1.539774
mapres80 | .0349051  .0114586    3.05  0.002   .0124215   .0573887
agemean | -.0424837  .0176105   -2.41  0.016  -.0770384  -.007929
agemean2 | -.0059131  .0007061   -8.37  0.000  -.0072987  -.0045275
agemean3 | .0002359  .0000293    8.05  0.000   .0001784   .0002934
 _cons |  17.58535   1.09589   16.05  0.000  15.43504  19.73566
-----+-----

```

```

. adjust educ born sex mapres80 if e(sample), gen(pred2)
-----+-----

```

```

Dependent variable: agekdbnr      Command: regress
Created variable: pred2
Variables left as is: agemean, agemean2, agemean3
Covariates set to mean: educ = 13.316804, born = 1.0707071, sex = 1.6244261, mapres80
= 39.440771
-----+-----

```

```

All |      xb
-----+-----
|    23.6648
-----+-----

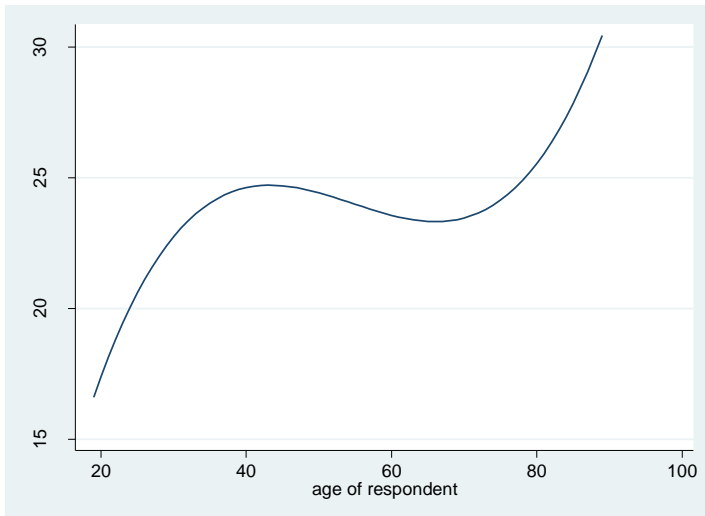
```

Key: xb = Linear Prediction

```

. line pred2 age, sort

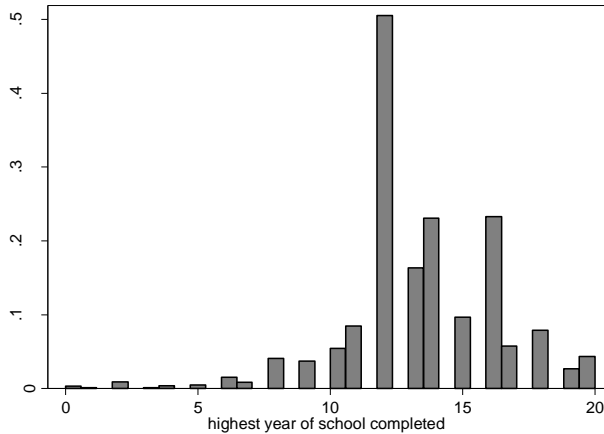
```



### C. Screening for Univariate and Bivariate Outliers

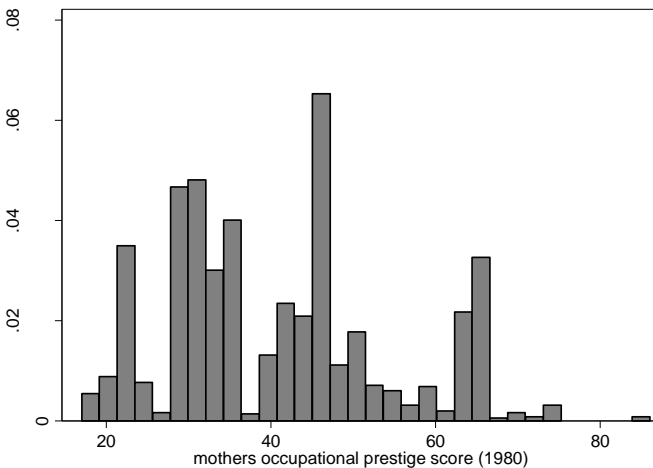
We usually start identifying potential outliers when conducting univariate and bivariate examination of the data. For example, when examining the distribution of educ, we would be concerned about those with very few years of education:

```
. histogram educ
```



When examining the distribution of mother's prestige, we'd be concerned about those with very high values:

```
. histogram mapres80
```



Such observations are likely high leverage points and we might want to deal with them early on, oftentimes by topcoding or bottomcoding:

```
. tab educ
highest |
year of |
school  |
completed |      Freq.      Percent      Cum.
-----+-----
      0 |           5       0.18       0.18
      1 |           2       0.07       0.25
      2 |          15       0.54       0.80
      3 |           2       0.07       0.87
```

|       |  |       |        |        |
|-------|--|-------|--------|--------|
| 4     |  | 6     | 0.22   | 1.09   |
| 5     |  | 8     | 0.29   | 1.38   |
| 6     |  | 25    | 0.91   | 2.29   |
| 7     |  | 14    | 0.51   | 2.80   |
| 8     |  | 66    | 2.40   | 5.19   |
| 9     |  | 60    | 2.18   | 7.37   |
| 10    |  | 88    | 3.20   | 10.57  |
| 11    |  | 137   | 4.98   | 15.55  |
| 12    |  | 818   | 29.71  | 45.26  |
| 13    |  | 265   | 9.63   | 54.89  |
| 14    |  | 374   | 13.59  | 68.47  |
| 15    |  | 157   | 5.70   | 74.17  |
| 16    |  | 377   | 13.69  | 87.87  |
| 17    |  | 93    | 3.38   | 91.25  |
| 18    |  | 128   | 4.65   | 95.90  |
| 19    |  | 43    | 1.56   | 97.46  |
| 20    |  | 70    | 2.54   | 100.00 |
| ----- |  |       |        |        |
| Total |  | 2,753 | 100.00 |        |

```
. gen educb=educ
(12 missing values generated)
```

```
. drop educb
```

```
. gen educb7=educ
(12 missing values generated)
```

```
. replace educb7=7 if educ<7
(63 real changes made)
```

```
. tab educb7
```

| educb7 |  | Freq. | Percent | Cum.   |
|--------|--|-------|---------|--------|
| 7      |  | 77    | 2.80    | 2.80   |
| 8      |  | 66    | 2.40    | 5.19   |
| 9      |  | 60    | 2.18    | 7.37   |
| 10     |  | 88    | 3.20    | 10.57  |
| 11     |  | 137   | 4.98    | 15.55  |
| 12     |  | 818   | 29.71   | 45.26  |
| 13     |  | 265   | 9.63    | 54.89  |
| 14     |  | 374   | 13.59   | 68.47  |
| 15     |  | 157   | 5.70    | 74.17  |
| 16     |  | 377   | 13.69   | 87.87  |
| 17     |  | 93    | 3.38    | 91.25  |
| 18     |  | 128   | 4.65    | 95.90  |
| 19     |  | 43    | 1.56    | 97.46  |
| 20     |  | 70    | 2.54    | 100.00 |
| -----  |  |       |         |        |
| Total  |  | 2,753 | 100.00  |        |

```
. sum mapres80
```

| Variable |  | Obs  | Mean     | Std. Dev. | Min | Max |
|----------|--|------|----------|-----------|-----|-----|
| mapres80 |  | 1619 | 40.96912 | 13.63189  | 17  | 86  |

```
. tab mapres80
```

```
mothers |
occupationa |
l prestige |
```

| score<br>(1980) | Freq. | Percent | Cum.   |
|-----------------|-------|---------|--------|
| 17              | 14    | 0.86    | 0.86   |
| 19              | 5     | 0.31    | 1.17   |
| 20              | 22    | 1.36    | 2.53   |
| 21              | 9     | 0.56    | 3.09   |
| 22              | 39    | 2.41    | 5.50   |
| 23              | 83    | 5.13    | 10.62  |
| 24              | 14    | 0.86    | 11.49  |
| 25              | 13    | 0.80    | 12.29  |
| 26              | 3     | 0.19    | 12.48  |
| 27              | 3     | 0.19    | 12.66  |
| 28              | 125   | 7.72    | 20.38  |
| 29              | 38    | 2.35    | 22.73  |
| 30              | 28    | 1.73    | 24.46  |
| 31              | 53    | 3.27    | 27.73  |
| 32              | 87    | 5.37    | 33.11  |
| 33              | 57    | 3.52    | 36.63  |
| 34              | 48    | 2.96    | 39.59  |
| 35              | 63    | 3.89    | 43.48  |
| 36              | 77    | 4.76    | 48.24  |
| 37              | 1     | 0.06    | 48.30  |
| 38              | 4     | 0.25    | 48.55  |
| 39              | 16    | 0.99    | 49.54  |
| 40              | 30    | 1.85    | 51.39  |
| 41              | 5     | 0.31    | 51.70  |
| 42              | 77    | 4.76    | 56.45  |
| 43              | 21    | 1.30    | 57.75  |
| 44              | 39    | 2.41    | 60.16  |
| 45              | 13    | 0.80    | 60.96  |
| 46              | 160   | 9.88    | 70.85  |
| 47              | 68    | 4.20    | 75.05  |
| 48              | 9     | 0.56    | 75.60  |
| 49              | 30    | 1.85    | 77.46  |
| 50              | 2     | 0.12    | 77.58  |
| 51              | 60    | 3.71    | 81.28  |
| 52              | 19    | 1.17    | 82.46  |
| 53              | 6     | 0.37    | 82.83  |
| 54              | 10    | 0.62    | 83.45  |
| 55              | 11    | 0.68    | 84.13  |
| 56              | 4     | 0.25    | 84.37  |
| 57              | 7     | 0.43    | 84.81  |
| 59              | 8     | 0.49    | 85.30  |
| 60              | 16    | 0.99    | 86.29  |
| 61              | 7     | 0.43    | 86.72  |
| 63              | 2     | 0.12    | 86.84  |
| 64              | 74    | 4.57    | 91.41  |
| 65              | 14    | 0.86    | 92.28  |
| 66              | 100   | 6.18    | 98.46  |
| 67              | 1     | 0.06    | 98.52  |
| 68              | 1     | 0.06    | 98.58  |
| 69              | 6     | 0.37    | 98.95  |
| 73              | 3     | 0.19    | 99.14  |
| 74              | 9     | 0.56    | 99.69  |
| 75              | 2     | 0.12    | 99.81  |
| 86              | 3     | 0.19    | 100.00 |
| Total           | 1,619 | 100.00  |        |

. gen mapres80t66=mapres80  
(1146 missing values generated)

